# The Global Increase in the Socioeconomic Achievement Gap, 1964-2015: Online Appendices

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ABSTRACT

These appendices report additional details of the results from the main text of the paper, as well as supplementary analyses undertaken to test the sensitivity of results to a number of different limitations of the data. The finding of global increases in SES achievement gaps is generally robust to differences across test instruments, changes in the distribution of achievement and of SES, and changes in the measurement error of achievement and of SES. The multivariate findings predicting changing country achievement gaps from changing country characteristics and policies are generally robust across a variety of model specifications.

VERSION

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**Online Appendices**

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The contents of the appendices are as follows:

A. List of countries and datasets included in the study
B. Combining different test instruments
C. Changing distribution of achievement
D. Changing measurement error of achievement
E. Changing distribution of SES
F. Achievement gaps by mother's and father's SES characteristics
G. Achievement gaps conditional on other SES variables
H. Changing measurement error of SES
J. Trend models run separately by SES variable
K. SES achievement gap trends by school level and subject
L. Specification of trend model
M. Specification of multivariate model
N. Trends in SES achievement gaps using a rank-based measure of achievement

**A1.** List of Included Countries and Datasets

| code | country | region | mid/low-income in 1980 | fims1964 | fiss1970gr4 | fiss1970gr8 | fircs1970gr4 | fircs1970gr8 | sims1980 | siss1984gr4 | siss1984gr8 | rls1991gr4 | rls1991gr8 | timss1995gr4 | timss2003gr4 | timss2007gr4 | timss2011gr4 | timss2015gr4 | timss1995gr8 | timss1999 | timss2003gr8 | timss2007gr8 | timss2011gr8 | timss2015gr8 | pirls2001 | pirls2006 | pirls2011 | pisa2000 | pisa2003 | pisa2006 | pisa2009 | pisa2012 | pisa2015 |
|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|
| ALB | Albania | E | x | | | | | | | | | | | | | | | | | | | | | | | | | x | | | x | (x) | (x) |
| DZA | Algeria | M | x | | | | | | | | | | | | | x | | | | | | x | | | | | | | | | | | x |
| ARG | Argentina | L | x | | | | | | | | | | | | | | | | | | | | | | x | | | x | | x | x | x | |
| ARM | Armenia | E | x | | | | | | | | | x | (x) | x | | | | | | x | x | x | | | | | | | | | | | |
| AUS | Australia | W | x | | | | x | | | x | x | | | x | x | x | x | x | x | x | x | x | x | x | | x | x | x | x | x | x | x | x |
| AUT | Austria | W | | | | | | | | | | | | x | | x | x | | | | x | | | | | x | x | x | x | x | x | x | x |
| AZE | Azerbaijan | E | x | | | | | | | | | | | | | x | | | | | | | | | | x | | | | | x | x | |
| BHR | Bahrain | M | | | | | | | | | | | | | (x) | x | | | | | x | x | x | x | | | | | | | | | |
| BEL | Belgium | W | | x | | | | | | | | | | | | | | | | | | | | | | | | | | | | | |
| BFL | Belgium-Flemish | W | | | x | x | x | x | x | | | | | x | | | | | x | x | x | x | x | | | x | | x | x | x | x | x | x |
| BFR | Belgium-French | W | | | x | x | x | x | | x | x | | | | | | | | | x | | | | | | x | | x | x | x | x | x | x |
| BLZ | Belize | L | x | | | | | | | | | | | | | | | | | | | | | | x | | | | | | | | |
| BIH | Bosnia & Herzegovina | E | x | | | | | | | | | | | | | | | | | | | x | | | | | | | | | | | |
| BWA | Botswana | Af | x | | | | | | | | x | | | | x | | | | | | x | x | x | | | x | | | | | | | |
| BRA | Brazil | L | x | | | | | | | | | | | | | | | | | | | | | | | | | x | x | x | x | x | x |
| BGR | Bulgaria | E | x | | | | | | | | | | | | | x | (m) | x | | x | x | x | | | x | x | x | x | x | x | x | x | x |
| CAN | Canada | W | | | | | | | (m) | (m) | | x | | x | x | x | | | | x | | x | | | x | x | x | x | x | x | x | x | x |
| CHL | Chile | L | x | | | (x) | x | x | x | | | | | | x | x | | | | | x | x | | | | x | | x | | x | x | x | x |
| CHN | China | As | x | | | | | | | x | | | | | | | | | | | | | | | | | | | | | | | |
| TWN | Chinese Taipei | As | | | | | | | | | | | | | x | x | x | x | | x | x | x | x | x | | x | x | | | x | x | x | x |
| COL | Colombia | L | x | | | | | | | | | | | x | | | x | | | | x | | x | | | x | | | | x | x | x | x |
| CRI | Costa Rica | L | x | | | | | | | | | | | | | | | | | | | | | | | | | | | | x | x | x |
| HRV | Croatia | E | | | | | | | | | | | | | | | x | x | | | | | | | | x | | | | x | x | x | x |
| CYP | Cyprus | W | x | | | | | | | | | (x) | x | x | x | | | | x | x | x | x | x | | | x | | | | | x | | |
| CZE | Czech Rep. | E | | | | | | | | | | | | x | | | | | x | x | x | x | x | | x | | x | x | x | x | x | x | x |
| DNK | Denmark | W | | | | | | | | x | x | | | | | | | | x | x | x | | | | | x | x | x | x | x | x | x | x |
| DOM | Dominican Republic | L | x | | | | | | | | | | | | | | | | | | | | | | | | | | | | | | x |
| EGY | Egypt | M | x | | | | | | | | | | | | | | | | | | x | x | | x | | | | | | | | | |
| SLV | El Salvador | L | x | | | | | | | | | | | | | x | | | | | | x | | | | | | | | | | | |
| ENG | England | W | | x | x | x | x | x | x | x | x | | | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x |
| EST | Estonia | E | | | | | | | | | | | | | | | | | | | x | | | | | | | | | x | x | x | x |
| FIN | Finland | W | | | x | x | x | x | x | x | x | x | x | | | | x | x | | | | | x | | | | x | x | x | x | x | x | x |
| FRA | France | W | | x | | | | | x | | | x | x | x | x | | | | | | | | | | x | x | | x | x | x | x | x | x |
| GEO | Georgia | E | x | | | | | | | | | | | | | x | x | x | | | x | x | x | | | x | x | | | | | x | | x |
| DEU | Germany | W | | | | | | | | | | | | x | x | x | x | | | | | | | | x | x | x | x | x | x | x | x | x |
| GDR | Germany-East | W | | | | | | | | x | x | | | | | | | | | | | | | | | | | | | | | | |
| FRG | Germany-West | W | | | x | x | x | | | x | x | | | | | | | | | | | | | | | | | | | | | | |
| GHA | Ghana | Af | x | | | | | | | | x | | | | | | | | | | x | x | x | | | | | | | | | | |
| GRC | Greece | W | | | | | | | | | | x | x | x | | | | | | x | | | | | | x | | x | x | x | x | x | x |
| HND | Honduras | L | x | | | | | | | | | | | | x | | | | | | x | | | | | x | | | | | | | |
| HKG | Hong Kong | As | | | | | | | | | x | x | x | (x) | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x |
| HUN | Hungary | E | x | | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x |
| ISL | Iceland | W | | | | | | | | | | x | x | x | | | | | | | | x | | | x | x | | x | x | x | x | x | x |
| IND | India | As | x | | x | x | x | x | | | | | | | | | | | | | | | | | | | | | | | | | |
| IDN | Indonesia | As | x | | | | | | | x | | | | (m) | | | x | | | x | x | x | x | | | x | | x | x | x | x | x | x |
| IRN | Iran | M | x | | | x | x | x | | | | | | x | x | x | x | x | x | x | x | x | x | x | x | x | x | | | | | | |
| IRL | Ireland | W | | | | | | | | x | x | x | | x | | | x | x | | | | | | | | | x | x | x | x | x | x | x |
| ISR | Israel | W | | x | | | x | x | x | x | (m) | | | x | | | | | x | x | x | x | x | | | x | | x | | x | x | x | x |
| ITA | Italy | W | | | x | x | x | x | | x | x | (x) | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x |
| JPN | Japan | As | | x | x | x | | | x | x | x | | | (m) | x | x | x | x | x | (m) | x | x | x | x | | | | x | x | x | x | x | x |
| JOR | Jordan | M | x | | | | | | | | | | | | | | | | | x | x | x | x | x | | | | | | x | x | x | x |
| KAZ | Kazakhstan | E | x | | | | | | | | | | | | | x | x | x | | | | x | x | | | | | | | | x | x | |
| KOR | Korea, Rep. | As | x | | | | | | | x | x | | | x | | | | | x | x | x | x | x | x | | | | x | x | x | x | x | x |
| KSV | Kosovo | E | x | | | | | | | | | | | | | | | | | | | | | | | | | | | | | | x |
| KWT | Kuwait | M | | | | | | | | | | | | x | | | | | x | x | x | x | | x | x | x | | | | | | | |
| KGZ | Kyrgyzstan | E | x | | | | | | | | | | | | | | | | | | | | | | | | | | | x | x | | |
| LVA | Latvia | E | x | | | | | | | | | | | x | x | x | | | x | x | x | | | | x | x | | x | x | x | x | x | x |

Notes: Shaded countries were excluded because they participated in only one test. (m) denotes missing SES data; (x) denotes that gaps could not be computed, usually because of low-quality SES data. Regions include: Af=sub-Saharan Africa, As=east and southeast Asian and Pacific countries, M=Middle Eastern and North African countries, E=Eastern Europe and the Commonwealth of Independent States, L=Latin America and the Caribbean, and W=Western countries (Western Europe and Anglophone countries).

| code | country | region | mid/low-income in 1980 | fims1964 | fiss1970gr4 | fiss1970gr8 | fircs1970gr4 | fircs1970gr8 | sims1980 | siss1984gr4 | siss1984gr8 | rls1991gr4 | rls1991gr8 | timss1995gr4 | timss2003gr4 | timss2007gr4 | timss2011gr4 | timss2015gr4 | timss1995gr8 | timss1999 | timss2003gr8 | timss2007gr8 | timss2011gr8 | timss2015gr8 | pirls2001 | pirls2006 | pirls2011 | pisa2000 | pisa2003 | pisa2006 | pisa2009 | pisa2012 | pisa2015 |
|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|
| LBN | Lebanon | M | x | | | | | | | | | | | | | | | | | | x | x | x | x | | | | | | | | | x |
| LIE | Liechtenstein | W | | | | | | | | | | | | | | | | | | | | | | | | | | x | x | x | x | x | |
| LTU | Lithuania | E | x | | | | | | | | | | | x | x | x | x | x | x | x | x | x | x | x | x | x | x | (m) | | x | x | x | x |
| LUX | Luxembourg | W | | | | | | x | | | | | | | | | | | | | | | | | | | | x | x | x | x | x | x |
| MAC | Macao-China | As | | | | | | | | | | | | | | | | | | | | | | | | | | | x | x | x | x | x |
| MKD | Macedonia | E | x | | | | | | | | | | | | | | | | | x | x | | x | | x | x | | x | | | | | x |
| MYS | Malaysia | As | x | | | | | | | | | | | | | | | | | x | x | x | x | x | | | | | | | x | x | |
| MLT | Malta | W | x | | | | | | | | | | x | | | | | | | | | | | x | | | x | | | | | x | x |
| MUS | Mauritius | Af | x | | | | | | | | | | | | | | | | | | | | | | | | | | | | | x | |
| MEX | Mexico | L | x | | | | | | (m) | | | | | | | | | | | | | | | | | | | x | x | x | x | x | x |
| MDA | Moldova | E | x | | | | | | | | | | | | | | | | | x | | x | x | | x | x | | | | | | | x |
| MNG | Mongolia | As | x | | | | | | | | | | | | | | | | | x | | x | | | | | | | | | | | |
| MNE | Montenegro | E | | | | | | | | | | | | | | | | | | | | | | | | | | | | x | x | x | x |
| MAR | Morocco | M | x | | | | | | | | | | | | | | | | x | x | x | x | x | x | x | x | x | | | | (x) | x | x |
| NLD | Netherlands | W | | | x | x | x | x | x | | x | x | x | x | x | x | x | x | x | x | x | x | x | | x | x | x | x | x | x | x | x | x |
| NZL | New Zealand | W | | | | x | | x | x | | x | x | x | x | x | x | x | x | x | x | x | x | x | | x | x | x | x | x | x | x | x | x |
| NGA | Nigeria | Af | x | | | | | | (x) | x | x | | x | | | | | | | | | | | | | | | | | | | | |
| NIR | Northern Ireland | W | | | | | | | | | | | | | | x | x | | | | | | | | | | x | | | | | | |
| NOR | Norway | W | | | | | | | | | x | x | x | x | x | x | x | x | x | x | x | | x | | x | x | x | x | x | x | x | x | x |
| OMN | Oman | M | x | | | | | | | | | | | | | x | x | | | | | | x | x | | | x | | | | | | |
| PAN | Panama | L | x | | | | | | | | | | | | | | | | | | | | | | | | | | | | x | | |
| PNG | Papua New Guinea | As | x | | | | | | | x | | | | | | | | | | | | | | | | | | | | | | | |
| PER | Peru | L | x | | | | | | | | | | | | | | | | | | | | | | | | | x | | | x | x | x |
| PHL | Philippines | As | x | | | | | | | x | x | | x | | | | | | x | x | x | | | | | | | | | | | | |
| POL | Poland | E | x | | | | | | | x | x | | | | | | | | | | | x | x | | x | x | x | x | x | x | x | x | x |
| PRT | Portugal | W | x | | | | | | | | | x | x | x | | | | | x | x | x | | | | | | x | x | x | x | x | x | x |
| QAT | Qatar | M | x | | | | | | | | | | | | | x | x | x | | | | x | x | x | | | x | | | x | x | x | x |
| ROM | Romania | E | x | | | | | | | | | | | | | | | | x | x | x | x | x | | x | x | x | | | x | x | x | x |
| RUS | Russian Fed. | E | | | | | | | | | | | | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x |
| SAU | Saudi Arabia | M | | | | | | | | | | | | | | | | | x | x | x | x | x | x | | | x | | | | | | |
| SCO | Scotland | W | | | x | x | x | x | x | x | | x | x | x | | | | | x | x | x | | | | | x | x | x | x | x | x | x | x |
| SRB | Serbia | E | x | | | | | | | | | | | | | | | | | | x | x | x | | | | | | x | x | x | x |
| QCN | Shanghai-China | As | x | | | | | | | | | | | | | | | | | | | | | | | | | | | | x | x | |
| SGP | Singapore | As | | | | | | | | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | | | | x | x | x |
| SVK | Slovak Rep. | E | x | | | | | | | | | | | | | | | | x | x | x | x | x | | x | x | x | | x | x | x | x | x |
| SVN | Slovenia | E | | | | | | | | | | | | x | x | x | x | x | x | x | x | x | x | x | x | x | x | | | x | x | x | x |
| ZAF | South Africa | Af | x | | | | | | | | | | | | | | | | x | x | x | | x | x | | x | x | | | | | | |
| ESP | Spain | W | x | | | | | | | x | x | | | | | | | | x | x | x | | | | x | x | x | x | x | x | x | x | x |
| SWZ | Swaziland | Af | x | | | | | | x | | | | | | | | | | | | | | | | | | | | | | | | |
| SWE | Sweden | W | | (m) | x | x | x | x | x | x | x | x | x | x | | | | | x | x | x | | | | x | x | x | x | x | x | x | x | x |
| CHE | Switzerland | W | | | | | | | | | | | | | | | | | x | x | | | x | | | | | x | x | x | x | x | x |
| SYR | Syria | M | x | | | | | | | | | | | | | | | | | | x | x | x | | | | | | | | | | |
| THA | Thailand | As | x | | | x | | x | x | | x | | | | | | | | x | x | | | x | | | | | x | x | x | x | x | x |
| TTO | Trinidad and Tobago | L | | | | | | | | x | x | | | | | | | | | | | | | | | x | x | | | | x | | x |
| TUN | Tunisia | M | x | | | | | | | | | | | | | | | | | x | x | x | x | | | | | | x | x | x | x |
| TUR | Turkey | W | x | | | | | | | | | | | | | | | | | x | | | x | | | | | | x | x | x | x | x |
| UKR | Ukraine | E | x | | | | | | | | | | | | | | | | | x | | x | x | | | | | | | | | | |
| ARE | United Arab Emirates | M | | | | | | | | | | | | | | | | | x | x | | x | x | | | | x | | | | x | x | x |
| USA | United States | W | | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | x | (x) | x | x | x | x | x |
| URY | Uruguay | L | x | | | | | | | | | | | | | | | | | | | | | | | | | | x | x | x | x | x |
| VEN | Venezuela | L | x | | | | | | x | x | | | | | | | | | | | | | | | | | | | | | | | |
| VNM | Vietnam | As | x | | | | | | | | | | | | | | | | | | | | | | | | | | | | (x) | x | x |
| PSE | West Bank and Gaza | M | x | | | | | | | | | | | | | | | | | | x | x | x | | | | | | | | | | |
| YEM | Yemen | M | x | | | | | | | | | | | (x) | x | x | | | | | | | | | | | | | | | | | |
| ZWE | Zimbabwe | Af | x | | | | | | | | x | | x | | | | | | | | | | | | | | | | | | | | |

Notes: Shaded countries were excluded because they participated in only one test. (m) denotes missing SES data; (x) denotes that gaps could not be computed, usually because of low-quality SES data. Regions include: Af=sub-Saharan Africa, As=east and southeast Asian and Pacific countries, M=Middle Eastern and North African countries, E=Eastern Europe and the Commonwealth of Independent States, L=Latin America and the Caribbean, and W=Western countries (Western Europe and Anglophone countries).

**B. Combining different test instruments**

       In order to estimate international trends in SES achievement gaps over a 50-year period, this paper combines data from a variety of different international assessments of math, science, and reading. However, among tests of the same subject, a comparison of skills frameworks from the official study reports reveals differences. For example, the IEA math and science tests—FIMS, FISS, SIMS, SISS, and TIMSS—are curriculum-based, while the PISA math and science tests, as well as all reading tests—PISA, FIRCS, RLS, and PIRLS—are literacy-based. Though the early IEA tests contained anchor items to enable studying trends in achievement, the scores were not placed on common scales, and they did not have the advantage of improvements to testing methodology in the 1990s; thus, early and recent IEA tests are not strictly comparable. The analyses in the main text of the paper deal with this issue by standardizing achievement within each study and each country and assuming only that each test is interval-scaled and that different tests rank students similarly.

       However, there are six recent studies that repeat the same test instrument to enable measuring trends over time: TIMSS $4^{th}$ and $8^{th}$ grade math and science; PIRLS; and PISA reading, math, and science. These trend studies allow us to investigate the sensitivity of gap trend results to differences across test instruments—but only over the recent 9-20 years that the studies have been conducted. TIMSS trends (for both grades and subjects) can be estimated for test years 1995 to 2015 (birth cohorts 1981 to 2001 for $8^{th}$ grade and birth cohorts 1985 to 2005 for $4^{th}$ grade); PIRLS trends for test years 2001 to 2011 (birth cohorts 1991 to 2001); PISA reading trends for test years 2000 to 2015 (birth cohorts 1985 to 2000); PISA math trends for test years 2003 to 2015 (birth cohorts 1988 to 2000); and PISA science trends for test years 2006 to 2015 (birth cohorts 1991 to 2000). In addition, because each instrument remains the same over time, it is not necessary to standardize achievement within studies or countries, meaning we can examine changes in SES achievement gaps in light of possible changes in the variance of skills (which will be addressed in Appendix C).

       Tables B1-B4 compute trends in SES achievement gaps separately for each test instrument. For each of the eight test instruments, Model 1 is a hierarchical multivariate variance-known model that estimates a different cohort slope for each gap type (parent education, parent occupation, or household books) using interactions between cohort birth year and gap type indicators. Model 2 estimates a single pooled cohort slope for all gap types. Thus, the specifications of Models 1 and 2 are the same as Models 1 and 2 in the main text but without control variables (as the subject and age at testing dummies from the models in the main text do not vary within each test instrument). Tables B1 and B2 report trends in gaps estimated without standardizing achievement within each country-study-year. Tables B3 and B4 report gaps estimated when achievement is standardized within each county-study-year, as in the main text of the paper.[1,2] It can be seen from the reported country sample sizes that the number of participating countries varies widely across the different test instruments. It is not possible to reliably estimate trends across all test instruments for a core group of countries that has participated in every test, as there are too few countries that have done so. Thus, the reported

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[1] The Standardized Achievement models adjust each gap estimate for the reliabilities of test instruments and SES report; the Unstandardized Achievement models adjust each gap estimate only for the reliability of SES report. It is not necessary to adjust for test reliability in these models as gap estimates are not attenuated since they have not been divided by the test score variance.

[2] When computing these gaps, rather than using all available categories of each SES variable as in the main text of the paper, each SES variable was recoded to ensure that the SES instrument remained the same across test years. There were six categories for parent education and parent occupation and five categories for household books.

trends for each test instrument should be interpreted only as a rough indication of the sensitivity of the general finding of increasing SES achievement gaps over time. The size of coefficients can be compared across different test instruments only for the models using standardized achievement, not for those using unstandardized achievement, as they are in different metrics and 1 point in PISA, for example, is not the same as 1 point in PIRLS or TIMSS. Significance levels should be interpreted with caution because of changing sample sizes and the large number of significance tests conducted; significance is reported only as a rough indication of the precision of each estimate.

Overall, the estimated gap trends are positive for most test instruments. Gaps are consistently increasing for all three SES variables in PIRLS and for both available SES variables in both subjects of TIMSS at both the 4th and the 8th grades. However, gap trends for PISA are more mixed. In particular, trends in gaps based on unstandardized math and science achievement are negative, while trends in gaps based on standardized math and science achievement and based on reading achievement (standardized or unstandardized) are more consistently positive. This is likely due to substantial declines in the variance of PISA math and science scores (see Appendix C). Additionally, for all three PISA subject tests, trends in gaps based on parent education are negative. Further analysis shows that this pattern may be due to problems in the measurement of parent education in the PISA student survey (see Appendix H). Therefore, with the exception of some unreliable trends, positive increases in SES achievement gaps over time are quite robust across the different test instruments that are combined in the main text of the paper.

Table B1. Coefficients from Hierarchical Growth Models Predicting Achievement Gaps between 90th and 10th Percentiles of SES, Run Separately by Test Instrument (Unstandardized Achievement) – PISA and PIRLS

| Test instrument | PISA Math | | PISA Reading | | PISA Science | | PIRLS Reading | |
|---|---|---|---|---|---|---|---|---|
| Test years | 2003-2015 | | 2000-2015 | | 2006-2015 | | 2001-2011 | |
| Cohort birth years | 1988-2000 | | 1985-2000 | | 1991-2000 | | 1991-2001 | |
| | (1) | (2) | (1) | (2) | (1) | (2) | (1) | (2) |
| Parent education gaps intercept | 88.580 *** | 86.591 *** | 85.725 *** | 82.817 *** | 92.806 *** | 86.595 *** | 76.698 *** | 78.551 *** |
| | (3.696) | (3.381) | (3.337) | (3.130) | (4.189) | (3.472) | (3.715) | (3.840) |
| Parent occupation gaps intercept | 86.131 *** | 85.944 *** | 83.782 *** | 84.089 *** | 84.459 *** | 85.742 *** | 68.610 *** | 68.730 *** |
| | (3.328) | (3.278) | (3.180) | (3.193) | (3.602) | (3.620) | (3.579) | (3.447) |
| Household books gaps intercept | 127.448 *** | 127.860 *** | 125.759 *** | 128.238 *** | 132.753 *** | 132.500 *** | 87.755 *** | 86.356 *** |
| | (4.460) | (4.364) | (4.230) | (4.380) | (5.621) | (5.252) | (3.864) | (3.975) |
| Cohort birth year × Parent education | -0.491 * | | -0.374 * | | -0.865 ** | | 0.769 + | |
| | (0.225) | | (0.188) | | (0.263) | | (0.407) | |
| Cohort birth year × Parent occupation | -0.288 + | | 0.049 | | -0.124 | | 0.507 | |
| | (0.173) | | (0.172) | | (0.189) | | (0.358) | |
| Cohort birth year × Books | -0.206 | | 0.377 | | -0.261 | | 0.224 | |
| | (0.272) | | (0.276) | | (0.286) | | (0.324) | |
| Cohort birth year | | -0.266 | | 0.027 | | -0.242 | | 0.470 |
| | | (0.174) | | (0.174) | | (0.194) | | (0.319) |
| **Random effects** | | | | | | | | |
| *Level 2 - Residual variance between studies in…* | | | | | | | | |
| Parent education intercepts | 83.585 | 80.729 | 59.861 | 55.654 | 49.099 | 58.929 | 46.742 | 45.620 |
| Parent occupation intercepts | 34.517 | 35.085 | 32.022 | 31.198 | 29.029 | 29.485 | 35.950 | 40.488 |
| Books intercepts | 78.836 | 98.737 | 55.998 | 73.371 | 44.217 | 55.873 | 86.374 | 86.604 |
| *Level 3 - Residual variance between countries in…* | | | | | | | | |
| Parent education intercepts | 629.437 | 601.335 | 605.245 | 560.178 | 713.887 | 564.316 | 408.697 | 465.914 |
| Parent occupation intercepts | 629.187 | 600.721 | 614.484 | 630.121 | 689.660 | 679.301 | 408.276 | 364.008 |
| Books intercepts | 1048.458 | 1152.394 | 1057.018 | 1245.643 | 1737.470 | 1668.747 | 433.859 | 491.341 |
| Parent education cohort slopes | 0.649 | | 0.806 | | 0.845 | | 3.323 | |
| Parent occupation cohort slopes | 0.802 | | 1.015 | | 0.769 | | 2.778 | |
| Books cohort slopes | 1.838 | | 2.996 | | 1.757 | | 0.708 | |
| Cohort slopes | | 0.737 | | 1.198 | | 0.684 | | 1.671 |
| % Parent education cohort slopes > 0 | 0.271 | | 0.339 | | 0.173 | | 0.663 | |
| % Parent occupation cohort slopes > 0 | 0.374 | | 0.519 | | 0.444 | | 0.620 | |
| % Books cohort slopes > 0 | 0.440 | | 0.586 | | 0.422 | | 0.605 | |
| % Cohort slopes > 0 | | 0.378 | | 0.510 | | 0.385 | | 0.642 |
| N (Level 1 - gaps) | 893 | 893 | 1030 | 1030 | 764 | 764 | 300 | 300 |
| N (Level 2 - study-years) | 298 | 298 | 344 | 344 | 255 | 255 | 103 | 103 |
| N (Level 3 - countries) | 70 | 70 | 72 | 72 | 70 | 70 | 41 | 41 |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: To avoid very long computation times, all models in this table specify known level 1 error variances estimated using conventional non-bootstrap formulas and omit error covariances. This simplified specification appears to produce very similar results to models using bootstrapped error variances and covariances (see Appendix L). "% cohort slopes > 0" is computed by comparing cohort coefficient estimate and cohort slope variance component to a normal distribution.

Table B2. Coefficients from Hierarchical Growth Models Predicting Achievement Gaps between 90th and 10th Percentiles of SES, Run Separately by Test Instrument (Unstandardized Achievement) – TIMSS

| Test instrument | TIMSS Grade 4 Math | | TIMSS Grade 4 Science | | TIMSS Grade 8 Math | | TIMSS Grade 8 Science | |
|---|---|---|---|---|---|---|---|---|
| Test years | 1995-2015 | | 1995-2015 | | 1995-2015 | | 1995-2015 | |
| Cohort birth years | 1985-2005 | | 1985-2005 | | 1981-2001 | | 1981-2001 | |
| | (1) | (2) | (1) | (2) | (1) | (2) | (1) | (2) |
| Parent education gaps intercept | | | | | 70.929 *** | 70.212 *** | 70.426 *** | 68.680 *** |
| | | | | | (3.629) | (3.561) | (4.224) | (3.998) |
| Household books gaps intercept | 74.438 *** | 74.438 *** | 80.765 *** | 80.765 *** | 80.117 *** | 80.820 *** | 81.421 *** | 83.011 *** |
| | (4.988) | (4.988) | (5.859) | (5.859) | (4.473) | (4.505) | (5.034) | (5.132) |
| Cohort birth year × Parent education | | | | | 1.225 *** | | 1.267 *** | |
| | | | | | (0.255) | | (0.269) | |
| Cohort birth year × Books | 0.775 ** | | 0.682 * | | 1.389 *** | | 1.591 *** | |
| | (0.275) | | (0.312) | | (0.309) | | (0.271) | |
| Cohort birth year | | 0.775 ** | | 0.682 * | | 1.318 *** | | 1.450 *** |
| | | (0.275) | | (0.312) | | (0.257) | | (0.240) |
| Random effects | | | | | | | | |
| *Level 2 - Residual variance between studies in…* | | | | | | | | |
| Parent education intercepts | | | | | 127.994 | 122.721 | 157.985 | 163.727 |
| Books intercepts | 38.198 | 38.198 | 109.020 | 109.020 | 138.450 | 157.292 | 162.137 | 167.889 |
| *Level 3 - Residual variance between countries in…* | | | | | | | | |
| Parent education intercepts | | | | | 441.812 | 460.884 | 670.049 | 624.631 |
| Books intercepts | 681.626 | 681.626 | 903.670 | 903.670 | 817.700 | 925.355 | 1116.629 | 1258.709 |
| Parent education cohort slopes | | | | | 1.501 | | 1.572 | |
| Books cohort slopes | 1.526 | | 1.513 | | 2.974 | | 1.590 | |
| Cohort slopes | | 1.526 | | 1.513 | | 1.990 | | 1.345 |
| % Parent education cohort slopes > 0 | | | | | 0.841 | | 0.844 | |
| % Books cohort slopes > 0 | 0.735 | | 0.710 | | 0.790 | | 0.897 | |
| % Cohort slopes > 0 | | 0.735 | | 0.710 | | 0.825 | | 0.894 |
| N (Level 1 - gaps) | 163 | 163 | 165 | 165 | 485 | 485 | 485 | 485 |
| N (Level 2 - study-years) | 163 | 163 | 165 | 165 | 245 | 245 | 245 | 245 |
| N (Level 3 - countries) | 49 | 49 | 50 | 50 | 61 | 61 | 61 | 61 |

+ $p<.1$, * $p<.05$, ** $p<.01$, *** $p < .001$. Note: To avoid very long computation times, all models in this table specify known level 1 error variances estimated using conventional non-bootstrap formulas and omit error covariances. This simplified specification appears to produce very similar results to models using bootstrapped error variances and covariances (see Appendix L). "% cohort slopes > 0" is computed by comparing cohort coefficient estimate and cohort slope variance component to a normal distribution.

Table B3. Coefficients from Hierarchical Growth Models Predicting Achievement Gaps between 90th and 10th Percentiles of SES, Run Separately by Test Instrument (Standardized Achievement) – PISA and PIRLS

| Test instrument | PISA Math | | PISA Reading | | PISA Science | | PIRLS Reading | |
|---|---|---|---|---|---|---|---|---|
| Test years | 2003-2015 | | 2000-2015 | | 2006-2015 | | 2001-2011 | |
| Cohort birth years | 1988-2000 | | 1985-2000 | | 1991-2000 | | 1991-2001 | |
| | (1) | (2) | (1) | (2) | (1) | (2) | (1) | (2) |
| Parent education gaps intercept | 1.000 *** | 0.972 *** | 0.946 *** | 0.912 *** | 1.044 *** | 0.980 *** | 1.042 *** | 1.066 *** |
| | (0.037) | (0.034) | (0.032) | (0.030) | (0.041) | (0.034) | (0.043) | (0.039) |
| Parent occupation gaps intercept | 0.971 *** | 0.964 *** | 0.923 *** | 0.925 *** | 0.953 *** | 0.968 *** | 0.928 *** | 0.934 *** |
| | (0.034) | (0.033) | (0.029) | (0.029) | (0.034) | (0.034) | (0.043) | (0.040) |
| Household books gaps intercept | 1.438 *** | 1.455 *** | 1.401 *** | 1.424 *** | 1.506 *** | 1.505 *** | 1.209 *** | 1.192 *** |
| | (0.044) | (0.043) | (0.040) | (0.042) | (0.051) | (0.048) | (0.041) | (0.046) |
| Cohort birth year × Parent education | -0.001 | | -0.002 | | -0.007 ** | | 0.013 ** | |
| | (0.002) | | (0.002) | | (0.003) | | (0.004) | |
| Cohort birth year × Parent occupation | 0.001 | | 0.002 | | 0.001 | | 0.010 ** | |
| | (0.002) | | (0.002) | | (0.002) | | (0.004) | |
| Cohort birth year × Books | 0.004 | | 0.006 * | | -0.001 | | 0.006 | |
| | (0.003) | | (0.003) | | (0.002) | | (0.004) | |
| Cohort birth year | | 0.002 | | 0.002 + | | -0.001 | | 0.009 ** |
| | | (0.002) | | (0.001) | | (0.002) | | (0.003) |
| **Random effects** | | | | | | | | |
| *Level 2 - Residual variance between studies in…* | | | | | | | | |
| Parent education intercepts | 0.00693 | 0.00662 | 0.00754 | 0.00721 | 0.00498 | 0.00513 | 0.00647 | 0.00963 |
| Parent occupation intercepts | 0.00217 | 0.00212 | 0.00227 | 0.00239 | 0.00251 | 0.00239 | 0.00457 | 0.00589 |
| Books intercepts | 0.00413 | 0.00534 | 0.00353 | 0.00472 | 0.00274 | 0.00288 | 0.01014 | 0.01203 |
| *Level 3 - Residual variance between countries in…* | | | | | | | | |
| Parent education intercepts | 0.06126 | 0.05959 | 0.05010 | 0.04805 | 0.06348 | 0.05186 | 0.04715 | 0.03904 |
| Parent occupation intercepts | 0.06315 | 0.05965 | 0.04929 | 0.05019 | 0.05750 | 0.05809 | 0.06038 | 0.04729 |
| Books intercepts | 0.10118 | 0.11315 | 0.08929 | 0.11116 | 0.13984 | 0.13857 | 0.04441 | 0.06193 |
| Parent education cohort slopes | 0.00006 | | 0.00004 | | 0.00002 | | 0.00021 | |
| Parent occupation cohort slopes | 0.00007 | | 0.00005 | | 0.00003 | | 0.00023 | |
| Books cohort slopes | 0.00015 | | 0.00019 | | 0.00002 | | 0.00007 | |
| Cohort slopes | | 0.00007 | | 0.00006 | | 0.00002 | | 0.00005 |
| % Parent education cohort slopes > 0 | 0.432 | | 0.354 | | 0.060 | | 0.818 | |
| % Parent occupation cohort slopes > 0 | 0.550 | | 0.629 | | 0.555 | | 0.750 | |
| % Books cohort slopes > 0 | 0.624 | | 0.660 | | 0.431 | | 0.776 | |
| % Cohort slopes > 0 | | 0.586 | | 0.621 | | 0.440 | | 0.900 |
| N (Level 1 - gaps) | 893 | 893 | 1030 | 1030 | 764 | 764 | 300 | 300 |
| N (Level 2 - study-years) | 298 | 298 | 344 | 344 | 255 | 255 | 103 | 103 |
| N (Level 3 - countries) | 70 | 70 | 72 | 72 | 70 | 70 | 41 | 41 |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: To avoid very long computation times, all models in this table specify known level 1 error variances estimated using conventional non-bootstrap formulas and omit error covariances. This simplified specification appears to produce very similar results to models using bootstrapped error variances and covariances (see Appendix L). "% cohort slopes > 0" is computed by comparing cohort coefficient estimate and cohort slope variance component to a normal distribution.

Table B4. Coefficients from Hierarchical Growth Models Predicting Achievement Gaps between 90th and 10th Percentiles of SES, Run Separately by Test Instrument (Standardized Achievement) – TIMSS

| Test instrument | TIMSS Grade 4 Math | | TIMSS Grade 4 Science | | TIMSS Grade 8 Math | | TIMSS Grade 8 Science | |
|---|---|---|---|---|---|---|---|---|
| Test years | 1995-2015 | | 1995-2015 | | 1995-2015 | | 1995-2015 | |
| Cohort birth years | 1985-2005 | | 1985-2005 | | 1981-2001 | | 1981-2001 | |
| | (1) | (2) | (1) | (2) | (1) | (2) | (1) | (2) |
| Parent education gaps intercept | | | | | 0.933 *** | 0.934 *** | 0.927 *** | 0.918 *** |
| | | | | | (0.044) | (0.042) | (0.049) | (0.043) |
| Household books gaps intercept | 1.011 *** | 1.011 *** | 1.046 *** | 1.046 *** | 1.066 *** | 1.068 *** | 1.100 *** | 1.113 *** |
| | (0.068) | (0.068) | (0.077) | (0.077) | (0.052) | (0.053) | (0.058) | (0.061) |
| Cohort birth year × Parent education | | | | | 0.012 *** | | 0.013 *** | |
| | | | | | (0.003) | | (0.003) | |
| Cohort birth year × Books | 0.013 *** | | 0.015 *** | | 0.013 *** | | 0.016 *** | |
| | (0.004) | | (0.004) | | (0.003) | | (0.003) | |
| Cohort birth year | | 0.013 *** | | 0.015 *** | | 0.013 *** | | 0.014 *** |
| | | (0.004) | | (0.004) | | (0.003) | | (0.002) |
| Random effects | | | | | | | | |
| *Level 2 - Residual variance between studies in…* | | | | | | | | |
| Parent education intercepts | | | | | 0.01337 | 0.01365 | 0.01571 | 0.01665 |
| Books intercepts | 0.00339 | 0.00339 | 0.01708 | 0.01708 | 0.01152 | 0.01331 | 0.01766 | 0.01863 |
| *Level 3 - Residual variance between countries in…* | | | | | | | | |
| Parent education intercepts | | | | | 0.06882 | 0.06648 | 0.09139 | 0.07122 |
| Books intercepts | 0.14078 | 0.14078 | 0.16034 | 0.16034 | 0.11734 | 0.13454 | 0.15155 | 0.18369 |
| Parent education cohort slopes | | | | | 0.00011 | | 0.00009 | |
| Books cohort slopes | 0.00026 | | 0.00026 | | 0.00034 | | 0.00020 | |
| Cohort slopes | | 0.00026 | | 0.00026 | | 0.00018 | | 0.00009 |
| % Parent education cohort slopes > 0 | | | | | 0.884 | | 0.918 | |
| % Books cohort slopes > 0 | 0.784 | | 0.827 | | 0.765 | | 0.870 | |
| % Cohort slopes > 0 | | 0.784 | | 0.827 | | 0.829 | | 0.931 |
| N (Level 1 - gaps) | 163 | 163 | 165 | 165 | 485 | 485 | 485 | 485 |
| N (Level 2 - study-years) | 163 | 163 | 165 | 165 | 245 | 245 | 245 | 245 |
| N (Level 3 - countries) | 49 | 49 | 50 | 50 | 61 | 61 | 61 | 61 |

+ $p<.1$, * $p<.05$, ** $p<.01$, *** $p < .001$. Note: To avoid very long computation times, all models in this table specify known level 1 error variances estimated using conventional non-bootstrap formulas and omit error covariances. This simplified specification appears to produce very similar results to models using bootstrapped error variances and covariances (see Appendix L). "% cohort slopes > 0" is computed by comparing cohort coefficient estimate and cohort slope variance component to a normal distribution.

Figures B1 and B2 illustrate average global trends in 90/10 SES achievement gaps for each of the eight test instruments, estimated from the cohort birth year slope in Model 2 (based on standardized achievement). Figure B1 displays cohort slopes for the four instruments in Table B3, and Figure B2 displays cohort slopes for the four instruments in Table B4. The figures show that trends for all test instruments are positive except for PISA science, which is the PISA instrument that is available for the shortest span of cohort years. The trend in SES gaps for PISA science scores is very close to flat.

Figure B1
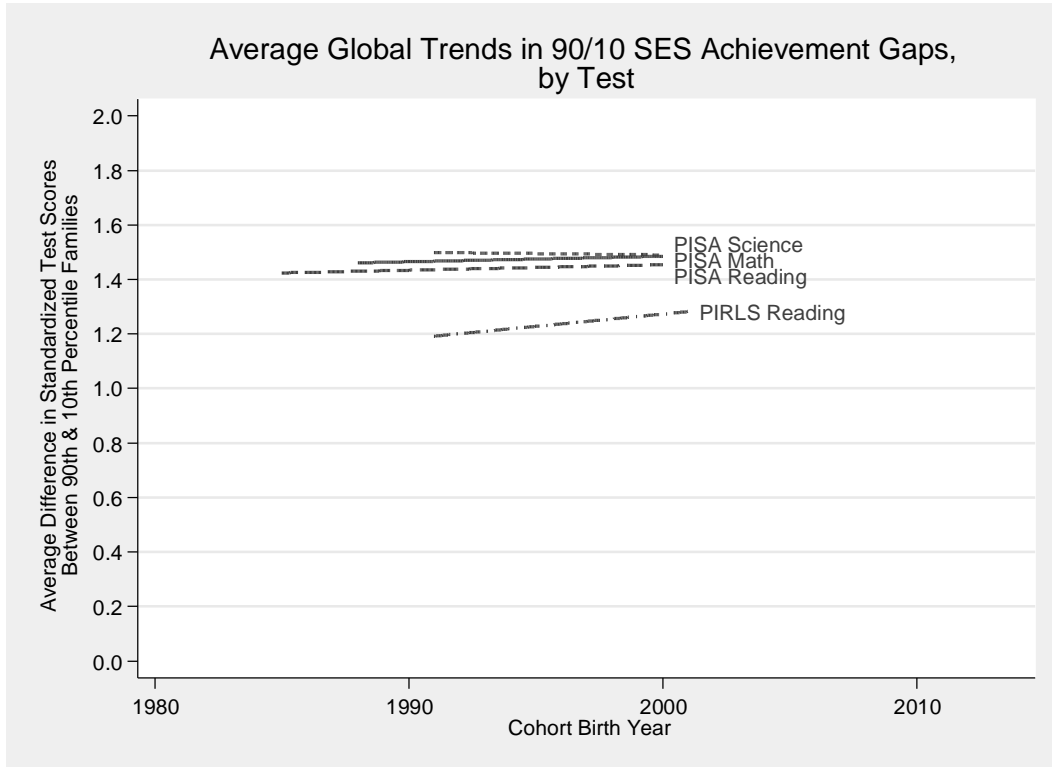


Note: Trend lines are estimates from Table B3, setting SES variable equal to parent education.

Figure B2
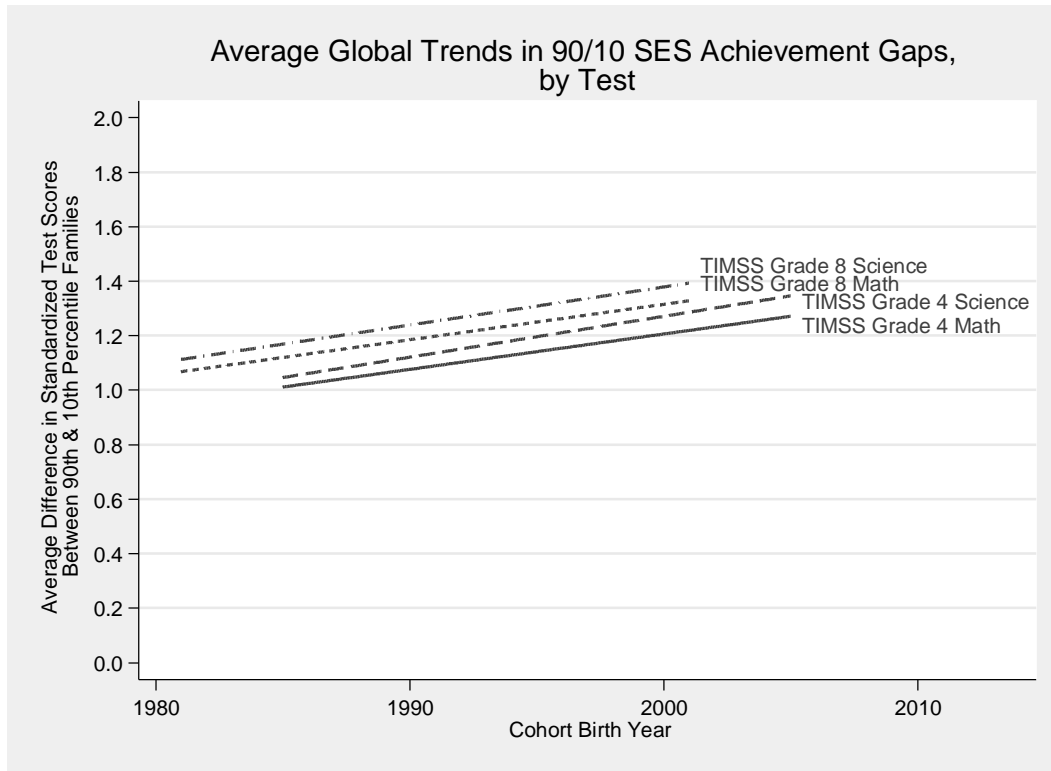


Note: Trend lines are estimates from Table B4, setting SES variable equal to parent education.

The lower 'Random effects' portions of Tables B1-B4 contain the residual country-level variance in cohort slopes, which quantifies how much country gap trends vary around the average global trend for each test instrument. By comparing the cohort coefficient estimate and cohort slope variance component from each model to a normal distribution, we can compute the estimated proportion of countries with positive gap trends. This is reported in each table below the random effects. Naturally, for all test instruments where average cohort coefficient estimates are positive, we estimate that a majority of countries have increasing gaps. Even for PISA, where negative average cohort coefficients indicate that a majority of countries have decreasing gaps, we still estimate that a substantial number of countries have increasing gaps.

To further understand cross-national variation by test instrument, Tables B5-B12 present trend estimates by country for each test. In each table, I present trends only for countries with at least 4 available study-years (except in PIRLS, where only 3 study-years are available). Even for this restricted sample, country trend estimates are highly unreliable due to the small number of observed years, which can be seen by the large discrepancies between OLS and shrunken empirical Bayes (EB) estimates. I prefer OLS estimates for these estimates based on very sparse country data because EB estimates tend to be biased toward the grand mean trend, and thus will underestimate cross-national variability in gap trends. Consistent with the model results reported above, for every test instrument except PISA science, a majority of countries have increasing gaps when achievement is standardized within country-study-year. When achievement is unstandardized, a smaller share of countries has positive gap trends, though still the majority for all tests except PISA.

Table B5. Estimated Change in 90/10 SES Achievement Gaps, PISA Math 2003-2015 (1988-2000 Birth Cohorts)

| Country | N (study-years) | Standardized Achievement OLS cohort trend | EB cohort trend | (SE) | Unstandardized Achievement OLS cohort trend | EB cohort trend | (SE) |
|---|---|---|---|---|---|---|---|
| Mexico | 5 | -0.024 | -0.011 | (0.006) | -2.757 | -1.470 | (0.496) |
| Chile | 4 | -0.024 | -0.006 | (0.006) | -2.038 | -0.919 | (0.566) |
| Slovenia | 4 | -0.021 | -0.002 | (0.006) | -2.304 | -0.395 | (0.529) |
| Lithuania | 4 | -0.020 | -0.003 | (0.006) | -1.869 | -0.423 | (0.505) |
| Brazil | 5 | -0.019 | -0.011 | (0.005) | -3.255 | -1.880 | (0.469) |
| Germany | 5 | -0.019 | -0.007 | (0.005) | -3.114 | -1.120 | (0.470) |
| Netherlands | 5 | -0.017 | -0.004 | (0.005) | -1.540 | -0.418 | (0.474) |
| Liechtenstein | 4 | -0.014 | -0.003 | (0.007) | -1.749 | -0.877 | (0.627) |
| Qatar | 4 | -0.014 | -0.006 | (0.006) | -0.691 | -1.130 | (0.521) |
| Jordan | 4 | -0.013 | -0.005 | (0.006) | -1.339 | -0.974 | (0.524) |
| Tunisia | 5 | -0.013 | -0.007 | (0.006) | -1.391 | -1.170 | (0.525) |
| Thailand | 5 | -0.011 | -0.004 | (0.006) | -0.931 | -0.809 | (0.533) |
| Poland | 5 | -0.010 | -0.002 | (0.005) | -1.128 | -0.444 | (0.474) |
| Estonia | 4 | -0.009 | -0.002 | (0.006) | -1.333 | -0.637 | (0.494) |
| England | 5 | -0.008 | 0.000 | (0.005) | -0.821 | -0.182 | (0.446) |
| Bulgaria | 4 | -0.005 | -0.002 | (0.007) | -1.440 | -0.623 | (0.593) |
| Norway | 5 | -0.005 | 0.002 | (0.005) | -1.015 | -0.191 | (0.449) |
| United States | 5 | -0.002 | 0.003 | (0.005) | -0.631 | 0.038 | (0.439) |
| Montenegro | 4 | -0.001 | 0.002 | (0.006) | -0.150 | -0.180 | (0.499) |
| Russian Fed. | 5 | -0.001 | 0.002 | (0.005) | -0.458 | -0.126 | (0.448) |
| Portugal | 5 | -0.001 | 0.000 | (0.005) | 0.686 | -0.160 | (0.467) |
| Scotland | 5 | -0.001 | 0.004 | (0.005) | -0.241 | 0.254 | (0.462) |
| Slovak Rep. | 5 | 0.000 | 0.003 | (0.005) | 0.354 | 0.086 | (0.478) |
| Austria | 5 | 0.001 | 0.004 | (0.006) | -0.064 | 0.231 | (0.525) |
| Switzerland | 5 | 0.001 | 0.003 | (0.005) | -0.411 | -0.062 | (0.500) |
| Japan | 5 | 0.001 | 0.006 | (0.005) | -0.569 | 0.246 | (0.474) |
| Colombia | 4 | 0.001 | 0.003 | (0.006) | -0.705 | -0.175 | (0.502) |
| Turkey | 5 | 0.002 | 0.007 | (0.006) | -1.355 | 0.260 | (0.493) |
| Greece | 5 | 0.003 | 0.001 | (0.005) | -0.291 | -0.514 | (0.456) |
| France | 5 | 0.004 | 0.003 | (0.005) | 0.569 | 0.148 | (0.452) |
| Romania | 4 | 0.004 | 0.003 | (0.007) | 0.552 | -0.005 | (0.580) |
| Czech Rep. | 5 | 0.005 | 0.002 | (0.006) | -0.492 | -0.425 | (0.506) |
| Australia | 5 | 0.006 | 0.005 | (0.005) | 0.390 | 0.057 | (0.446) |
| Hong Kong | 5 | 0.008 | 0.010 | (0.005) | 0.251 | 0.660 | (0.466) |
| Canada | 5 | 0.008 | 0.006 | (0.005) | 0.536 | 0.144 | (0.423) |
| Ireland | 5 | 0.009 | 0.008 | (0.005) | 0.212 | 0.350 | (0.446) |
| Belgium-French | 5 | 0.009 | -0.001 | (0.005) | -1.348 | -1.170 | (0.476) |
| Chinese Taipei | 4 | 0.010 | 0.010 | (0.006) | 1.513 | 0.897 | (0.516) |
| Uruguay | 5 | 0.010 | 0.002 | (0.005) | -0.665 | -0.678 | (0.482) |
| Finland | 5 | 0.010 | 0.008 | (0.005) | 0.666 | 0.366 | (0.438) |
| Serbia | 4 | 0.011 | 0.004 | (0.006) | 1.657 | 0.013 | (0.473) |
| Latvia | 5 | 0.012 | 0.004 | (0.005) | 0.325 | -0.302 | (0.437) |
| Croatia | 4 | 0.012 | 0.003 | (0.006) | 1.422 | -0.209 | (0.499) |
| Hungary | 5 | 0.013 | 0.005 | (0.006) | 1.213 | 0.209 | (0.499) |
| Indonesia | 5 | 0.013 | -0.002 | (0.006) | 0.424 | -0.787 | (0.485) |
| Sweden | 5 | 0.014 | 0.007 | (0.005) | 0.740 | 0.111 | (0.453) |
| Spain | 5 | 0.014 | 0.009 | (0.005) | 0.786 | 0.468 | (0.476) |
| Belgium-Flemish | 5 | 0.014 | 0.001 | (0.005) | 0.826 | -0.907 | (0.485) |
| Korea, Rep. | 5 | 0.015 | 0.013 | (0.006) | 1.925 | 1.200 | (0.503) |
| Denmark | 5 | 0.016 | 0.008 | (0.005) | 0.511 | 0.129 | (0.450) |
| Iceland | 5 | 0.016 | 0.011 | (0.005) | 1.329 | 0.702 | (0.469) |
| Italy | 5 | 0.019 | 0.012 | (0.005) | 1.365 | 0.602 | (0.446) |
| New Zealand | 5 | 0.020 | 0.006 | (0.005) | 1.343 | -0.115 | (0.493) |
| Israel | 4 | 0.021 | 0.002 | (0.006) | 1.752 | -0.669 | (0.521) |
| Luxembourg | 5 | 0.023 | 0.008 | (0.005) | 1.768 | 0.033 | (0.470) |
| Macao-China | 5 | 0.024 | 0.016 | (0.005) | 1.792 | 0.953 | (0.459) |

Note: Counties sorted by OLS cohort trend in gaps using standardized achievement. OLS and shrunken EB (empirical Bayes) cohort trends are derived from Model 2 in Table B1 (unstandardized achievement) and Table B3 (standardized achievement). Trends are shown only for countries with at least 4 available study-years.

Table B6. Estimated Change in 90/10 SES Achievement Gaps, PISA Reading 2000-2015 (1985-2000 Birth Cohorts)

| Country | N (study-years) | Standardized Achievement | | | Unstandardized Achievement | | |
|---|---|---|---|---|---|---|---|
| | | OLS cohort trend | EB cohort trend | (SE) | OLS cohort trend | EB cohort trend | (SE) |
| Jordan | 4 | -0.026 | -0.008 | (0.006) | -2.469 | -1.320 | (0.772) |
| Mexico | 6 | -0.023 | -0.014 | (0.005) | -2.440 | -1.950 | (0.544) |
| Argentina | 4 | -0.021 | -0.005 | (0.006) | -3.797 | -1.540 | (0.728) |
| Chile | 5 | -0.017 | -0.008 | (0.005) | -2.358 | -1.550 | (0.617) |
| Scotland | 6 | -0.014 | -0.004 | (0.005) | -1.654 | -0.917 | (0.546) |
| Brazil | 6 | -0.013 | -0.008 | (0.005) | -1.297 | -1.110 | (0.563) |
| Germany | 6 | -0.012 | -0.005 | (0.005) | -2.927 | -1.700 | (0.599) |
| Slovenia | 4 | -0.011 | 0.000 | (0.006) | -0.798 | -0.141 | (0.778) |
| England | 6 | -0.008 | -0.003 | (0.004) | -1.093 | -0.649 | (0.503) |
| Chinese Taipei | 4 | -0.007 | 0.004 | (0.006) | 0.572 | 0.828 | (0.728) |
| United States | 5 | -0.007 | -0.001 | (0.005) | -1.197 | -0.526 | (0.534) |
| Tunisia | 5 | -0.007 | -0.003 | (0.006) | -1.123 | -0.907 | (0.678) |
| Poland | 6 | -0.006 | -0.003 | (0.005) | -1.653 | -1.070 | (0.546) |
| Switzerland | 6 | -0.006 | -0.002 | (0.005) | -0.965 | -0.568 | (0.570) |
| Netherlands | 6 | -0.003 | 0.001 | (0.005) | 0.687 | 0.531 | (0.545) |
| Indonesia | 6 | -0.003 | -0.003 | (0.006) | -0.226 | -0.427 | (0.578) |
| Estonia | 4 | -0.002 | -0.001 | (0.006) | -0.468 | -0.554 | (0.736) |
| Thailand | 6 | -0.001 | -0.002 | (0.006) | 0.118 | -0.151 | (0.628) |
| Belgium-French | 6 | -0.001 | -0.002 | (0.005) | -1.376 | -1.160 | (0.581) |
| Australia | 6 | -0.001 | 0.000 | (0.004) | -0.129 | -0.088 | (0.492) |
| Serbia | 4 | 0.001 | 0.002 | (0.006) | 1.565 | 0.494 | (0.724) |
| Turkey | 5 | 0.001 | 0.005 | (0.006) | -0.805 | -0.067 | (0.670) |
| Czech Rep. | 6 | 0.001 | 0.002 | (0.005) | -0.204 | -0.091 | (0.604) |
| Lithuania | 4 | 0.002 | 0.003 | (0.006) | -0.249 | 0.067 | (0.747) |
| Russian Fed. | 6 | 0.003 | 0.003 | (0.005) | 0.162 | 0.171 | (0.545) |
| Peru | 4 | 0.003 | -0.002 | (0.005) | -0.211 | -0.534 | (0.585) |
| Liechtenstein | 5 | 0.003 | 0.000 | (0.007) | -0.845 | -0.535 | (0.829) |
| Korea, Rep. | 6 | 0.003 | 0.006 | (0.005) | 1.218 | 1.120 | (0.523) |
| Austria | 6 | 0.004 | 0.005 | (0.005) | 0.482 | 0.518 | (0.580) |
| Sweden | 6 | 0.004 | 0.004 | (0.005) | 1.100 | 0.699 | (0.577) |
| Qatar | 4 | 0.005 | -0.001 | (0.006) | 0.428 | -0.443 | (0.806) |
| Canada | 6 | 0.005 | 0.005 | (0.005) | 0.322 | 0.393 | (0.539) |
| Norway | 6 | 0.005 | 0.005 | (0.005) | 0.006 | 0.181 | (0.567) |
| Israel | 5 | 0.005 | 0.001 | (0.006) | 0.557 | -0.087 | (0.691) |
| Hong Kong | 6 | 0.006 | 0.007 | (0.005) | 0.398 | 0.570 | (0.550) |
| France | 6 | 0.007 | 0.005 | (0.005) | 2.238 | 1.600 | (0.528) |
| Hungary | 6 | 0.007 | 0.005 | (0.005) | 0.812 | 0.618 | (0.581) |
| Spain | 6 | 0.007 | 0.006 | (0.004) | 0.512 | 0.481 | (0.492) |
| Luxembourg | 6 | 0.008 | 0.005 | (0.005) | 1.322 | 0.801 | (0.630) |
| Portugal | 6 | 0.010 | 0.006 | (0.005) | 0.247 | 0.099 | (0.537) |
| Japan | 6 | 0.011 | 0.009 | (0.005) | 1.036 | 0.928 | (0.563) |
| Slovak Rep. | 5 | 0.011 | 0.006 | (0.006) | 2.651 | 1.210 | (0.712) |
| New Zealand | 6 | 0.012 | 0.006 | (0.005) | 0.931 | 0.440 | (0.642) |
| Croatia | 4 | 0.012 | 0.005 | (0.006) | 0.963 | 0.358 | (0.727) |
| Belgium-Flemish | 6 | 0.012 | 0.004 | (0.005) | 1.030 | 0.330 | (0.614) |
| Greece | 6 | 0.012 | 0.007 | (0.005) | 1.017 | 0.558 | (0.530) |
| Finland | 6 | 0.013 | 0.010 | (0.005) | 1.591 | 1.240 | (0.527) |
| Iceland | 6 | 0.013 | 0.010 | (0.005) | 1.238 | 1.040 | (0.579) |
| Ireland | 6 | 0.013 | 0.010 | (0.004) | 0.531 | 0.542 | (0.499) |
| Colombia | 4 | 0.014 | 0.006 | (0.006) | 0.078 | 0.084 | (0.729) |
| Uruguay | 5 | 0.014 | 0.005 | (0.006) | -0.627 | -0.497 | (0.669) |
| Bulgaria | 5 | 0.015 | 0.006 | (0.006) | 2.684 | 1.250 | (0.690) |
| Denmark | 6 | 0.016 | 0.011 | (0.005) | 0.762 | 0.618 | (0.520) |
| Latvia | 6 | 0.016 | 0.008 | (0.005) | 0.739 | 0.349 | (0.542) |
| Italy | 6 | 0.021 | 0.013 | (0.005) | 1.516 | 1.130 | (0.550) |
| Montenegro | 4 | 0.021 | 0.007 | (0.006) | 1.389 | 0.419 | (0.740) |
| Macao-China | 5 | 0.022 | 0.012 | (0.006) | 1.987 | 1.300 | (0.646) |
| Romania | 5 | 0.028 | 0.011 | (0.006) | 1.731 | 0.730 | (0.707) |

Note: Counties sorted by OLS cohort trend in gaps using standardized achievement. OLS and shrunken EB (empirical Bayes) cohort trends are derived from Model 2 in Table B1 (unstandardized achievement) and Table B3 (standardized achievement). Trends are shown only for countries with at least 4 available study-years.

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Table B7. Estimated Change in 90/10 SES Achievement Gaps, PISA Science 2006-2015 (1991-2000 Birth Cohorts)

| Country | N (study-years) | Standardized Achievement | | | Unstandardized Achievement | | |
|---|---|---|---|---|---|---|---|
| | | OLS cohort trend | EB cohort trend | (SE) | OLS cohort trend | EB cohort trend | (SE) |
| Mexico | 4 | -0.032 | -0.008 | (0.001) | -3.451 | -1.530 | (0.608) |
| Netherlands | 4 | -0.020 | -0.001 | (0.002) | -1.240 | -0.496 | (0.619) |
| United States | 4 | -0.019 | 0.002 | (0.001) | -2.886 | -0.628 | (0.593) |
| Scotland | 4 | -0.018 | 0.004 | (0.001) | -2.491 | -0.448 | (0.610) |
| Jordan | 4 | -0.017 | -0.004 | (0.002) | -1.981 | -0.921 | (0.613) |
| Iceland | 4 | -0.016 | 0.003 | (0.002) | -1.953 | -0.250 | (0.667) |
| Chile | 4 | -0.015 | -0.004 | (0.002) | -1.558 | -0.670 | (0.647) |
| Germany | 4 | -0.014 | -0.002 | (0.002) | -1.790 | -0.608 | (0.615) |
| Thailand | 4 | -0.014 | -0.004 | (0.002) | -0.966 | -0.658 | (0.646) |
| Brazil | 4 | -0.014 | -0.005 | (0.002) | -1.431 | -0.837 | (0.632) |
| Uruguay | 4 | -0.012 | -0.003 | (0.002) | -1.978 | -0.687 | (0.624) |
| Poland | 4 | -0.009 | 0.000 | (0.002) | -0.415 | -0.050 | (0.629) |
| Lithuania | 4 | -0.008 | -0.001 | (0.001) | -0.415 | -0.244 | (0.597) |
| Slovenia | 4 | -0.007 | -0.001 | (0.002) | -1.064 | -0.397 | (0.627) |
| Austria | 4 | -0.007 | 0.001 | (0.002) | -1.094 | -0.105 | (0.653) |
| Bulgaria | 4 | -0.006 | -0.004 | (0.002) | -1.550 | -0.500 | (0.690) |
| England | 4 | -0.006 | 0.002 | (0.001) | -1.354 | -0.413 | (0.612) |
| Estonia | 4 | -0.005 | -0.002 | (0.002) | -0.308 | -0.599 | (0.610) |
| Croatia | 4 | -0.004 | -0.001 | (0.002) | 0.296 | -0.134 | (0.595) |
| Spain | 4 | -0.003 | 0.001 | (0.002) | -0.472 | -0.144 | (0.627) |
| France | 4 | -0.003 | 0.000 | (0.001) | -0.365 | -0.085 | (0.631) |
| Tunisia | 4 | -0.002 | -0.006 | (0.002) | -1.863 | -1.090 | (0.639) |
| Turkey | 4 | -0.001 | 0.005 | (0.002) | -0.461 | 0.045 | (0.622) |
| Sweden | 4 | -0.001 | -0.001 | (0.002) | 0.730 | -0.116 | (0.633) |
| Australia | 4 | -0.001 | 0.002 | (0.002) | 0.099 | 0.104 | (0.599) |
| Hong Kong | 4 | 0.000 | 0.006 | (0.001) | -0.868 | 0.014 | (0.588) |
| Colombia | 4 | 0.000 | 0.000 | (0.002) | -0.194 | -0.251 | (0.596) |
| Latvia | 4 | 0.000 | -0.003 | (0.002) | 0.062 | -0.523 | (0.594) |
| Qatar | 4 | 0.001 | -0.005 | (0.001) | 1.014 | -0.337 | (0.620) |
| Norway | 4 | 0.001 | 0.002 | (0.002) | 0.412 | 0.018 | (0.641) |
| Denmark | 4 | 0.001 | 0.000 | (0.002) | -0.209 | -0.159 | (0.619) |
| Switzerland | 4 | 0.001 | -0.001 | (0.002) | 0.172 | -0.106 | (0.641) |
| Portugal | 4 | 0.002 | -0.003 | (0.002) | 0.637 | -0.229 | (0.587) |
| Romania | 4 | 0.003 | 0.000 | (0.002) | -0.342 | -0.321 | (0.638) |
| Greece | 4 | 0.003 | -0.003 | (0.002) | 0.095 | -0.339 | (0.612) |
| Luxembourg | 4 | 0.005 | -0.003 | (0.002) | 0.751 | -0.080 | (0.645) |
| Japan | 4 | 0.005 | 0.006 | (0.002) | -0.133 | 0.405 | (0.614) |
| New Zealand | 4 | 0.005 | -0.003 | (0.002) | -0.036 | -0.438 | (0.651) |
| Canada | 4 | 0.006 | 0.004 | (0.002) | 0.391 | 0.186 | (0.614) |
| Indonesia | 4 | 0.006 | -0.004 | (0.002) | 0.085 | -0.412 | (0.598) |
| Belgium-French | 4 | 0.006 | -0.004 | (0.002) | -0.849 | -0.808 | (0.617) |
| Slovak Rep. | 4 | 0.006 | 0.000 | (0.002) | 1.901 | 0.284 | (0.637) |
| Montenegro | 4 | 0.008 | -0.001 | (0.001) | 0.664 | -0.154 | (0.592) |
| Chinese Taipei | 4 | 0.009 | 0.005 | (0.002) | 1.045 | 0.786 | (0.599) |
| Czech Rep. | 4 | 0.011 | -0.001 | (0.002) | 0.458 | -0.141 | (0.653) |
| Russian Fed. | 4 | 0.012 | 0.000 | (0.001) | 0.190 | -0.123 | (0.582) |
| Ireland | 4 | 0.013 | 0.004 | (0.002) | 0.444 | 0.324 | (0.604) |
| Macao-China | 4 | 0.014 | 0.007 | (0.001) | 1.033 | 0.411 | (0.587) |
| Finland | 4 | 0.015 | 0.003 | (0.002) | 2.442 | 0.609 | (0.635) |
| Belgium-Flemish | 4 | 0.015 | -0.007 | (0.001) | 2.737 | 0.002 | (0.615) |
| Korea, Rep. | 4 | 0.018 | 0.007 | (0.002) | 1.899 | 0.825 | (0.604) |
| Israel | 4 | 0.020 | -0.004 | (0.002) | 1.784 | -0.064 | (0.625) |
| Italy | 4 | 0.021 | 0.003 | (0.001) | 1.494 | 0.481 | (0.593) |
| Hungary | 4 | 0.024 | 0.000 | (0.002) | 3.309 | 0.779 | (0.627) |

Note: Counties sorted by OLS cohort trend in gaps using standardized achievement. OLS and shrunken EB (empirical Bayes) cohort trends are derived from Model 2 in Table B1 (unstandardized achievement) and Table B3 (standardized achievement). Trends are shown only for countries with at least 4 available study-years.

Table B8. Estimated Change in 90/10 SES Achievement Gaps, PIRLS Reading 2001-2011 (1991-2001 Birth Cohorts)

| Country | N (study-years) | Standardized Achievement | | | Unstandardized Achievement | | |
|---|---|---|---|---|---|---|---|
| | | OLS cohort trend | EB cohort trend | (SE) | OLS cohort trend | EB cohort trend | (SE) |
| Singapore | 3 | -0.018 | 0.007 | (0.004) | -2.736 | -0.694 | (0.808) |
| Iran | 3 | -0.018 | 0.005 | (0.005) | -2.327 | -0.652 | (0.843) |
| Netherlands | 3 | -0.012 | 0.006 | (0.005) | -0.853 | -0.283 | (0.741) |
| Germany | 3 | -0.006 | 0.002 | (0.005) | -0.263 | -0.189 | (0.796) |
| Norway | 3 | -0.002 | 0.005 | (0.005) | -1.997 | -1.090 | (0.764) |
| Slovenia | 3 | 0.005 | 0.012 | (0.005) | -0.184 | 0.502 | (0.782) |
| Italy | 3 | 0.005 | 0.008 | (0.005) | 0.232 | 0.340 | (0.757) |
| Hong Kong | 3 | 0.006 | 0.010 | (0.005) | 0.476 | 0.469 | (0.709) |
| Hungary | 3 | 0.007 | 0.008 | (0.005) | 2.453 | 1.430 | (0.850) |
| France | 3 | 0.007 | 0.007 | (0.005) | 0.173 | 0.295 | (0.773) |
| England | 3 | 0.009 | 0.005 | (0.005) | 0.266 | 0.051 | (0.993) |
| Bulgaria | 3 | 0.013 | 0.010 | (0.005) | 1.002 | 0.764 | (0.833) |
| Slovak Rep. | 3 | 0.013 | 0.011 | (0.005) | 0.553 | 0.922 | (0.833) |
| Sweden | 3 | 0.014 | 0.006 | (0.005) | 0.725 | 0.232 | (0.787) |
| Kuwait | 3 | 0.018 | 0.019 | (0.005) | 3.356 | 2.390 | (0.863) |
| New Zealand | 3 | 0.020 | 0.003 | (0.005) | 1.190 | 0.047 | (0.891) |
| Israel | 3 | 0.021 | 0.019 | (0.005) | 0.911 | 1.660 | (0.836) |
| Lithuania | 3 | 0.022 | 0.009 | (0.005) | 1.697 | 0.924 | (0.753) |
| Russian Fed. | 3 | 0.042 | 0.014 | (0.005) | 2.711 | 1.520 | (0.747) |
| Romania | 3 | 0.053 | 0.009 | (0.005) | 4.896 | 1.990 | (0.852) |

Note: Counties sorted by OLS cohort trend in gaps using standardized achievement. OLS and shrunken EB (empirical Bayes) cohort trends are derived from Model 2 in Table B1 (unstandardized achievement) and Table B3 (standardized achievement). Trends are shown only for countries with at least 3 available study-years.

Table B9. Estimated Change in 90/10 SES Achievement Gaps, TIMSS Grade 4 Math 1995-2015 (1985-2005 Birth Cohorts)

| Country | N (study-years) | Standardized Achievement | | | Unstandardized Achievement | | |
|---|---|---|---|---|---|---|---|
| | | OLS cohort trend | EB cohort trend | (SE) | OLS cohort trend | EB cohort trend | (SE) |
| New Zealand | 5 | -0.005 | 0.000 | (0.009) | -0.076 | 0.319 | (0.742) |
| Netherlands | 5 | -0.004 | 0.003 | (0.011) | -0.678 | -0.236 | (0.723) |
| United States | 5 | -0.003 | 0.002 | (0.010) | 0.120 | 0.415 | (0.795) |
| Chinese Taipei | 4 | -0.003 | 0.001 | (0.008) | 0.799 | 0.786 | (0.696) |
| Singapore | 5 | 0.002 | 0.003 | (0.005) | -0.228 | -0.055 | (0.514) |
| Slovenia | 5 | 0.009 | 0.009 | (0.008) | -0.334 | -0.043 | (0.633) |
| Morocco | 4 | 0.009 | 0.014 | (0.014) | 1.189 | 0.636 | (1.075) |
| England | 5 | 0.014 | 0.013 | (0.009) | 0.720 | 0.804 | (0.769) |
| Czech Rep. | 4 | 0.015 | 0.013 | (0.012) | 0.313 | 0.595 | (0.915) |
| Norway | 5 | 0.015 | 0.014 | (0.009) | 0.301 | 0.448 | (0.715) |
| Russian Fed. | 4 | 0.018 | 0.018 | (0.008) | 0.886 | 0.813 | (0.705) |
| Lithuania | 4 | 0.021 | 0.017 | (0.011) | 1.224 | 0.996 | (0.872) |
| Australia | 5 | 0.023 | 0.018 | (0.011) | 1.856 | 1.380 | (0.847) |
| Iran | 5 | 0.026 | 0.020 | (0.009) | 2.902 | 2.060 | (0.767) |
| Japan | 4 | 0.027 | 0.023 | (0.008) | 1.189 | 1.080 | (0.676) |
| Hungary | 5 | 0.033 | 0.029 | (0.007) | 3.884 | 3.160 | (0.628) |
| Hong Kong | 5 | 0.036 | 0.030 | (0.009) | 2.113 | 1.690 | (0.652) |
| Italy | 4 | 0.064 | 0.050 | (0.009) | 3.696 | 2.620 | (0.725) |

Note: Counties sorted by OLS cohort trend in gaps using standardized achievement. OLS and shrunken EB (empirical Bayes) cohort trends are derived from Model 2 in Table B2 (unstandardized achievement) and Table B4 (standardized achievement). Trends are shown only for countries with at least 4 available study-years.

Table B10. Estimated Change in 90/10 SES Achievement Gaps, TIMSS Grade 4 Science 1995-2015 (1985-2015 Birth Cohorts)

| Country | N (study-years) | Standardized Achievement | | | Unstandardized Achievement | | |
|---|---|---|---|---|---|---|---|
| | | OLS cohort trend | EB cohort trend | (SE) | OLS cohort trend | EB cohort trend | (SE) |
| Netherlands | 5 | -0.013 | 0.000 | (0.011) | -0.422 | 0.146 | (0.762) |
| Singapore | 5 | -0.004 | 0.000 | (0.009) | -1.110 | -0.739 | (0.682) |
| Slovenia | 5 | -0.002 | 0.005 | (0.011) | -0.209 | 0.186 | (0.778) |
| United States | 5 | 0.004 | 0.008 | (0.011) | 0.019 | 0.210 | (0.862) |
| Chinese Taipei | 4 | 0.008 | 0.012 | (0.011) | 0.502 | 0.614 | (0.874) |
| New Zealand | 5 | 0.013 | 0.013 | (0.011) | 0.673 | 0.500 | (0.869) |
| Czech Rep. | 4 | 0.015 | 0.014 | (0.012) | 0.469 | 0.539 | (0.848) |
| Australia | 5 | 0.019 | 0.017 | (0.012) | 0.767 | 0.711 | (0.945) |
| Iran | 5 | 0.022 | 0.018 | (0.010) | 3.005 | 1.640 | (0.792) |
| England | 5 | 0.023 | 0.017 | (0.011) | -0.062 | 0.138 | (0.881) |
| Morocco | 4 | 0.026 | 0.022 | (0.014) | 2.809 | 1.190 | (1.090) |
| Norway | 5 | 0.026 | 0.021 | (0.011) | 0.258 | 0.501 | (0.852) |
| Hong Kong | 5 | 0.032 | 0.027 | (0.010) | 2.170 | 1.810 | (0.738) |
| Hungary | 5 | 0.042 | 0.030 | (0.010) | 3.700 | 2.250 | (0.774) |
| Russian Fed. | 4 | 0.043 | 0.029 | (0.012) | 1.931 | 1.390 | (0.906) |
| Lithuania | 4 | 0.045 | 0.025 | (0.013) | 2.887 | 1.570 | (0.974) |
| Japan | 4 | 0.046 | 0.028 | (0.012) | 1.950 | 1.270 | (0.897) |
| Italy | 4 | 0.073 | 0.043 | (0.012) | 3.437 | 2.060 | (0.887) |

Note: Counties sorted by OLS cohort trend in gaps using standardized achievement. OLS and shrunken EB (empirical Bayes) cohort trends are derived from Model 2 in Table B2 (unstandardized achievement) and Table B4 (standardized achievement). Trends are shown only for countries with at least 4 available study-years.

Table B11. Estimated Change in 90/10 SES Achievement Gaps, TIMSS Grade 8 Math 1995-2015 (1981-2001 Birth Cohorts)

| Country | N (study-years) | Standardized Achievement | | | Unstandardized Achievement | | |
|---|---|---|---|---|---|---|---|
| | | OLS cohort trend | EB cohort trend | (SE) | OLS cohort trend | EB cohort trend | (SE) |
| Chile | 4 | -0.024 | -0.006 | (0.009) | -2.150 | -0.355 | (0.841) |
| Russian Fed. | 6 | -0.018 | -0.009 | (0.007) | -1.083 | -0.407 | (0.729) |
| Jordan | 5 | -0.011 | -0.003 | (0.008) | -1.137 | -0.043 | (0.822) |
| Lebanon | 4 | -0.008 | 0.003 | (0.010) | 0.069 | 0.602 | (0.900) |
| Cyprus | 4 | -0.004 | 0.006 | (0.010) | -1.122 | 0.338 | (1.077) |
| Botswana | 4 | -0.003 | 0.003 | (0.011) | 0.780 | 0.354 | (1.022) |
| Saudi Arabia | 4 | -0.002 | 0.006 | (0.011) | 1.744 | 1.140 | (0.965) |
| Slovenia | 6 | -0.002 | 0.003 | (0.007) | -0.772 | -0.280 | (0.638) |
| Malaysia | 5 | 0.000 | 0.006 | (0.009) | 0.767 | 0.922 | (0.822) |
| Korea, Rep. | 6 | 0.002 | 0.008 | (0.008) | 0.551 | 0.859 | (0.786) |
| United States | 6 | 0.003 | 0.007 | (0.007) | -0.039 | 0.429 | (0.695) |
| Indonesia | 4 | 0.003 | 0.006 | (0.011) | -0.784 | 0.209 | (1.084) |
| Lithuania | 6 | 0.005 | 0.007 | (0.007) | 0.285 | 0.670 | (0.695) |
| Chinese Taipei | 5 | 0.005 | 0.011 | (0.009) | 0.355 | 1.300 | (0.896) |
| Sweden | 5 | 0.012 | 0.014 | (0.008) | 0.800 | 0.853 | (0.758) |
| Italy | 6 | 0.016 | 0.015 | (0.008) | 0.103 | 0.446 | (0.721) |
| Japan | 5 | 0.017 | 0.015 | (0.009) | 2.211 | 1.960 | (0.847) |
| South Africa | 5 | 0.018 | 0.016 | (0.008) | 0.449 | 0.874 | (0.813) |
| Hungary | 6 | 0.018 | 0.016 | (0.009) | 2.577 | 2.380 | (0.863) |
| Singapore | 6 | 0.019 | 0.018 | (0.007) | 2.857 | 2.470 | (0.687) |
| Bahrain | 4 | 0.020 | 0.015 | (0.009) | 2.229 | 1.600 | (0.906) |
| Norway | 5 | 0.020 | 0.018 | (0.008) | 0.806 | 0.792 | (0.741) |
| Australia | 6 | 0.021 | 0.018 | (0.008) | 1.677 | 1.650 | (0.738) |
| New Zealand | 5 | 0.024 | 0.022 | (0.008) | 2.113 | 2.010 | (0.742) |
| Thailand | 5 | 0.024 | 0.019 | (0.009) | 2.169 | 1.870 | (0.828) |
| Iran | 6 | 0.024 | 0.021 | (0.008) | 3.290 | 2.710 | (0.754) |
| Hong Kong | 6 | 0.027 | 0.024 | (0.007) | 2.184 | 1.840 | (0.679) |
| Israel | 6 | 0.027 | 0.022 | (0.007) | 3.715 | 3.150 | (0.734) |
| England | 6 | 0.029 | 0.024 | (0.008) | 2.375 | 1.930 | (0.802) |
| Tunisia | 4 | 0.030 | 0.018 | (0.011) | 2.757 | 1.750 | (0.980) |
| Canada | 4 | 0.033 | 0.026 | (0.007) | 2.108 | 1.750 | (0.714) |
| Morocco | 5 | 0.034 | 0.021 | (0.009) | 2.576 | 1.570 | (0.889) |
| Romania | 5 | 0.046 | 0.033 | (0.009) | 5.213 | 3.830 | (0.890) |
| Turkey | 4 | 0.052 | 0.033 | (0.009) | 6.999 | 4.580 | (0.927) |

Note: Counties sorted by OLS cohort trend in gaps using standardized achievement. OLS and shrunken EB (empirical Bayes) cohort trends are derived from Model 2 in Table B2 (unstandardized achievement) and Table B4 (standardized achievement). Trends are shown only for countries with at least 4 available study-years.

Table B12. Estimated Change in 90/10 SES Achievement Gaps, TIMSS Grade 8 Science 1995-2015 (1982-2001 Birth Cohorts)

| Country | N (study-years) | Standardized Achievement | | | Unstandardized Achievement | | |
|---|---|---|---|---|---|---|---|
| | | OLS cohort trend | EB cohort trend | (SE) | OLS cohort trend | EB cohort trend | (SE) |
| Lebanon | 4 | -0.017 | 0.007 | (0.008) | -0.607 | 1.170 | (0.931) |
| Cyprus | 4 | -0.017 | 0.006 | (0.008) | -1.981 | 0.405 | (0.954) |
| Botswana | 4 | -0.015 | 0.008 | (0.009) | 0.308 | 1.100 | (0.982) |
| Chile | 4 | -0.013 | 0.004 | (0.008) | -1.578 | 0.050 | (0.803) |
| Jordan | 5 | -0.010 | 0.005 | (0.008) | -0.600 | 0.575 | (0.832) |
| Russian Fed. | 6 | 0.000 | 0.007 | (0.007) | -0.482 | 0.323 | (0.741) |
| Korea, Rep. | 6 | 0.001 | 0.011 | (0.007) | 0.013 | 0.504 | (0.747) |
| Malaysia | 5 | 0.002 | 0.010 | (0.007) | 2.537 | 1.970 | (0.791) |
| United States | 6 | 0.004 | 0.011 | (0.007) | -0.469 | 0.272 | (0.722) |
| Slovenia | 6 | 0.004 | 0.009 | (0.006) | 0.423 | 0.747 | (0.666) |
| Saudi Arabia | 4 | 0.005 | 0.011 | (0.008) | 2.578 | 1.710 | (0.917) |
| Indonesia | 4 | 0.006 | 0.011 | (0.009) | 0.052 | 1.150 | (0.962) |
| Singapore | 6 | 0.008 | 0.011 | (0.006) | 1.732 | 1.640 | (0.672) |
| Sweden | 5 | 0.012 | 0.016 | (0.008) | 1.904 | 1.550 | (0.785) |
| Chinese Taipei | 5 | 0.012 | 0.015 | (0.008) | 1.014 | 1.220 | (0.840) |
| Japan | 5 | 0.014 | 0.015 | (0.008) | 1.383 | 1.340 | (0.829) |
| South Africa | 5 | 0.017 | 0.015 | (0.007) | 0.705 | 1.270 | (0.815) |
| Hungary | 6 | 0.017 | 0.014 | (0.008) | 1.983 | 1.720 | (0.801) |
| Lithuania | 6 | 0.019 | 0.016 | (0.007) | 1.419 | 1.440 | (0.712) |
| Canada | 4 | 0.019 | 0.018 | (0.007) | 0.937 | 1.100 | (0.720) |
| Italy | 6 | 0.020 | 0.018 | (0.007) | 0.809 | 0.997 | (0.699) |
| Israel | 6 | 0.020 | 0.016 | (0.007) | 2.496 | 2.160 | (0.725) |
| New Zealand | 5 | 0.021 | 0.019 | (0.007) | 2.141 | 1.840 | (0.742) |
| Tunisia | 4 | 0.023 | 0.016 | (0.009) | 1.479 | 1.470 | (0.917) |
| Australia | 6 | 0.024 | 0.019 | (0.007) | 1.530 | 1.520 | (0.744) |
| Hong Kong | 6 | 0.024 | 0.020 | (0.007) | 1.690 | 1.540 | (0.689) |
| Norway | 5 | 0.025 | 0.020 | (0.007) | 1.997 | 1.670 | (0.768) |
| Morocco | 5 | 0.025 | 0.017 | (0.008) | 2.092 | 1.620 | (0.860) |
| England | 6 | 0.027 | 0.020 | (0.008) | 1.716 | 1.390 | (0.796) |
| Thailand | 5 | 0.027 | 0.019 | (0.008) | 2.723 | 2.220 | (0.747) |
| Bahrain | 4 | 0.034 | 0.018 | (0.008) | 5.645 | 2.930 | (0.911) |
| Iran | 6 | 0.038 | 0.026 | (0.007) | 4.099 | 3.110 | (0.705) |
| Turkey | 4 | 0.047 | 0.024 | (0.008) | 5.975 | 3.490 | (0.864) |
| Romania | 5 | 0.048 | 0.027 | (0.008) | 2.854 | 2.060 | (0.838) |

Note: Counties sorted by OLS cohort trend in gaps using standardized achievement. OLS and shrunken EB (empirical Bayes) cohort trends are derived from Model 2 in Table B2 (unstandardized achievement) and Table B4 (standardized achievement). Trends are shown only for countries with at least 4 available study-years.

Country sampling practices in international large-scale assessments have generally improved over time, resulting in greater target population coverage in more recent study-years. This may tend to bias gap trend estimates upward, to the extent that early samples with poor coverage are positively selected, resulting in smaller gaps in older cohorts. As population coverage information is inconsistently provided in early studies, I retain all available data in all analyses in the main paper and throughout these appendices, as I do not wish to possibly bias results by inappropriately excluding data.

However, Tables B14 and B15 report analyses to test the robustness of results to this decision. These tables show gap trends for each TIMSS test instrument, after dropping 10 country-study-years in TIMSS 4$^{th}$ grade and 31 country-study-years in TIMSS 8$^{th}$ grade that the IEA considers not comparable for trend analysis. This is usually due to improved translations or increased population coverage. The list of non-comparable country-study-years is drawn from Mullis et al. (2016) Appendix A.1. Comparing to models for the full sample of TIMSS country-study-years in Tables B2 and B4, results are similar, with large and significant positive cohort coefficients. If gap trends are biased upward by including early studies with poor population coverage, we would expect cohort coefficients to be smaller in Tables B13 and B14 than in Tables B2 and B4. Indeed, cohort coefficients are about 20-50% smaller for TIMSS 8$^{th}$ grade. However, coefficients are about 8-20% *larger* for TIMSS 4$^{th}$ grade, implying that excluding early years with poor population coverage may bias trends *downward*.

Thus, I conclude that, while trends reported in the main analyses of this paper are likely biased by the inclusion of some country-study-years with poor population coverage, the exact magnitude or direction of bias is difficult to ascertain. The average global increase in the 90/10 SES achievement gap reported in the main results is 0.4 SD over 55 cohort years. The gap trend changes in Tables B13 and B14 imply that the true gap increase may be between 0.2 and 0.48 SD. The most conservative lower bound of this range, 0.2 SD, is still a substantial increase in inequality. As it is difficult to apply a single exclusion rule consistently across the full sample of 1026 country-study-years due to inconsistencies in documentation—particularly in early years—I choose to retain all available data in all other analyses.

Table B13. Coefficients from Hierarchical Growth Models Predicting Achievement Gaps between 90th and 10th Percentiles of SES, Run Separately by Test Instrument (Unstandardized Achievement) – TIMSS, dropping country-study-years that IEA considers not comparable for trend analysis

| Test instrument | PISA Math | | PISA Reading | | PISA Science | | PIRLS Reading | |
|---|---|---|---|---|---|---|---|---|
| Test years | 2003-2015 | | 2000-2015 | | 2006-2015 | | 2001-2011 | |
| Cohort birth years | 1988-2000 | | 1985-2000 | | 1991-2000 | | 1991-2001 | |
| | (1) | (2) | (1) | (2) | (1) | (2) | (1) | (2) |
| Parent education gaps intercept | | | | | 76.235 *** | 76.606 *** | 76.146 *** | 74.211 *** |
| | | | | | (3.646) | (3.712) | (4.203) | (4.116) |
| Household books gaps intercept | 73.555 *** | 73.555 *** | 79.516 *** | 79.516 *** | 88.331 *** | 87.939 *** | 86.46 *** | 88.503 *** |
| | (5.008) | (5.008) | (6.018) | (6.018) | (4.237) | (4.276) | (4.986) | (5.017) |
| Cohort birth year × Parent education | | | | | 0.945 *** | | 0.95 *** | |
| | | | | | (0.253) | | (0.260) | |
| Cohort birth year × Books | 0.841 ** | | 0.778 * | | 0.929 ** | | 1.308 *** | |
| | (0.272) | | (0.318) | | (0.311) | | (0.275) | |
| Cohort birth year | | 0.841 ** | | 0.778 * | | 0.946 *** | | 1.13 *** |
| | | (0.272) | | (0.318) | | (0.256) | | (0.241) |
| **Random effects** | | | | | | | | |
| *Level 2 - Residual variance between studies in…* | | | | | | | | |
| Parent education intercepts | | | | | 115.756 | 108.252 | 141.517 | 137.663 |
| Books intercepts | 37.166 | 37.166 | 100.862 | 100.862 | 104.775 | 121.067 | 139.097 | 148.133 |
| *Level 3 - Residual variance between countries in…* | | | | | | | | |
| Parent education intercepts | | | | | 347.558 | 431.017 | 529.285 | 564.589 |
| Books intercepts | 617.864 | 617.864 | 863.261 | 863.261 | 595.754 | 708.006 | 938.432 | 1051.900 |
| Parent education cohort slopes | | | | | 1.088 | | 0.924 | |
| Books cohort slopes | 1.305 | | 1.409 | | 2.768 | | 1.312 | |
| Cohort slopes | | 1.305 | | 1.409 | | 1.730 | | 1.032 |
| % Parent education cohort slopes > 0 | | | | | 0.818 | | 0.839 | |
| % Books cohort slopes > 0 | 0.769 | | 0.744 | | 0.712 | | 0.873 | |
| % Cohort slopes > 0 | | 0.769 | | 0.744 | | 0.764 | | 0.867 |
| N (Level 1 - gaps) | 153 | 153 | 153 | 153 | 423 | 423 | 423 | 423 |
| N (Level 2 - study-years) | 153 | 153 | 153 | 153 | 214 | 214 | 214 | 214 |
| N (Level 3 - countries) | 47 | 47 | 47 | 47 | 59 | 59 | 59 | 59 |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: To avoid very long computation times, all models in this table specify known level 1 error variances estimated using conventional non-bootstrap formulas and omit error covariances. This simplified specification appears to produce very similar results to models using bootstrapped error variances and covariances (see Appendix L). "% cohort slopes > 0" is computed by comparing cohort coefficient estimate and cohort slope variance component to a normal distribution. See Mullis et al. (2016) Appendix A.1 for list of excluded country-study-years.

Table B14. Coefficients from Hierarchical Growth Models Predicting Achievement Gaps between 90th and 10th Percentiles of SES, Run Separately by Test Instrument (Standardized Achievement) – TIMSS, dropping country-study-years that IEA considers not comparable for trend analysis

| Test instrument | TIMSS Grade 4 Math | | TIMSS Grade 4 Science | | TIMSS Grade 8 Math | | TIMSS Grade 8 Science | |
|---|---|---|---|---|---|---|---|---|
| Test years | 1995-2015 | | 1995-2015 | | 1995-2015 | | 1995-2015 | |
| Cohort birth years | 1985-2005 | | 1985-2005 | | 1981-2001 | | 1981-2001 | |
| | (1) | (2) | (1) | (2) | (1) | (2) | (1) | (2) |
| Parent education gaps intercept | | | | | 1.003 *** | 1.020 *** | 1.016 *** | 0.996 *** |
| | | | | | (0.044) | (0.043) | (0.047) | (0.044) |
| Household books gaps intercept | 0.985 *** | 0.985 *** | 1.017 *** | 1.017 *** | 1.168 *** | 1.160 *** | 1.171 *** | 1.191 *** |
| | (0.070) | (0.070) | (0.081) | (0.081) | (0.047) | (0.048) | (0.058) | (0.058) |
| Cohort birth year × Parent education | | | | | 0.008 ** | | 0.008 ** | |
| | | | | | (0.003) | | (0.003) | |
| Cohort birth year × Books | 0.014 *** | | 0.018 *** | | 0.007 * | | 0.012 *** | |
| | (0.003) | | (0.004) | | (0.003) | | (0.003) | |
| Cohort birth year | | 0.014 *** | | 0.018 *** | | 0.007 ** | | 0.010 *** |
| | | (0.003) | | (0.004) | | (0.003) | | (0.003) |
| **Random effects** | | | | | | | | |
| *Level 2 - Residual variance between studies in…* | | | | | | | | |
| Parent education intercepts | | | | | 0.01257 | 0.01236 | 0.01247 | 0.01185 |
| Books intercepts | 0.00334 | 0.00334 | 0.01678 | 0.01678 | 0.00839 | 0.00981 | 0.01205 | 0.01295 |
| *Level 3 - Residual variance between countries in…* | | | | | | | | |
| Parent education intercepts | | | | | 0.05320 | 0.06072 | 0.06415 | 0.06454 |
| Books intercepts | 0.13752 | 0.13752 | 0.16607 | 0.16607 | 0.07529 | 0.09180 | 0.13271 | 0.15171 |
| Parent education cohort slopes | | | | | 0.00008 | | 0.00006 | |
| Books cohort slopes | 0.00021 | | 0.00019 | | 0.00036 | | 0.00024 | |
| Cohort slopes | | 0.00021 | | 0.00019 | | 0.00020 | | 0.00014 |
| % Parent education cohort slopes > 0 | | | | | 0.828 | | 0.834 | |
| % Books cohort slopes > 0 | 0.840 | | 0.904 | | 0.644 | | 0.771 | |
| % Cohort slopes > 0 | | 0.840 | | 0.904 | | 0.704 | | 0.789 |
| N (Level 1 - gaps) | 153 | 153 | 153 | 153 | 423 | 423 | 423 | 423 |
| N (Level 2 - study-years) | 153 | 153 | 153 | 153 | 214 | 214 | 214 | 214 |
| N (Level 3 - countries) | 47 | 47 | 47 | 47 | 59 | 59 | 59 | 59 |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: To avoid very long computation times, all models in this table specify known level 1 error variances estimated using conventional non-bootstrap formulas and omit error covariances. This simplified specification appears to produce very similar results to models using bootstrapped error variances and covariances (see Appendix L). "% cohort slopes > 0" is computed by comparing cohort coefficient estimate and cohort slope variance component to a normal distribution. See Mullis et al. (2016) Appendix A.1 for list of excluded country-study-years.

Tables B16 and B17 show a comparison of gap trend estimates from the models above and those from Broer et al. (in press). There are a number of methodological differences between the analyses reported in the paper and those in Broer et al. (in press). First, they classify trends as "increasing" or "decreasing" based only on the contrast between 1995 and 2015 gaps, rather than estimating a linear trend line. The column labeled "Broer diff./20" reproduces those reported differences, dividing by 20 to yield the estimated annual change in gaps. The column labeled "Broer coef." contains estimates that I computed by regressing gap estimates from each wave reported by Broer et al. on test year. Third, they do not standardize achievement within waves. The columns labeled "Unstandardized Achievement" also keep scores in the original TIMSS scale. Fourth, they exclude two country-study-years that the IEA does not consider comparable for trend analysis: TIMSS 1999 for Australia and Slovenia. These two country-study-years are omitted from all trends reported in Tables B16 and B17. They also exclude TIMSS 2015 Lithuanian students who were assessed in languages other than Lithuanian. In Tables B16 and B17, I exclude the entire TIMSS 2015 Lithuanian sample from Broer et al.'s and my estimates. After making all the adjustments above, the trend estimates in the "Broer coef." column and the "Unstandardized Achievement—OLS cohort trend" column are quite similar. They are similar in magnitude and have a correlation of 0.87 for both math and science. The remaining small discrepancies between these two columns are due to the different measurement of SES and computation of achievement gaps. Broer et al. compute gaps between the top and bottom quartile of an SES index composed from parent education, books, and two other household possessions (a computer and a desk for studying).

Table B16. Comparing Estimated Change in SES Achievement Gaps in Broer et al. and Current Study, TIMSS Grade 8 Math 1995-2015 (1982-2001 Birth Cohorts)

| Country | N (study-years) | Broer diff./20[1] | Broer coef.[2] | Unstandardized Achievement | | | Standardized Achievement | | |
|---|---|---|---|---|---|---|---|---|---|
| | | | | OLS cohort trend | EB cohort trend | (SE) | OLS cohort trend | EB cohort trend | (SE) |
| Australia | 5 | -1.205 | -0.284 | 1.255 | 1.420 | (0.805) | 0.016 | 0.015 | (0.009) |
| Hong Kong | 6 | 0.802 | 1.357 | 2.199 | 1.840 | (0.684) | 0.027 | 0.024 | (0.007) |
| Hungary | 6 | 3.403 | 1.721 | 2.584 | 2.390 | (0.864) | 0.018 | 0.015 | (0.009) |
| Iran | 6 | 5.108 | 3.414 | 3.285 | 2.690 | (0.756) | 0.024 | 0.020 | (0.008) |
| Korea, Rep. | 6 | -0.856 | 0.271 | 0.549 | 0.861 | (0.788) | 0.002 | 0.008 | (0.008) |
| Lithuania | 5 | -0.063 | -0.325 | 0.100 | 0.687 | (0.862) | 0.002 | 0.006 | (0.009) |
| New Zealand | 5 | 1.212 | 0.363 | 2.119 | 2.010 | (0.745) | 0.024 | 0.022 | (0.008) |
| Norway | 5 | -1.102 | -0.432 | 0.803 | 0.789 | (0.745) | 0.020 | 0.018 | (0.008) |
| Russian Fed. | 6 | -1.338 | -0.571 | -1.099 | -0.402 | (0.732) | -0.018 | -0.009 | (0.007) |
| Singapore | 6 | 3.125 | 1.986 | 2.865 | 2.460 | (0.691) | 0.019 | 0.018 | (0.007) |
| Slovenia | 5 | -1.183 | -0.422 | -0.719 | -0.163 | (0.680) | -0.005 | 0.001 | (0.007) |
| Sweden | 5 | -0.116 | -0.193 | 0.801 | 0.852 | (0.762) | 0.012 | 0.014 | (0.008) |
| United States | 6 | -0.871 | -0.679 | -0.034 | 0.443 | (0.700) | 0.003 | 0.007 | (0.007) |

[1] Broer diff./20 is derived from difference estimates reported Broer et al. (in press) Table 4.1., divided by 20. Lithuanian difference is divided by 16 (2011-1995).
[2] I computed Broer coefficients by regressing gap estimates reported in Broer et al. (in press) Table 4.2 on test year.
Note: OLS and shrunken EB (empirical Bayes) cohort trends are derived from Model 2 in Table B2 (unstandardized achievement) and Table B4 (standardized achievement), after omitting TIMSS 1999 for Australia and Slovenia and TIMSS 2015 for Lithuania.

Table B17. Comparing Estimated Change in SES Achievement Gaps in Broer et al. and Current Study, TIMSS Grade 8 Science 1995-2015 (1982-2001 Birth Cohorts)

| | | | | Unstandardized Achievement | | | Standardized Achievement | | |
|---|---|---|---|---|---|---|---|---|---|
| Country | N (study-years) | Broer diff./20[1] | Broer coef.[2] | OLS cohort trend | EB cohort trend | (SE) | OLS cohort trend | EB cohort trend | (SE) |
| Australia | 5 | -1.123 | -0.226 | 1.398 | 1.460 | (0.803) | 0.021 | 0.018 | (0.008) |
| Hong Kong | 6 | -0.288 | 0.607 | 1.690 | 1.530 | (0.694) | 0.024 | 0.020 | (0.007) |
| Hungary | 6 | 2.562 | 1.571 | 1.984 | 1.720 | (0.805) | 0.017 | 0.014 | (0.008) |
| Iran | 6 | 5.308 | 3.314 | 4.099 | 3.100 | (0.710) | 0.038 | 0.026 | (0.007) |
| Korea, Rep. | 6 | -0.288 | 0.057 | 0.012 | 0.500 | (0.751) | 0.001 | 0.010 | (0.007) |
| Lithuania | 5 | 0.313 | 0.150 | 1.123 | 1.270 | (0.837) | 0.018 | 0.015 | (0.008) |
| New Zealand | 5 | 1.093 | 0.291 | 2.137 | 1.830 | (0.746) | 0.021 | 0.019 | (0.007) |
| Norway | 5 | 0.538 | 0.730 | 1.997 | 1.660 | (0.774) | 0.025 | 0.020 | (0.008) |
| Russian Fed. | 6 | -1.074 | -0.400 | -0.482 | 0.325 | (0.746) | 0.000 | 0.007 | (0.007) |
| Singapore | 6 | 0.786 | 0.879 | 1.731 | 1.640 | (0.677) | 0.008 | 0.010 | (0.007) |
| Slovenia | 5 | 0.045 | 0.139 | 0.329 | 0.724 | (0.713) | -0.002 | 0.007 | (0.007) |
| Sweden | 5 | 0.702 | 0.530 | 1.902 | 1.540 | (0.790) | 0.012 | 0.015 | (0.008) |
| United States | 6 | -1.869 | -1.079 | -0.468 | 0.270 | (0.727) | 0.004 | 0.010 | (0.007) |

[1] Broer diff./20 is derived from difference estimates reported Broer et al. (in press) Table 4.1., divided by 20. Lithuanian difference is divided by 16 (2011-1995).

[2] I computed Broer coefficients by regressing gap estimates reported in Broer et al. (in press) Table 4.3 on test year.

Note: OLS and shrunken EB (empirical Bayes) cohort trends are derived from Model 2 in Table B2 (unstandardized achievement) and Table B4 (standardized achievement), after omitting TIMSS 1999 for Australia and Slovenia and TIMSS 2015 for Lithuania.

## C. Changing distribution of achievement

The models in the main text of the paper standardize achievement within each country-study-year before computing SES achievement gaps, both out of necessity since they combine different test instruments and also out of a theoretical preference for treating achievement as a positional good. Thus, trends in gaps are unaffected by changes in the variance of student achievement; they reflect changes in the relative *strength* of the SES-achievement association rather than the absolute *size* of the association. However, it is also interesting to ask whether the absolute size of the SES-achievement association is changing over time, as gaps in unstandardized achievement are assumed to reflect gaps in meaningful academic skills, rather than gaps in students' position within the achievement distribution.

In contrast to the main analyses that pool different studies, in separate analyses of trend studies (PISA, TIMSS, PIRLS), it is possible to estimate whether the variance of student achievement has changed over time. Figures C1-C8 display score variance at the student, school/classroom, and country levels for each year of each trend study, estimated from separate three-level hierarchical models, as follows:

$$\hat{T}_{ij} = \gamma_{00} + v_k + u_{jk} + \epsilon_{ijk},$$
$$v_k \sim N(0, \tau_{000}); u_{jk} \sim N(0, \tau_{00}); \ \epsilon_{ijk} \sim N(0, \sigma^2),$$

[C1]

where $\hat{T}_{ijk}$ is the estimated test score of student $i$ in school or classroom $j$ in country $k$, $\tau_{000}$ is the between-country variance of scores, $\tau_{00}$ is the between-school variance of scores, and $\sigma^2$ is the within-school student-level variance of scores. Total student weights are applied at the student level, meaning all students are weighted in proportion to their probability of selection, and all countries are weighted in proportion to the size of their target population (i.e., more populous countries receive greater weight). Models are estimated once for each plausible value and averaged. Only countries that participated in all years of a given trend study are included. Samples of included countries vary depending on the study. After estimating the student-, school/classroom-, and country-level variances, all three are adjusted for estimated test reliability ($\alpha$) for the relevant set of countries in each year, as follows:
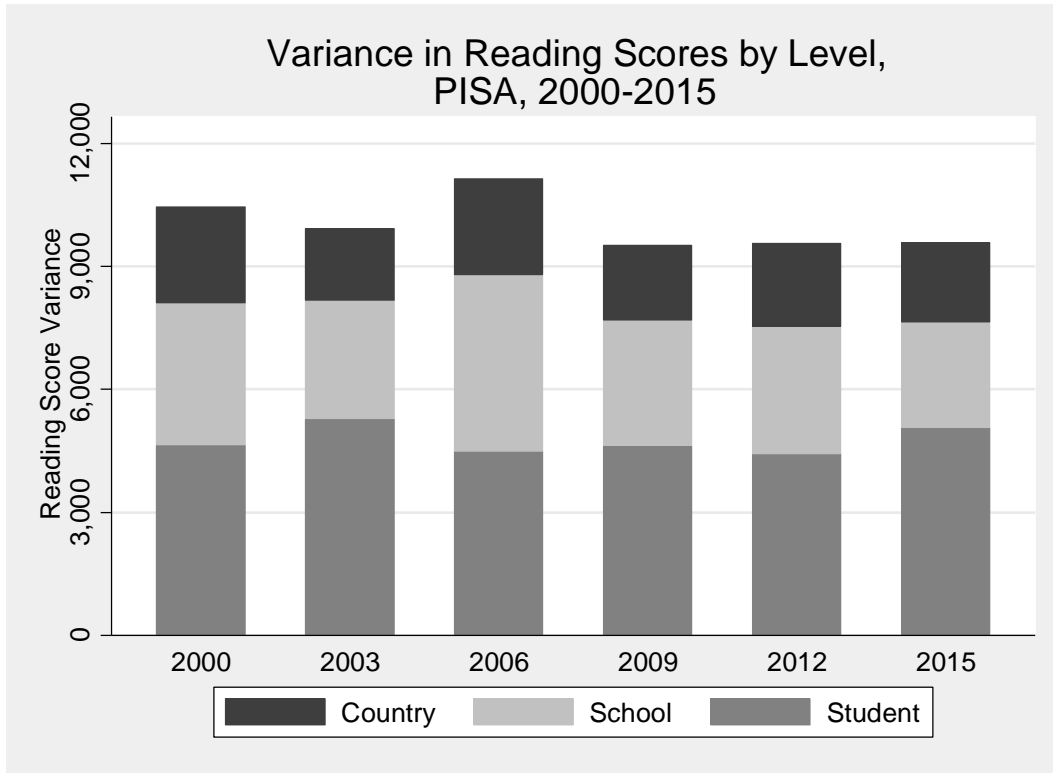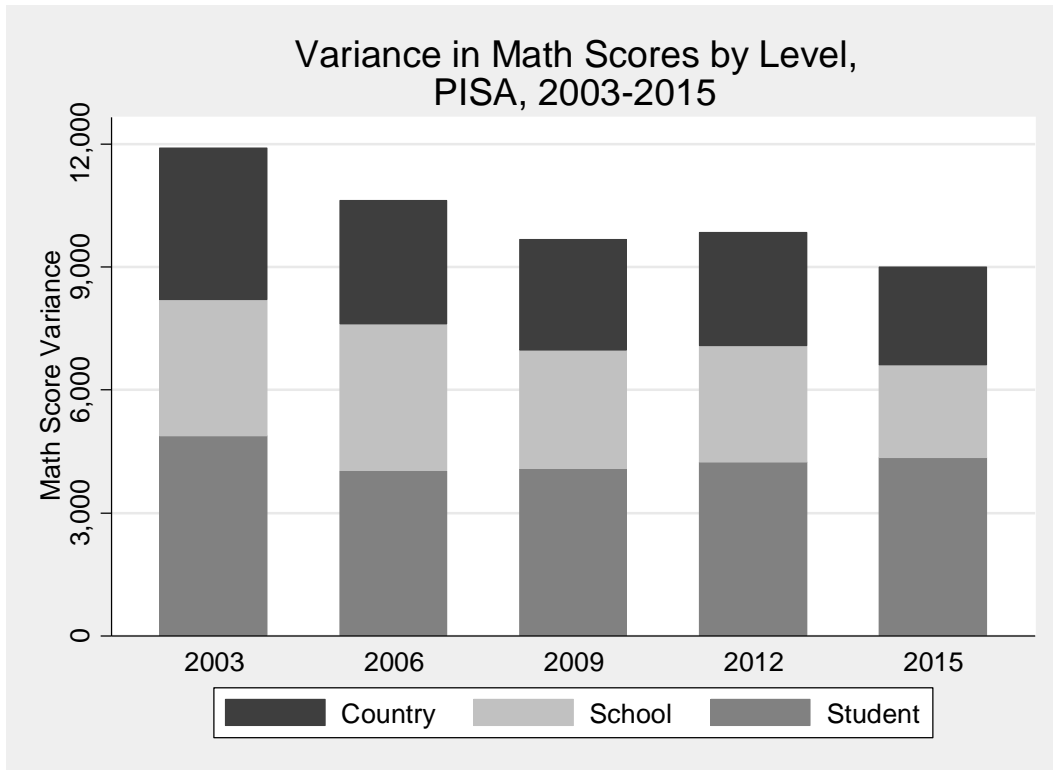
$$\sigma^2_{True} = \alpha * \sigma^2_{Total}$$

<div align="right">[C2]</div>

Figure C1

Variance in Reading Scores by Level,
PISA, 2000-2015



Figure C2

Variance in Math Scores by Level,
PISA, 2003-2015

Figure C3

Variance in Science Scores by Level,
PISA, 2006-2015



Figure C4

Variance in Math Scores by Level,
TIMSS Grade 4, 1995-2015

Figure C5



Variance in Science Scores by Level,
TIMSS Grade 4, 1995-2015

Figure C6



Variance in Math Scores by Level,
TIMSS Grade 8, 1995-2015

Figure C7



Variance in Science Scores by Level,
TIMSS Grade 8, 1995-2015

Figure C8



Variance in Reading Scores by Level,
PIRLS, 2001-2011
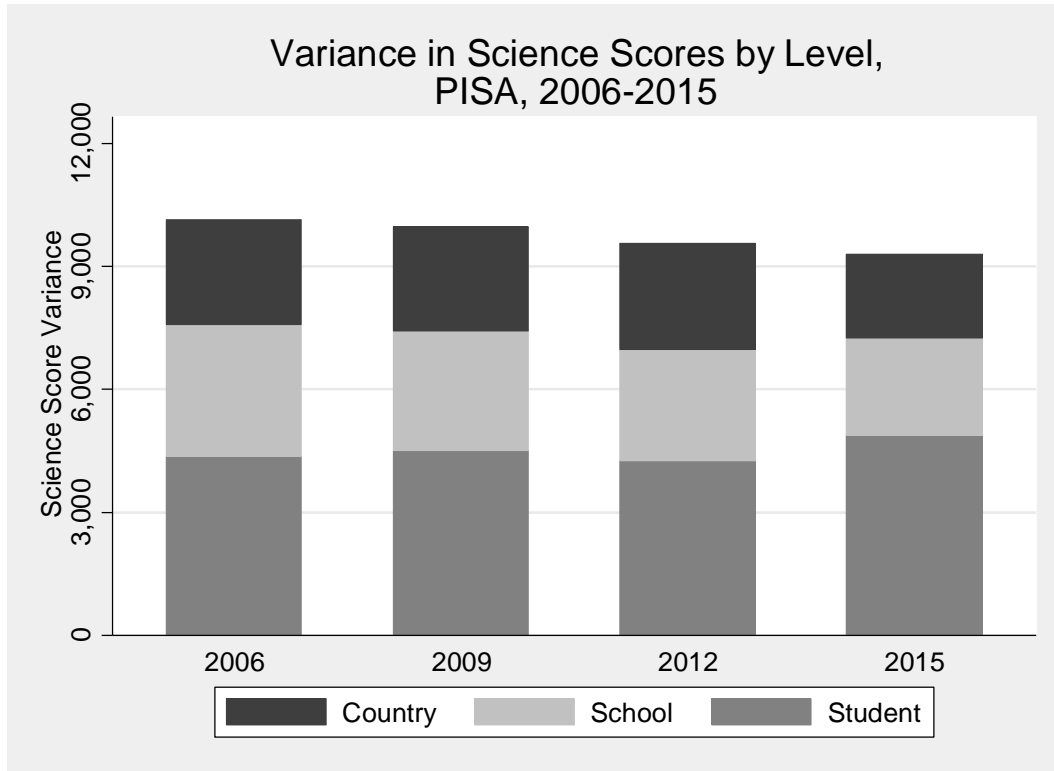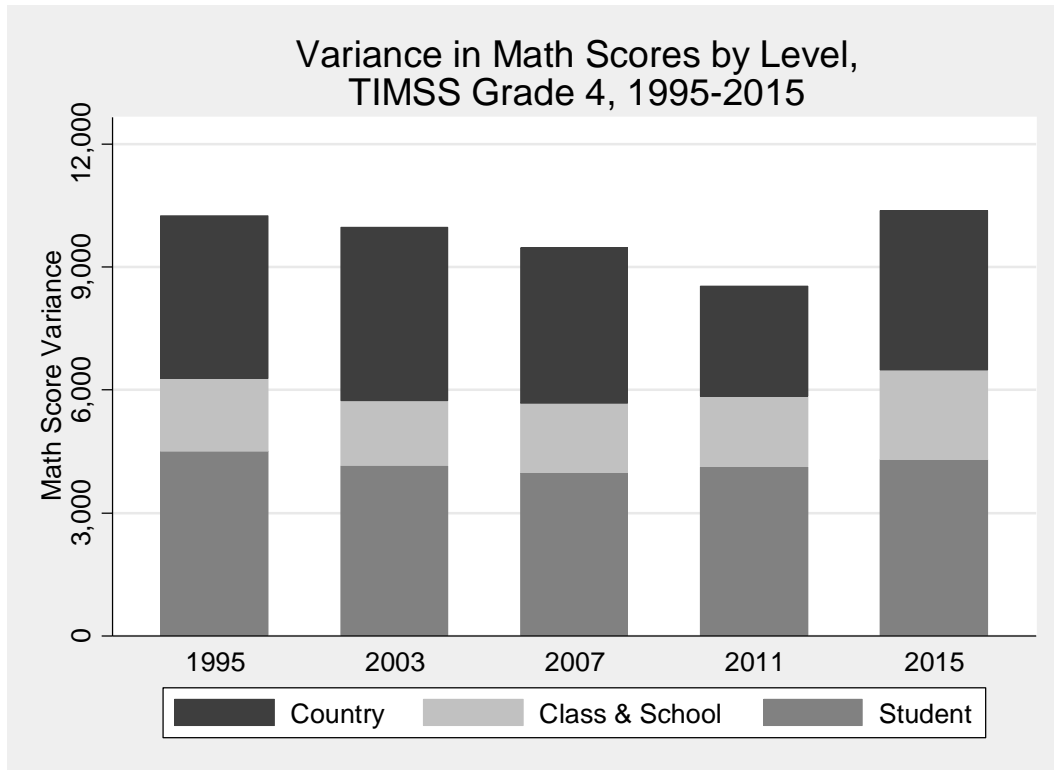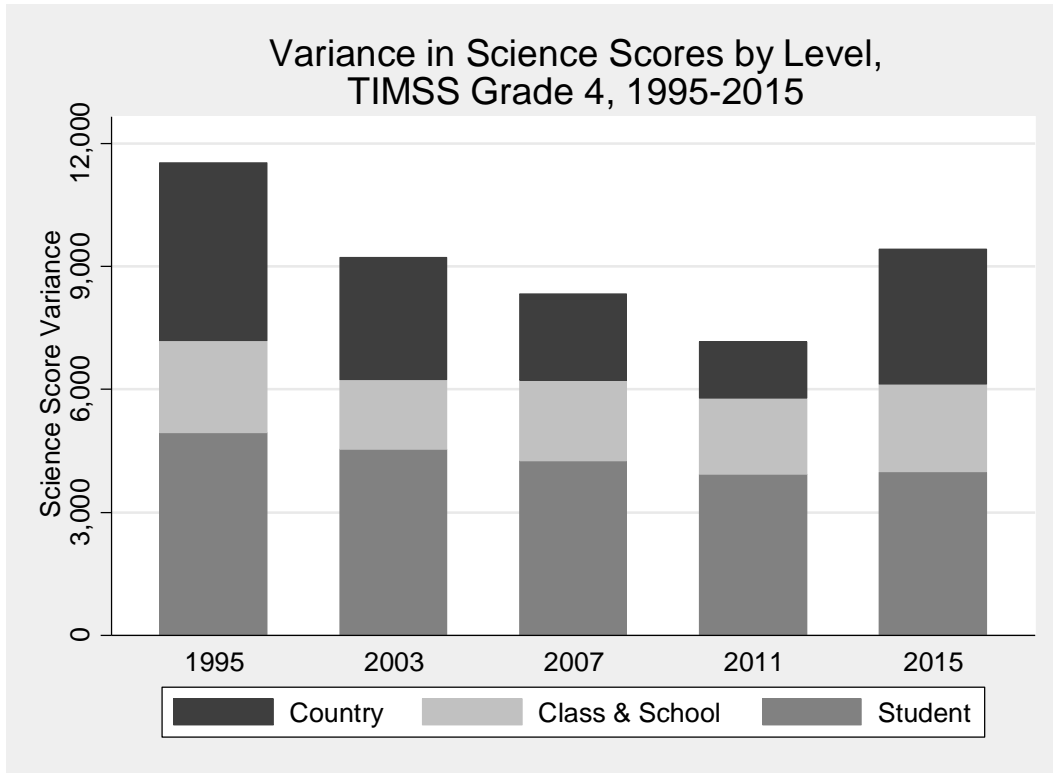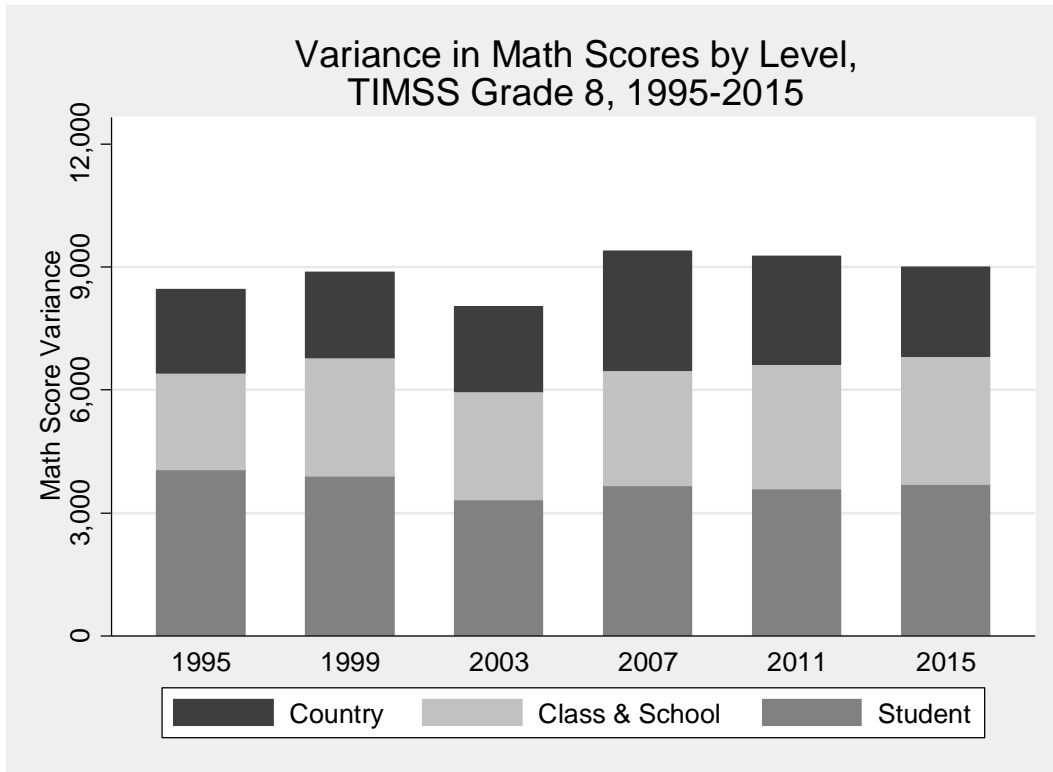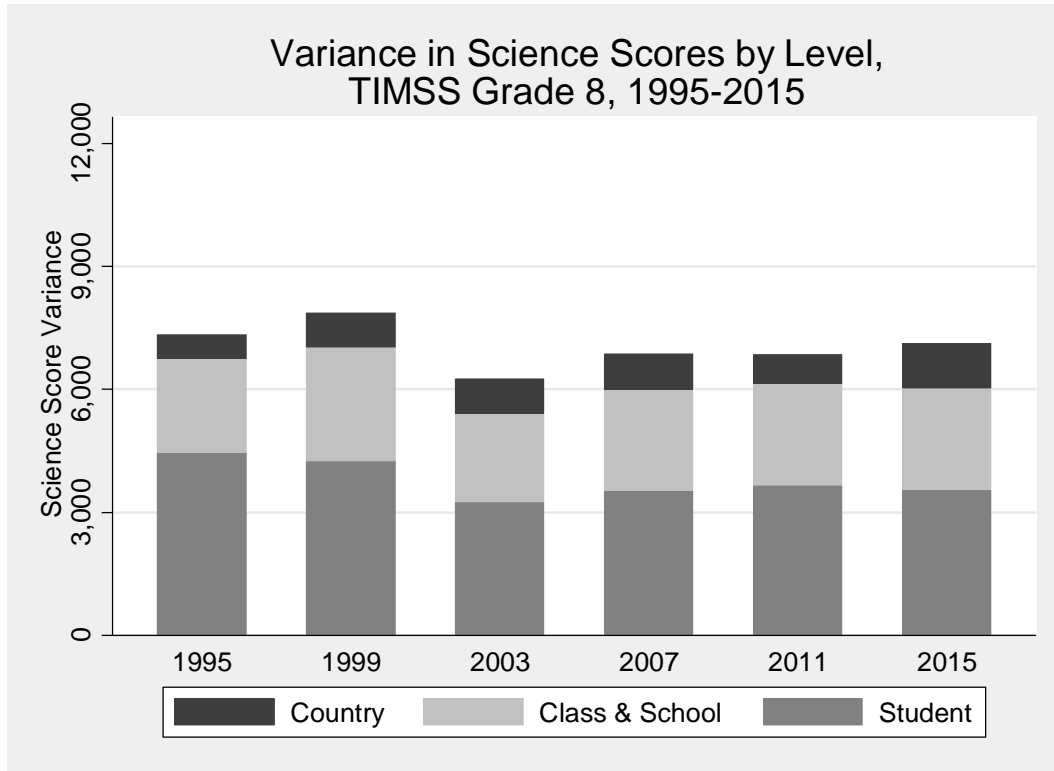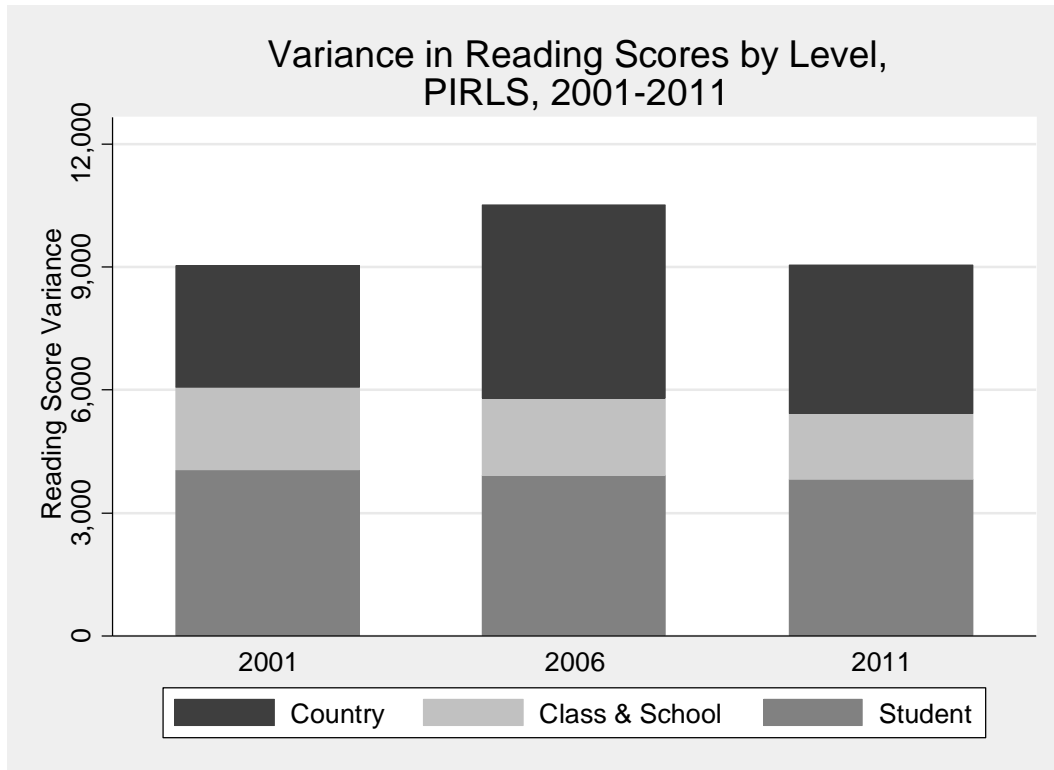
The hierarchal models reveal that within-country test score variance (the sum of the student- and school/classroom-level variances) has declined for all test instruments except TIMSS math (4th and 8th grades). These decreases in variance might lead us to question whether the absolute *size* of SES achievement gaps in terms of skills is declining even as their *strength* increases. However, the results for unstandardized achievement in Tables B1 and B2 show positive trends for most test instruments (except for PISA math and science). This indicates that, for the specific sets of reading, math, and science skills tested by the trend studies, differences in the degree to which high- and low-SES students have mastered those skills have mostly grown over the past 9-20 years.

The trend studies also allow us to examine not only whether achievement gaps have truly grown in size but also the changing *levels* of achievement for students of different SES. In other words, we can ask: Do SES achievement gaps increase because low-SES students' achievement is declining or because it is not rising as quickly as that of high-SES students? A series of models addressing this question, separately for each test instrument, are estimated as follows:

$$\hat{A}_{pjk} = \alpha(\mathbf{T}_{pjk}) + \gamma_{100}(P10_{pjk}) + \gamma_{010}(C_{jk}) + \gamma_{110}(C_{jk}P10_{pjk}) + \boldsymbol{u}_k(\mathbf{T}_{pjk}) + w_{1k}(C_{jk})$$
$$+ w_{2k}(C_{jk}P10_{pjk}) + \boldsymbol{r}_{jk}(\mathbf{T}_{pjk}) + s_{jk}(P10_{pjk}) + \epsilon_{pjk},$$

$$\epsilon_{pjk} \sim N(0, V_{jk}); \begin{bmatrix} r_{1jk} \\ r_{2jk} \\ r_{3jk} \\ s_{jk} \end{bmatrix} \sim MVN[\mathbf{0}, \boldsymbol{\Sigma}]; \begin{bmatrix} u_{1k} \\ u_{2k} \\ u_{3k} \\ w_{1k} \\ w_{2k} \end{bmatrix} \sim MVN[\mathbf{0}, \boldsymbol{\tau}]$$

[C3]

where $\hat{A}_{pjk}$ is the *p*th observed mean achievement of the 90th or the 10th SES percentile (level 1) in study-year *j* (level 2) in country *k* (level 3), $\boldsymbol{\alpha}$ is a vector of the true achievement means conditional on all covariates in the model, $\mathbf{T}_{pjk}$ is a vector of dummy variables indicating gap type (parent education, parent occupation, or household books), $\gamma_{100}$ is the coefficient for the dummy variable $P10_{ijk}$ indicating whether the mean achievement *p* was estimated for the 10th (1) or the 90th (0) SES percentile, $\gamma_{010}$ is the coefficient for cohort birth year $C_{jk}$, $\gamma_{110}$ is the coefficient for the interaction between cohort birth year $C_{jk}$ and the 10th percentile dummy $P10_{pjk}$, $\boldsymbol{\Sigma}$ and $\boldsymbol{\tau}$ are the within-country and between-country covariance matrices among the true gaps, and $V_{jk} = var(\hat{A}_{pjk})$ is the sampling variance of $\hat{A}_{pjk}$, which I compute using the squared standard error of $\hat{A}_{pjk}$.[3]

Table C1 reports these models for each test instrument. The coefficient for Cohort birth year represents the trend in achievement for the reference category, the 90th SES percentile. The interaction between Cohort birth year and the 10th percentile dummy represents the difference in

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[3] The level-1 known-variance portion of this model accounts only for the sampling variance of $\hat{A}_{pjk}$ and not for the sampling error covariances among different mean achievement estimates within the same country-study-year. Thus, the model assumes that all sampling error covariances among different mean achievement estimates are 0, which is unlikely to be true. However, the models in the main text yield very similar results when sampling error covariances are omitted (see Appendix L). Therefore, sampling error covariances are omitted from this and most appendix models due to the very long computation time of sampling error covariances via bootstrapping.

achievement trends for the two SES groups.[4] The sum of these two coefficients represents the trend in achievement for the 10th SES percentile, which is reported at the bottom of the table, along with significance of a Wald joint test of the hypothesis that the sum of coefficients is equal to 0. On average, the achievement of both high- and low-SES students has increased in all test instruments. However, in most test instruments, the achievement of low-SES students has increased by a smaller amount than that of high-SES students. The two exceptions are PISA math and science, where the achievement of low-SES students has increased more than that of high-SES students, consistent with declining SES achievement gaps for these test instruments when achievement is unstandardized, seen in Table B1. Thus, the results for the majority of test instruments show that SES achievement gaps are increasing not because low-SES students' achievement is declining, but because it is not rising as quickly as that of high-SES students.

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[4] Note that if the models in Tables B1 and C1 were simple OLS regression models, the coefficients for the Cohort birth year*p10 interactions in Table C1 would be equal to the Cohort birth year trends in the gaps in Table B1, but opposite in sign. This is not true in the reported models due to the precision weighting and random cohort slopes in the hierarchical growth models. However, estimated coefficients are generally similar in size and opposite in sign across Tables B1 and C1.

Table C1. Models Predicting Achievement Levels for 90<sup>th</sup> and 10<sup>th</sup> SES Percentiles, Run Separately by Test Instrument

| Test instrument | PISA Math | PISA Reading | PISA Science | PIRLS Reading | TIMSS Grade 4 Math | TIMSS Grade 4 Science | TIMSS Grade 8 Math | TIMSS Grade 8 Science |
|---|---|---|---|---|---|---|---|---|
| Test years | 2003-2015 | 2000-2015 | 2006-2015 | 2001-2011 | 1995-2015 | 1995-2015 | 1995-2015 | 1995-2015 |
| Cohort birth years | 1988-2000 | 1985-2000 | 1991-2000 | 1991-2001 | 1985-2005 | 1985-2005 | 1981-2001 | 1981-2001 |
| Parent education gaps intercept | 507.127 *** | 494.311 *** | 513.236 *** | 525.820 *** | | | 482.550 *** | 491.766 *** |
| | (8.761) | (7.744) | (8.354) | (10.582) | | | (11.904) | (10.809) |
| Parent occupation gaps intercept | 506.873 *** | 494.334 *** | 512.779 *** | 525.935 *** | | | | |
| | (8.675) | (7.665) | (8.252) | (10.573) | | | | |
| Household books gaps intercept | 505.730 *** | 492.764 *** | 511.313 *** | 525.379 *** | 493.917 *** | 497.314 *** | 480.153 *** | 489.396 *** |
| | (8.629) | (7.656) | (8.212) | (10.562) | (11.616) | (12.110) | (11.804) | (10.821) |
| p10 | -81.823 *** | -78.787 *** | -83.919 *** | -63.540 *** | -49.020 *** | -55.704 *** | -54.590 *** | -54.802 *** |
| | (2.872) | (2.622) | (3.334) | (2.800) | (3.684) | (4.240) | (3.073) | (3.563) |
| Cohort birth year | 0.167 | 0.841 *** | 0.018 | 1.836 *** | 1.444 *** | 1.194 *** | 1.297 *** | 1.317 *** |
| | (0.252) | (0.241) | (0.254) | (0.531) | (0.323) | (0.327) | (0.374) | (0.362) |
| Cohort birth year × p10 | 0.294 + | -0.023 | 0.377 + | -0.530 * | -0.448 * | -0.273 | -1.146 *** | -1.295 *** |
| | (0.166) | (0.158) | (0.201) | (0.248) | (0.209) | (0.232) | (0.198) | (0.185) |
| N (Level 1 - gaps) | 1786 | 2060 | 1528 | 602 | 340 | 340 | 970 | 970 |
| N (Level 2 - study-years) | 298 | 344 | 255 | 104 | 170 | 170 | 245 | 245 |
| N (Level 3 - countries) | 70 | 72 | 70 | 41 | 51 | 51 | 61 | 61 |
| | Wald test | Wald test | Wald test | Wald test | Wald test | Wald test | Wald test | Wald test |
| Estimated cohort birth year slope for p10 | 0.461 + | 0.818 ** | 0.395 | 1.306 * | 0.996 ** | 0.921 ** | 0.151 | 0.022 |
| | (0.282) | (0.252) | (0.261) | (0.509) | (0.338) | (0.337) | (0.360) | (0.350) |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: To avoid very long computation times, all models in this table specify known level 1 error variances estimated using conventional non-bootstrap formulas and omit error covariances. This simplified specification appears to produce very similar results to models using bootstrapped error variances and covariances (see Appendix L).

**D. Changing measurement error of achievement**

Even assuming the true variance of achievement had remained constant over time, if measurement error of achievement declines over time (e.g., because of improvements in testing methodology), the SES achievement gap estimates in the main text of the paper will artificially appear to increase because they were attenuated in early years where measurement error was higher. This is because the method of standardizing achievement in each country-year involves dividing by the standard deviation of achievement, which will be inflated due to measurement error. That SES achievement gaps are also increasing for most of the unstandardized scores reported in Appendix B is evidence that findings are robust, even when not standardizing achievement.

Tables D1-D3 report median, minimum, and maximum test reliabilities by age group for math, reading, and science tests. Median test reliabilities have not consistently increased over time for all test subjects and age groups. Reliabilities have increased for 4th grade tests and for secondary science tests, but appear to have declined somewhat for secondary math and reading tests. However, it should be kept in mind that the sample of countries participating in international assessments has become more diverse over time, and countries at a lower level of development often have lower test reliabilities.

Table D1. Median, Minimum, and Maximum Test Reliability for Math Tests

| Study | Year | Median | Minimum | Maximum |
|---|---|---|---|---|
| **Grade 4 Math** | | | | |
| TIMSS | 1995 | 0.84 | 0.74 | 0.88 |
| TIMSS | 2003 | 0.87 | 0.76 | 0.91 |
| TIMSS | 2007 | 0.83 | 0.55 | 0.88 |
| TIMSS | 2011 | 0.82 | 0.57 | 0.89 |
| TIMSS | 2015 | 0.88 | 0.78 | 0.92 |
| **Grade 8 Math** | | | | |
| FIMS | 1964 | 0.92 | 0.87 | 0.95 |
| SIMS[a] | 1980 | 0.85[a] | 0.81[a] | 0.85[a] |
| TIMSS | 1995 | 0.89 | 0.73 | 0.92 |
| TIMSS | 1999 | 0.89 | 0.69 | 0.94 |
| TIMSS | 2003 | 0.89 | 0.51 | 0.94 |
| TIMSS | 2007 | 0.88 | 0.62 | 0.93 |
| TIMSS | 2011 | 0.87 | 0.66 | 0.94 |
| TIMSS | 2015 | 0.91 | 0.81 | 0.94 |
| **Age 15 Math** | | | | |
| PISA | 2000 | 0.88 | 0.82 | 0.92 |
| PISA | 2003 | 0.90 | 0.83 | 0.93 |
| PISA | 2006 | 0.88 | 0.83 | 0.93 |
| PISA | 2009 | 0.88 | 0.77 | 0.92 |
| PISA | 2012 | 0.92 | 0.84 | 0.94 |
| PISA | 2015 | 0.85 | 0.67 | 0.89 |

[a] SIMS test reliability was not reported in the available documentation and was estimated for each country using a model that included age, subject, year, and countries' level of development.

Table D2. Median, Minimum, and Maximum Test Reliability for Reading Tests

| Study | Year | Median | Minimum | Maximum |
|-------|------|--------|---------|---------|
| **Grade 4 Reading** | | | | |
| FIRCS | 1970 | 0.85 | 0.74 | 0.89 |
| RLS | 1991 | 0.93 | 0.89 | 0.97 |
| PIRLS | 2001 | 0.88 | 0.83 | 0.91 |
| PIRLS | 2006 | 0.87 | 0.81 | 0.92 |
| PIRLS | 2011 | 0.88 | 0.79 | 0.93 |
| **Grade 8 Reading** | | | | |
| FIRCS | 1970 | 0.85 | 0.64 | 0.90 |
| RLS | 1991 | 0.92 | 0.77 | 0.95 |
| **Age 15 Reading** | | | | |
| PISA | 2000 | 0.92 | 0.87 | 0.94 |
| PISA | 2003 | 0.83 | 0.70 | 0.88 |
| PISA | 2006 | 0.88 | 0.80 | 0.93 |
| PISA | 2009 | 0.92 | 0.86 | 0.94 |
| PISA | 2012 | 0.89 | 0.81 | 0.93 |
| PISA | 2015 | 0.86 | 0.72 | 0.89 |

Table D3. Median, Minimum, and Maximum Test Reliability for Science Tests

| Study | Year | Median | Minimum | Maximum |
|-------|------|--------|---------|---------|
| **Grade 4 Science** | | | | |
| FISS | 1970 | 0.82 | 0.68 | 0.87 |
| SISS | 1984 | 0.74 | 0.70 | 0.79 |
| TIMSS | 1995 | 0.77 | 0.70 | 0.83 |
| TIMSS | 2003 | 0.84 | 0.74 | 0.87 |
| TIMSS | 2007 | 0.80 | 0.69 | 0.88 |
| TIMSS | 2011 | 0.78 | 0.62 | 0.85 |
| TIMSS | 2015 | 0.85 | 0.81 | 0.90 |
| **Grade 8 Science** | | | | |
| FISS | 1970 | 0.83 | 0.57 | 0.89 |
| SISS | 1984 | 0.75 | 0.60 | 0.80 |
| TIMSS | 1995 | 0.78 | 0.69 | 0.84 |
| TIMSS | 1999 | 0.80 | 0.62 | 0.86 |
| TIMSS | 2003 | 0.84 | 0.63 | 0.91 |
| TIMSS | 2007 | 0.84 | 0.65 | 0.91 |
| TIMSS | 2011 | 0.83 | 0.67 | 0.89 |
| TIMSS | 2015 | 0.89 | 0.81 | 0.92 |
| **Age 15 Science** | | | | |
| PISA | 2000 | 0.87 | 0.75 | 0.92 |
| PISA | 2003 | 0.82 | 0.68 | 0.88 |
| PISA | 2006 | 0.91 | 0.84 | 0.94 |
| PISA | 2009 | 0.89 | 0.79 | 0.93 |
| PISA | 2012 | 0.89 | 0.80 | 0.93 |
| PISA | 2015 | 0.91 | 0.77 | 0.93 |

For the models in the main text of the paper, which pool different tests with different scales and must standardize achievement, all SES achievement gaps (and their standard errors) are adjusted according to each country's test reliability for each study, as published in the corresponding technical reports (as well as for the estimated reliability of SES reports, which is explained in more detail in Appendix H). The adjustment is computed as follows:

$$\widehat{Gap}_{adj} = \widehat{Gap}_{raw} * \frac{1}{\sqrt{\alpha_{ach}} * \sqrt{\alpha_{SES}}}$$

[D1]

Table D4 reports estimated trends in SES achievement gaps without adjusting those gaps for differences in test reliability (and also without adjusting for the reliability of SES report, which is discussed in more detail in Appendix H). Gap trends are positive and significant, and are very similar to those reported in the models in the main text. Without adjusting for reliability, the increase in parent occupation gaps is nearly identical to that reported in the main text, and increases in parent education and books gaps are slightly smaller. Gaps adjusted for reliability are preferred, as they are likely more accurate.

Table D4. Estimated trends in 90/10 SES achievement gaps, without adjusting for test or SES reliability

| | Adjusted for reliability | | No adjustment | |
|---|---|---|---|---|
| | (1) | (2) | (1) | (2) |
| Parent education gaps intercept | 1.032 *** | 1.039 *** | 0.741 *** | 0.744 *** |
| | (0.030) | (0.030) | (0.022) | (0.022) |
| Parent occupation gaps intercept | 0.958 *** | 0.964 *** | 0.775 *** | 0.776 *** |
| | (0.030) | (0.030) | (0.024) | (0.024) |
| Household books gaps intercept | 1.299 *** | 1.294 *** | 0.851 *** | 0.850 *** |
| | (0.041) | (0.041) | (0.028) | (0.028) |
| **Level 1 - Gaps** | | | | |
| Subject (ref=Reading): | | | | |
| Math | 0.020 ** | 0.020 ** | 0.017 ** | 0.017 ** |
| | (0.007) | (0.007) | (0.006) | (0.006) |
| Science | 0.034 *** | 0.034 *** | 0.025 *** | 0.025 *** |
| | (0.005) | (0.005) | (0.004) | (0.004) |
| SES variable quality measures | | | | |
| Parent-reported × Parent education | 0.132 *** | 0.112 *** | 0.277 *** | 0.265 *** |
| | (0.030) | (0.031) | (0.025) | (0.026) |
| Parent-reported × Parent occupation | 0.075 ** | 0.073 ** | 0.082 *** | 0.089 *** |
| | (0.025) | (0.024) | (0.021) | (0.020) |
| Parent-reported × Books | -0.039 | -0.017 | 0.041 | 0.044 + |
| | (0.029) | (0.026) | (0.025) | (0.023) |
| Number of categories (centered at 7) | 0.003 | 0.002 | 0.004 | 0.004 + |
| | (0.003) | (0.003) | (0.002) | (0.002) |
| ≥ 20% in bottom category | -0.065 ** | -0.063 ** | -0.057 *** | -0.057 *** |
| | (0.021) | (0.021) | (0.015) | (0.015) |
| ≥ 20% in top category | -0.135 *** | -0.146 *** | -0.099 *** | -0.101 *** |
| | (0.013) | (0.013) | (0.009) | (0.009) |
| **Level 2 - Study-years** | | | | |
| Age at testing (ref=14) | | | | |
| Age 10 at testing | -0.170 *** | -0.168 *** | -0.159 *** | -0.159 *** |
| | (0.024) | (0.024) | (0.017) | (0.017) |
| Age 15 at testing | -0.024 | -0.023 | -0.003 | -0.002 |
| | (0.020) | (0.020) | (0.014) | (0.014) |
| Cohort birth year × Parent education | 0.007 *** | | 0.005 *** | |
| | (0.001) | | (0.001) | |
| Cohort birth year × Parent occupation | 0.007 *** | | 0.007 *** | |
| | (0.001) | | (0.001) | |
| Cohort birth year × Books | 0.008 *** | | 0.006 *** | |
| | (0.001) | | (0.001) | |
| Cohort birth year | | 0.007 *** | | 0.006 *** |
| | | (0.001) | | (0.001) |
| **Random effects** | | | | |
| *Level 2 - Residual variance between studies in…* | | | | |
| Parent education intercepts | 0.03736 | 0.03831 | 0.0219 | 0.02253 |
| Parent occupation intercepts | 0.02322 | 0.02284 | 0.01498 | 0.01496 |
| Books intercepts | 0.03698 | 0.03823 | 0.01607 | 0.01618 |
| *Level 3 - Residual variance between countries in…* | | | | |
| Parent education intercepts | 0.05426 | 0.05362 | 0.03274 | 0.0325 |
| Parent occupation intercepts | 0.05227 | 0.0533 | 0.03748 | 0.0373 |
| Books intercepts | 0.1159 | 0.12149 | 0.05473 | 0.05587 |
| Parent education cohort slopes | 0.00004 | | 0.00002 | |
| Parent occupation cohort slopes | 0.00003 | | 0.00002 | |
| Books cohort slopes | 0.00007 | | 0.00003 | |
| Cohort slopes | | 0.00003 | | 0.00002 |
| N (Level 1 - gaps) | 5541 | 5541 | 5541 | 5541 |
| N (Level 2 - study-years) | 1026 | 1026 | 1026 | 1026 |
| N (Level 3 - countries) | 100 | 100 | 100 | 100 |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: All models in this table specify known level 1 error variances and covariances, estimated via bootstrapping, consistent with models in the main text of the paper.

Aside from test reliability, measurement error in achievement could also decline if tests are administered in a more standardized way over time. Standardization of test administration could explain the reductions in between-country achievement variance reported in Appendix C above. The hierarchical models also found declining between-school and between-classroom achievement variance (for all trend studies except TIMSS 4th and 8th grade math), which may also be partially explained by standardization of test administration. However, achievement variance has also declined within schools and classrooms for most trend studies, and it is less clear how greater standardization of test administration could explain this change.

## E. Changing distribution of SES

Educational attainment and the occupational structure have shifted dramatically in most countries since 1964. The trends reported in the main text of the paper refer to changes in the achievement gap between the 90th, 50th, and 10th percentiles of each SES variable in each country-year, even though the meaning of the 90th, 50th and 10th percentiles has changed over time. Table E1 reports the median category at which the 90th, 50th, and 10th percentiles of each SES variable fall across all participating countries in an older study and the same set of countries in PISA 2015. For parent education and occupation, the older study is FIMS 1964; for household books, it is the SISS 1984 8th grade sample, the first available high-quality measure of household books. It is clear that dramatic upgrading has occurred for both parent education and occupation during this period. However, the distribution of household books has remained more stable.

Table E1. Median Category at which 90th, 50th, and 10th SES Percentiles Fall, by Study

|  |  | FIMS 1964 | PISA 2015 |
| --- | --- | --- | --- |
| Parent education | 90th percentile | 13 years | ISCED 5A+ |
|  | 50th percentile | 9 years | ISCED 5B |
|  | 10th percentile | 7 years | ISCED 3B, C |
| Parent occupation | 90th percentile | Semi-Professional | Professionals |
|  | 50th percentile | Clerical & Sales | Technicians and Associate Professionals |
|  | 10th percentile | Manual Workers, Skilled & Semi-Skilled | Craft Etc Trades Workers |
|  |  | SISS 1984 | PISA 2015 |
| Household books | 90th percentile | 251-500 books | 201-500 books |
|  | 50th percentile | 26-100 books | 26-100 books |
|  | 10th percentile | 11-25 books | 0-10 books |

Note: Countries included in FIMS-PISA sample (parent education and occupation) are Australia, Belgium-Flanders, Belgium-French, England, Finland, France, Germany, Israel, Japan, Netherlands, Scotland, and United States. Countries included in SISS-PISA sample (books) are Australia, Canada, England, Finland, Hong Kong, Hungary, Israel, Italy, Japan, Korea, Netherlands, Norway, Poland, Singapore, Sweden, Thailand, and United States.

Another way to describe the changing SES distribution is the changing share of the sample that falls into the same, consistently-defined set of high, medium, and low SES categories over time. Table E2 reports the average share of the sample falling into each of the three SES categories for a constant sample of countries participating in an older and more recent study. Each SES variable is coded into three categories for all studies: parent education is coded into (1) less than secondary [less than ISCED 3], (2) secondary or non-degree vocational postsecondary [ISCED 3 or 4], and (3) an academically- or vocationally-oriented higher educational degree [ISCED 5A or 5B or more]. Parent occupation is coded into (1) working class [unskilled, semiskilled, or agricultural labor], (2) intermediate class [skilled trades, service, clerical, or small business], (3) salariat class [semi-professional, managerial, or professional]. Household books

are coded into (1) 0-10 books, (2) 11-100 books, (3) 101 books or more. Once again, it is clear that the average level of parent education and occupation have increased dramatically, while the distribution of books has remained relatively constant or even *decreased* slightly. Comparing the SISS 1984 8th grade dataset to PISA 2015, the average share of students reporting 10 or fewer books at home has more than doubled (from 6% to nearly 14%), while the shares of students in the high and medium books categories have both declined by a few percentage points.

Table E2. Proportion of Sample in High, Medium, and Low Categories of SES Variables, by Study

|  |  | FIMS 1964 | PISA 2015 |
|---|---|---|---|
| Parent education | High (ISCED 5A, 5B) | 0.106 | 0.644 |
|  | Medium (ISCED 3, 4) | 0.168 | 0.288 |
|  | Low (<ISCED 3) | 0.726 | 0.067 |
| Parent occupation | High (Salariat) | 0.202 | 0.460 |
|  | Medium (Intermediate) | 0.313 | 0.343 |
|  | Low (Working Class) | 0.485 | 0.196 |
|  |  | SISS 1984 | PISA 2015 |
| Household books | High (101+) | 0.448 | 0.401 |
|  | Medium (11-100) | 0.492 | 0.461 |
|  | Low (0-10) | 0.060 | 0.138 |

Note: Countries included in FIMS-PISA sample (parent education and occupation) are Australia, Belgium-Flanders, Belgium-French, England, Finland, France, Germany, Israel, Japan, Netherlands, Scotland, and United States. Countries included in SISS-PISA sample (books) are Australia, Canada, England, Finland, Hong Kong, Hungary, Israel, Italy, Japan, Korea, Netherlands, Norway, Poland, Singapore, Sweden, Thailand, and United States.

The 90/10 percentile method (Reardon 2011b) was chosen to avoid changes in the selectivity of different SES categories as their frequencies changed over time. However, treating these historical and contemporary percentiles as equivalent also makes a theoretical assumption that these SES characteristics confer mainly positional advantages to children. Alternatively, it may be that, for example, having a parent with a university degree always confers the same absolute advantage, regardless of whether that parent was among the elite few who earned a degree in the mid-20th century or the larger share who earned a degree at the turn of the 21st century. One piece of evidence that the increasing SES achievement gaps reported here are not merely an artifact of the general upgrading of SES is that increases are found not only for parental education and occupation, whose levels have increased over time, but also for household books, whose levels have remained stable or even slightly declined, as seen above.

In addition to changing levels of the three SES variables, the dispersion of SES has also changed. The variance of parent education and occupation has declined somewhat in most countries over time, while the variance of household books has remained relatively constant. (The changing variances of these ordinal SES variables were computed after recoding into the same categories in every study—6 categories for parent education and occupation and 5 categories for books.) All else equal, if the variance of an independent variable—SES—decreases, then its unstandardized association with an outcome variable—achievement—will increase. The models in the main text of the paper avoid this problem, as converting the SES variables into percentiles is a form of standardization. However, these changes in variance should be kept in mind for the next set of models, where SES is unstandardized.

An additional piece of evidence that increasing SES achievement gaps are not an artifact of changing SES distributions comes from the models reported in Tables E3 and E4, which compute achievement gaps between three consistently-defined categories of each SES variable rather than percentiles. These analyses examine the robustness of the finding of increasing SES achievement gaps to treating SES as an absolute rather than a positional good. SES variables are coded as in Table E2 above. Table E3 reports trends in the gap between the top and bottom categories of each variable, while the top panel of Table E4 reports trends in the gap between the top and middle categories, and the bottom panel of Table E4 reports trends in the gap between the middle and bottom categories. Thus, these models are an alternative way to capture changing achievement gaps across the whole SES distribution, at the top, and at the bottom, rather than the 90/10, 90/50, and 50/10 gap trends reported in the main text.

In these models, unlike in the models predicting percentile-based gaps, the size of coefficients cannot be compared across different SES variables because gaps based on categories of different SES variables are not on equivalent scales. Therefore, we look only at whether cohort birth year coefficients are positive. The results in Table E3 show that gaps have increased between the top and bottom categories of all three variables. However, the increase for parent occupation gaps not significantly different from 0. Table E4 shows that, for all three variables, increases between the top and middle categories are larger than increases between the middle and bottom categories. This is true both for parent education and occupation, where the top category has become a larger share of students, and for household books, where the top category has become a slightly smaller share of students. Thus, the achievement advantage of students with college-educated or professional parents or many books at home has increased, even as the share of students with college-educated or professional parents has increased (and the share with many books has declined).

That the gap between the top and middle SES categories has increased more than between the middle and bottom categories may at first appear inconsistent with the results from the main text of the paper showing that 50/10 gaps have increased more than 90/50 gaps in most countries. However, these two findings are simultaneously true. Comparing between descriptive Tables E1 and E2 shows why this is the case. In PISA 2015, the $50^{th}$ and $10^{th}$ percentiles of parent education and occupation in fact correspond to the "high" and "medium" categories of parent education and occupation. This demonstrates the difficulty of studying trends in achievement gaps between categories with drastically changing distributions. I choose to treat SES as a positional rather than an absolute good in this study, as it seems a more tenable assumption that the advantage conferred by a particular level of parental education or occupation changes over time with the changing distributions of these variables. However, the results of these models indicate that achievement gaps between high and low SES categories still appear to increase even when SES is treated as an absolute good.

Table E3. Estimated Trends in Achievement Gaps between High and Low SES Categories

| | Education | Occupation | Books |
|---|---|---|---|
| *High-low* | | | |
| Age 10 at testing | -0.038 | 0.099 *** | -0.221 *** |
| | (0.032) | (0.020) | (0.028) |
| Age 15 at testing | -0.083 ** | | -0.001 |
| | (0.030) | | (0.025) |
| Math | 0.03 ** | 0.002 | -0.02 + |
| | (0.010) | (0.008) | (0.012) |
| Science | 0.032 *** | 0.001 | 0.044 *** |
| | (0.009) | (0.007) | (0.013) |
| Cohort birth year | 0.007 *** | 0.002 * | 0.003 + |
| | (0.001) | (0.001) | (0.002) |
| Intercept | 0.976 *** | 0.825 *** | 1.318 *** |
| | (0.029) | (0.021) | (0.044) |
| Residual variance (within countries) | 0.03481 | 0.00851 | 0.02560 |
| Residual variance (country intercepts) | 0.05793 | 0.02703 | 0.16224 |
| Residual variance (cohort slopes) | 0.00009 | 0.00007 | 0.00015 |
| N (observations) | 1889 | 1334 | 2086 |
| N (countries) | 93 | 80 | 95 |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: See text for definitions of high, middle, and low for each SES variable. To avoid very long computation times, all models in this table specify known level 1 error variances estimated using conventional non-bootstrap formulas and omit error covariances. This simplified specification appears to produce very similar results to models using bootstrapped error variances and covariances (see Appendix L).

Table E4. Estimated Trends in Achievement Gaps between High-Middle and Middle-Low SES Categories

| | Education | Occupation | Books |
|---|---|---|---|
| *High-middle* | | | |
| Age 10 at testing | -0.029 | -0.036 ** | -0.169 *** |
| | (0.019) | (0.012) | (0.015) |
| Age 15 at testing | -0.088 *** | | 0.011 |
| | (0.016) | | (0.013) |
| Math | 0.031 *** | 0.012 * | 0.001 |
| | (0.006) | (0.005) | (0.008) |
| Science | 0.019 *** | 0.012 *** | 0.0270 *** |
| | (0.005) | (0.004) | (0.007) |
| Cohort birth year | 0.005 *** | 0.002 + | 0.002 * |
| | (0.001) | (0.001) | (0.001) |
| Intercept | 0.496 *** | 0.548 *** | 0.62 *** |
| | (0.014) | (0.016) | (0.022) |
| Residual variance (within countries) | 0.01243 | 0.00319 | 0.00907 |
| Residual variance (country intercepts) | 0.01691 | 0.01686 | 0.04203 |
| Residual variance (cohort slopes) | 0.00003 | 0.00005 | 0.00007 |
| N (observations) | 1889 | 1334 | 2086 |
| N (countries) | 93 | 80 | 95 |
| *Middle-low* | | | |
| Age 10 at testing | -0.004 | 0.137 *** | -0.05 * |
| | (0.021) | (0.015) | (0.024) |
| Age 15 at testing | 0.004 | | -0.004 |
| | (0.023) | | (0.018) |
| Math | -0.001 | -0.009 + | -0.024 *** |
| | (0.007) | (0.005) | (0.007) |
| Science | 0.012 + | -0.013 ** | 0.015 * |
| | (0.007) | (0.004) | (0.007) |
| Cohort birth year | 0.001 | 0.001 | 0.001 |
| | (0.001) | (0.001) | (0.001) |
| Intercept | 0.484 *** | 0.273 *** | 0.698 *** |
| | (0.020) | (0.010) | (0.026) |
| Residual variance (within countries) | 0.01781 | 0.00297 | 0.01407 |
| Residual variance (country intercepts) | 0.02717 | 0.00471 | 0.05037 |
| Residual variance (cohort slopes) | 0.00004 | 0.00004 | 0.00005 |
| N (observations) | 1889 | 1334 | 2086 |
| N (countries) | 93 | 80 | 95 |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: See text for definitions of high, middle, and low for each SES variable. To avoid very long computation times, all models in this table specify known level 1 error variances estimated using conventional non-bootstrap formulas and omit error covariances. This simplified specification appears to produce very similar results to models using bootstrapped error variances and covariances (see Appendix L).

**F. Achievement gaps by mother's and father's SES characteristics**

One possible explanation for increasing SES achievement gaps is the increasing educational attainment and occupational status of mothers. The dramatic global rise in women's status since the mid-20th century is easily visible in the international assessment data. In early international assessments, most students reported a higher level of educational attainment for their father than their mother. The mother's occupation was not collected at all in the earliest three datasets (FIMS 1964, FISS 1970, and FIRCS 1970). In the most recent assessments, mothers and fathers are about equally educated, and mothers have somewhat higher occupational status than fathers (reflecting women's greater likelihood of working in white collar jobs). Since the main analyses in study use only the *highest* of the two parents' education and occupation as a measure of the child's SES, the increasing education and occupational status of mothers means that the highest parent education and occupation are increasingly likely to come from the mother. If children's achievement tends to be more strongly associated with their mother's than with their father's education and occupation (because mothers perform the majority of childcare), then the increasing educational and occupational attainment of mothers could explain why SES achievement gaps are increasing. (Children and/or parents were asked to report the education and occupation of both parents, whether or not both were present in the home, and most international assessments did not collect data on whether each parent lived in the home. For parents not currently working, the most recent occupation was reported. For parents who had never worked—who were very likely to be mothers performing home duties, particularly in earlier years—this study treats that parent's occupation as missing and imputes an occupation as part of the multiple imputation model described in the Methods section of the main text of the paper.)[5]

To check the robustness of the main results to changes in the relative status of mothers and fathers, Table F1 reports trends in SES achievement gaps based on fathers' and mothers' education and occupation separately. The results show that trends in gaps based on all four characteristics are positive and significant. As expected, increases in gaps based on mothers' education and occupation are larger than those based on fathers' characteristics—a Wald joint hypothesis test shows that the increase in the mother's education gap is significantly larger than the increase in the father's education gap (p < .001). The trends in gaps based on mothers' and fathers' occupation differ only slightly, by less than 0.001 SD per year, a difference that is not significant (p > 0.5). Between the 1950 and the 2005 birth cohorts, the father's education achievement gap grew from about 0.72 SDs to 1.10 SDs (about a 54% increase), while the mother's education achievement gap grew from 0.68 SDs to 1.12 SDs (about a 65% increase). Between the 1966 and 2005 birth cohorts (1966 is the birth cohort corresponding to SIMS 1980, the first cohort for which both mothers' and fathers' occupation are available), the father's occupation achievement gap grew from 0.81 to 1.04 SDs (about a 29% increase), and the mother's occupation achievement gap grew from 0.83 to 1.07 SDs (about a 28% increase). That achievement gaps have increased not only for mothers' but also fathers' education and occupation suggests that the global increase in parent education and occupation achievement gaps is not fully explained by the increasing educational or occupational attainment of mothers.

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[5] See Appendix L for models computing gap trends with listwise deletion of missing data rather than multiple imputation. In these models, for students with one missing and one nonmissing parent education or occupation, the nonmissing value was used as the "highest" parent education or occupation. Results are very similar to those from the models with imputed data.

Table F1. Trends in Achievement Gaps Based on Fathers' and Mothers' SES Characteristics

| | Coef | (se) | |
|---|---|---|---|
| Father's education gaps intercept | 0.991 | (0.028) | *** |
| Mother's education gaps intercept | 0.993 | (0.030) | *** |
| Father's occupation gaps intercept | 0.947 | (0.027) | *** |
| Mother's occupation gaps intercept | 0.970 | (0.029) | *** |
| Level 1 - Gaps | | | |
|     Subject (ref=Reading): | | | |
|       Math | 0.036 | (0.007) | *** |
|       Science | 0.028 | (0.005) | *** |
|     SES variable quality measures | | | |
|       Parent-reported × Parent education | 0.091 | (0.029) | ** |
|       Parent-reported × Parent occupation | -0.054 | (0.025) | * |
|       Number of categories (centered at 7) | 0.001 | (0.002) | |
|       ≥ 20% in bottom category | -0.044 | (0.011) | *** |
|       ≥ 20% in top category | -0.047 | (0.009) | *** |
| Level 2 - Study-years | | | |
|     Age at testing (ref=14) | | | |
|       Age 10 at testing | -0.091 | (0.033) | ** |
|       Age 15 at testing | -0.116 | (0.023) | *** |
|     Cohort birth year × Father's education | 0.007 | (0.001) | *** |
|     Cohort birth year × Mother's education | 0.008 | (0.001) | *** |
|     Cohort birth year × Father's occupation | 0.006 | (0.001) | *** |
|     Cohort birth year × Mother's occupation | 0.006 | (0.001) | *** |
| Random effects | | | |
| *Level 2 - Residual variance between studies in…* | | | |
|     Father's education intercepts | 0.03050 | | |
|     Mother's education intercepts | 0.03658 | | |
|     Father's occupation intercepts | 0.01361 | | |
|     Mother's occupation intercepts | 0.01496 | | |
| *Level 3 - Residual variance between countries in…* | | | |
|     Father's education intercepts | 0.04857 | | |
|     Mother's education intercepts | 0.05634 | | |
|     Father's occupation intercepts | 0.03237 | | |
|     Mother's occupation intercepts | 0.04390 | | |
|     Father's education cohort slopes | 0.00006 | | |
|     Mother's education cohort slopes | 0.00005 | | |
|     Father's occupation cohort slopes | 0.00004 | | |
|     Mother's occupation cohort slopes | 0.00005 | | |
| N (Level 1 - gaps) | 6502 | | |
| N (Level 2 - study-years) | 866 | | |
| N (Level 3 - countries) | 95 | | |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: All models in this table specify known level 1 error variances and covariances, estimated via bootstrapping, consistent with models in the main text of the paper.

Another possible explanation for increasing SES achievement gaps is increasing homogamy among the parents of participating students. That is, the SES characteristics (education and occupation) of mothers and fathers are likely growing more correlated over time. Students with two highly-educated or high-occupational-status parents may be more advantaged than students with only one highly-educated or high-occupational-status parent. Table F2 reports hierarchical growth models estimating trends in the correlation between mothers' and fathers' education and occupation. The results show that the average correlation between mothers' and fathers' education has increased only slightly (from about 0.56 in the 1950 birth cohort to about 0.60 in the 2005 cohort), and the average correlation between mothers' and fathers' occupation has remained relatively constant (declining slightly from about 0.40 in the 1966 birth cohort to about 0.39 in the 2005 cohort). These results demonstrate a less pronounced increase in homogamy than expected, which may be because this analysis treats parent education and occupation as continuous positional goods, converted into percentiles within each country-year, as in the main text of this paper. While it is true that an increasing number of children have two parents with higher education degrees or professional occupations, the associations between the *relative* positions of mothers and fathers within their own gender distributions have not increased dramatically over time.

Table F2. Trends in Correlations between Mothers' and Fathers' Education and Occupation

|  | Education | Occupation |
|---|---|---|
| Intercept | 0.589 *** | 0.392 *** |
|  | (0.006) | (0.009) |
| Level 1 - Study-years |  |  |
| Age at testing (ref=14) |  |  |
| Age 10 at testing | -0.025 *** | 0.090 *** |
|  | (0.007) | (0.008) |
| Age 15 at testing | -0.054 *** |  |
|  | (0.006) |  |
| Cohort birth year | 0.0006 * | -0.0003 |
|  | (0.0003) | (0.0005) |
| Level 1 residual variance | 0.00234 | 0.00365 |
| Level 2 residual variance in intercepts | 0.00341 | 0.00523 |
| Level 2 residual variance in cohort slopes | 0.00000 | 0.00001 |
| N (Level 1 - study-years) | 866 | 550 |
| N (Level 2 - countries) | 95 | 82 |

+ p<.1, * p<.05, ** p<.01, *** p < .001

# G. Achievement gaps conditional on other SES variables

Results in the main paper text estimate achievement gaps separately for each SES variable (parent education, parent occupation, and household books) rather than constructing an SES index to avoid loss of information because not all SES variables are available in every dataset. However, it could be the case that only one of the three SES variables is growing more strongly associated with achievement over time, while the other two SES variables only appear to be growing more strongly associated with achievement due to their correlations with this one most salient SES variable. Or it may even be the case that the independent associations between each SES variable and achievement have remained constant over time, but correlations between

the three SES variables are growing stronger over time. This would create the appearance of increasing SES achievement gaps for all three variables because an increasing share of students would experience "double-" or "triple disadvantage." That is, students with university-educated parents would be more likely also to have parents with professional occupations and to have a large number of books at home. Conversely, students whose parents have not completed secondary education would be more likely to have parents with working-class occupations and very few books at home. Thus, there may be a pattern of increasing polarization of socioeconomic advantage and disadvantage among schoolchildren, which may completely explain away increasing SES achievement gaps for all three SES variables.

Table G1 reports hierarchical growth models estimating trends in the correlation between pairs of SES variables. The results show that the average correlation between parent education and parent occupation has increased quite substantially (from about 0.43 in the 1950 birth cohort to about 0.56 in the 2005 cohort). This finding is consistent with international research showing an increasing association between education and occupation across most countries (Kreidl, Ganzeboom and Treiman 2014).[6] In contrast, the average correlations between household books and each of parent education and occupation, respectively, have increased more moderately (the correlation between household books and parent education increased from about 0.31 in the 1956 birth cohort to about 0.35 in the 2005 cohort; the correlation between household books and parent occupation increased from about 0.31 in the 1956 birth cohort to about 0.39 in the 2005 birth cohort).

Table G1. Trends in Correlations between Three SES Variables

|  | Education & Occupation | Education & Books | Occupation & Books |
|---|---|---|---|
| Intercept | 0.523 *** | 0.338 *** | 0.361 *** |
|  | (0.042) | (0.007) | (0.014) |
| Level 1 - Study-years |  |  |  |
| Age at testing (ref=14) |  |  |  |
| Age 10 at testing | 0.057 | 0.064 *** | -0.007 |
|  | (0.048) | (0.007) | (0.012) |
| Age 15 at testing | -0.048 | -0.031 *** | -0.059 *** |
|  | (0.044) | (0.005) | 0.012 |
| Cohort birth year | 0.0025 ** | 0.0008 + | 0.0016 *** |
|  | (0.0008) | (0.0004) | (0.0004) |
| Level 1 residual variance | 0.00529 | 0.00195 | 0.0018 |
| Level 2 residual variance in intercepts | 0.00437 | 0.00339 | 0.00541 |
| Level 2 residual variance in cohort slopes | 0.00001 | 0.00001 | 0.00001 |
| N (Level 1 - study-years) | 576 | 835 | 567 |
| N (Level 2 - countries) | 82 | 95 | 83 |

+ p<.1, * p<.05, ** p<.01, *** p < .001

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[6] Note the literature shows an increasing association between education and occupation only when both variables are treated as linear (as here); the true pattern of change may be more complex. For example, research on over-education shows a declining relationship between attainment of tertiary education and a professional occupation when treating both variables as categorical rather than linear—that is, assuming education and occupational status are absolute rather than positional goods (the opposite of the assumption made in this paper).

But do these increasing correlations fully explain increasing SES achievement gaps for all three variables? And even if increasing correlations do not fully explain increasing gaps, could it be the case that only one or two SES variables are growing more strongly associated with achievement, while the other SES variable(s) only appear to be growing more strongly associated with achievement due to their correlation(s) with the most salient SES variable(s)? One way to address both of these questions is by computing SES achievement gaps for each variable conditional on one or both of the other SES variables. The results of these models are presented in Table G2. Conditional gaps can be estimated only from studies that collected more than one SES variable, meaning sample sizes are reduced. To obtain an accurate comparison to cohort trends in unconditional SES achievement gaps, trends in unconditional gaps are also estimated using the same reduced set of studies. In Table G2, each row reports two models estimating gaps based on a particular SES variable. The "Conditional gaps" columns report the intercept and cohort birth year coefficients for a model whose gaps are conditional on one or both of the other SES variables. The "Unconditional gaps" columns report the coefficients for a model based on unconditional gaps using the same sample of studies. In each model, the intercept is the estimated gap for the variable in question (conditional or unconditional) for the 1989 birth cohort, while the cohort trend is the estimated annual change in the gap. All models control for subject and age at testing. Since there is no established method to adjust conditional associations for attenuation due to measurement error, conditional gaps are not adjusted for test or SES reliability. For an accurate comparison, the unconditional gaps are also not adjusted for reliability.

Table G2. Intercept and Cohort Birth Year Coefficients from Models Predicting 90/10 SES Achievement Gaps Conditional on Other SES Variables

| Variable | Conditional on | Conditional gaps | | | Unconditional gaps | | |
|---|---|---|---|---|---|---|---|
| | | Intercept | Cohort trend | (N) | Intercept | Cohort trend | (N) |
| Education | Occupation | 0.444 *** | 0.002 | (81) | 0.760 *** | 0.005 ** | (81) |
| Education | Books | 0.389 *** | 0.003 ** | (95) | 0.729 *** | 0.006 *** | (95) |
| Education | Occupation & Books | 0.302 *** | 0.001 | (81) | 0.750 *** | 0.003 * | (81) |
| Occupation | Education | 0.361 *** | 0.001 | (81) | 0.684 *** | 0.004 *** | (81) |
| Occupation | Books | 0.356 *** | 0.001 | (82) | 0.672 *** | 0.004 *** | (82) |
| Occupation | Education & Books | 0.278 *** | -0.001 | (81) | 0.729 *** | 0.003 * | (81) |
| Books | Education | 0.483 *** | 0.004 *** | (95) | 0.806 *** | 0.007 *** | (95) |
| Books | Occupation | 0.502 *** | 0.002 ** | (82) | 0.785 *** | 0.005 *** | (82) |
| Books | Education & Occupation | 0.475 *** | 0.004 *** | (81) | 0.797 *** | 0.006 *** | (81) |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: To avoid very long computation times, all models in this table specify known level 1 error variances estimated using conventional non-bootstrap formulas and omit error covariances. This simplified specification appears to produce very similar results to models using bootstrapped error variances and covariances (see Appendix L).

By computing the predictions of the model with all three SES variables (rows 3, 6, and 9) for the 1956 birth cohort (the first cohort with all three variables available) and the 2005 cohort, it can be seen that in both years, books had the strongest independent association with achievement (0.54), followed by education (0.32) and then occupation (0.26). However, over this time period, this ranking became even more pronounced. The independent 90/10 books gap increased markedly from 0.34 to 0.54, while the independent 90/10 parent education gap increased more modestly from 0.27 to 0.32, and the independent parent occupation gap

*decreased* from 0.31 to 0.26. (Note this does not mean parent occupation is not significantly related to student achievement after controlling for parent education and household books, only that the relationship between parent education and achievement is not growing, after accounting for the other two SES variables). Thus, it appears that the independent associations of each SES variable with achievement have changed at different rates over time.

That the conditional 90/10 parent education and household books gaps show increases suggests that the global increases in unconditional SES achievement gaps for these variables are not fully explained by their correlations with other SES characteristics. In contrast, the declining conditional 90/10 parent occupation gap suggests that the growing parent occupation gap may be fully explained by the correlations between parent occupation and other SES characteristics. However, it should be noted that the model with all three SES variables included is based on a substantially reduced sample of only those datasets that collected all three SES variables, and thus may not be representative.

Regardless of the relative importance of parent education, parent occupation, and books, the results in Table G2 clearly indicate that increasing correlations between SES variables do not fully explain increasing SES achievement gaps. Nearly all trends in conditional SES achievement gaps are positive, with the exception of parent occupation conditional on education and books, as noted above. These results suggest that, even after accounting for the growing number of children with "double-" or "triple-disadvantage" due to increasing correlations among SES variables, each SES variable (or at least parent education and books) has become more consequential for students' academic achievement.

One further piece of evidence that increasing SES achievement gaps for each SES variable are not entirely due to increasing correlations between variables is that the $R^2$ of the models used to compute conditional gaps with all three variables has increased over time. Table G3 shows results from hierarchical growth models (country-subject-years within countries) predicting $R^2$ from cohort birth year (with controls for age and subject). ($R^2$ is adjusted for test reliability before running these models.) Results indicate that the $R^2$ of the model including all three SES variables nearly doubled from about 0.10 in the 1956 birth cohort to about 0.18 in the 2005 birth cohort. Thus, it appears that the overall predictive power of SES on achievement has grown substantially stronger over this 49-year time period.

Table G3. Trends in $R^2$ from Models with Two or Three SES Variables

| | $R^2$ (Education & Occupation) | $R^2$ (Education & Books) | $R^2$ (Occupation & Books) | $R^2$ (Education, Occupation & Books) |
|---|---|---|---|---|
| Intercept | 0.107 *** | 0.138 *** | 0.129 *** | 0.154 *** |
| | (0.013) | (0.006) | (0.010) | (0.016) |
| Level 1 - $R^2$ measures | | | | |
| Subject (ref=Reading): | | | | |
| Math | 0.006 *** | 0.002 | 0.001 | 0.003 + |
| | (0.002) | (0.002) | (0.002) | (0.002) |
| Science | 0.005 *** | 0.009 *** | 0.005 ** | 0.006 *** |
| | (0.001) | (0.001) | (0.002) | (0.001) |
| Level 2 - Study-years | | | | |
| Age at testing (ref=14) | | | | |
| Age 10 at testing | 0.022 | -0.015 ** | -0.008 | -0.010 |
| | (0.013) | (0.006) | (0.008) | (0.015) |
| Age 15 at testing | 0.002 | -0.001 | 0.037 *** | 0.02 |
| | (0.014) | (0.005) | 0.01 | 0.016 |
| Cohort birth year | 0.0010 *** | 0.0021 *** | 0.0014 *** | 0.0017 *** |
| | (0.0003) | (0.0002) | (0.0003) | (0.0003) |
| Level 1 residual variance | 0.000198 | 0.000322 | 0.000293 | 0.000323 |
| Level 2 residual variance in intercepts | 0.000806 | 0.001049 | 0.000728 | 0.000853 |
| Level 3 residual variance in intercepts | 0.001443 | 0.002646 | 0.002701 | 0.002623 |
| Level 3 residual variance in cohort slopes | 0.000002 | 0.000002 | 0.000002 | 0.000002 |
| N (Level 1 - $R^2$ measures) | 1382 | 1910 | 1389 | 1351 |
| N (Level 2 - study-years) | 576 | 836 | 567 | 545 |
| N (Level 3 - countries) | 82 | 95 | 83 | 82 |

+ p<.1, * p<.05, ** p<.01, *** p < .001

**H. Changing measurement error of SES**

Another reason why SES achievement gaps may artificially appear to be increasing over time is that the reliability with which SES is measured may be increasing over time. Lower reliability of SES in early years could cause the estimated association between SES and achievement in early years to be attenuated, creating the appearance of increasing SES achievement gaps over time. Reliability is defined as the ratio of the variance in true SES to the total variance in SES (including both true variance and the variance of errors of measurement). As reported in Appendix E, the total variance of SES has declined for two SES variables—parent education and occupation—while the total variance has remained relatively constant for household books. It is likely that much of the decline in total variance in parent education and occupation reflects a decline in true variance, due to the large increase in the average *levels* of both variables as more parents attain higher education and professional occupations.

However, it is also likely that some of the reduction in the total variance of parent education and occupation is due to declining measurement error in these variables. Measurement error could potentially decline, for example, because current students have more accurate knowledge of their families' SES characteristics than in the past, because of improvements in survey wording, or because more recent SES data are more likely to be reported by parents rather than students. As stated in the main paper text, some international assessments have added parent questionnaires in recent years (PIRLS 2001-2011, TIMSS 4th grade 2011-2015, and PISA 2006-2012). The main models use parent-reported SES variables when available and student-reported SES otherwise. However, it is expected that parents report SES variables more reliably than their children. Some other patterns reported in previous appendix sections are consistent with increasing reliability of SES due to declining measurement error, while other patterns are inconsistent with this story. The large increase in the correlation between parent education and occupation (reported in Appendix G) is consistent with increasing reliability in these variables. However, the increase in the correlations between household books and each of the other two SES variables have been more modest. An increasing correlation between mothers' and fathers' education and occupation (reported in Appendix F) could also be evidence for increasing reliability of these variables. However, the correlation between mothers' and fathers' education has increased only moderately, and the correlation between mothers' and fathers' occupations has *declined* slightly.

As described in the Methods section of the main paper text, all SES achievement gaps have been adjusted for estimated reporting reliability of SES variables. Following Reardon (2011a), this adjustment consists of by multiplying each gap estimate by $\frac{1}{\sqrt{r}}$, where $r$ is the reliability of the SES measure. In order to estimate the reliability $r$ of each student- or parent-reported SES measure, we can take advantage of having two measures of the same variable reported by different sources (i.e., students and parents) (Jerrim and Micklewright 2014). The reliability of students' and parents' SES reports can be computed from the following formulas (Reardon 2011a):

$$r_s = corr(s,p) \cdot \frac{corr(s,y)}{corr(p,y)}$$
$$r_p = corr(s,p) \cdot \frac{corr(p,y)}{corr(s,y)}$$

where $r_s$ is the reliability of student-reported SES variable $s$, $r_p$ is the reliability of parent-reported SES variable $p$, and $y$ is a third variable that would have a particular correlation with true SES, were SES measured without error. In reliability calculations from PISA and TIMSS, math achievement is used for this third variable $y$; in reliability calculations from PIRLS, reading achievement is used for $y$. Parent education is reported by both 15-year-old students and parents in PISA 2006, 2009, and 2012. Parent occupation is reported by both 15-year-old students and parents in PISA 2006 and 2012. Household books are reported by both 4th grade students and parents in PIRLS 2001, 2006, and 2011 and in TIMSS 2011 and 2015. I estimate the average reliability of parents' reports of their own educational attainment (across PISA 2006-2012) at 0.84 and students' reports of their parents' education at 0.62. I estimate the average reliability of parents' reports of their own occupational category (across PISA 2006 and 2012) at 0.81 and students' reports of their parents' occupation at 0.79. I estimate the average reliability of parents' reports of the number of household books (across PIRLS 2001-2011 and TIMSS 2011-2015) at 0.52 and students' reports of household books at 0.46. The higher accuracy of parent occupation reports and low accuracy of household books reports is consistent with findings by Jerrim and Micklewright (2014) using some of the same international datasets. In order to estimate reliabilities for other age groups, I assume that 8th grade students report all SES variables with the same reliability as 15-year-old students, but 4th grade students report parent education 80% as reliably and parent occupation and household books 90% as reliably as 15-year-old students. Finally, in order to estimate reliabilities for other years where parent reports are unavailable, I use the average reliabilities for each of these age groups. Since the reliabilities applied to all years are derived from parent reports in recent years, this procedure adjusts only for differences in reliability between parents and children, but cannot account for possible changes in reliability over time.

Table H1 reports estimated trends in 90/10 SES achievement gaps, using only student-reported data for all three SES variables (gaps are not adjusted for SES or test reliability). Sample sizes are reduced, mainly due to the omission of 4th grade assessments from recent years with parent-reported education and occupation. In these models, the estimated positive trends in 90/10 SES achievement gaps are reduced compared to those reported in the main paper text but are still positive and highly significant for all three SES variables. Using these results, it is possible to estimate the sensitivity of estimated gap trends to potential increases in the accuracy of students' reports of their parents' SES characteristics. For the increases in SES achievement gaps reported in Table H1 to be fully accounted for by measurement error alone, the reliability of students' reports of parental education, estimated at 0.62 for recent cohorts, would have to be only 0.46 for the 1950 cohort. The reliability of students' reports of parental occupation, estimated at 0.77 for recent cohorts, would have to be 0.50 for the 1950 cohort; and the reliability of students' reports of household books, estimated at 0.46 for recent cohorts, would have to be 0.34 for the 1956 cohort (the first cohort for which the household books variable was collected). Without parental reports, it is impossible to know from these data whether the reliability of students' reports could have increased by 35-55% over this 50 year period. However, a thorough literature search did not reveal published evidence that survey reporting of SES characteristics by either adults or children has become more accurate over time.

Table H1. Estimated trends in 90/10 SES achievement gaps, student-reported SES data only

| | No adjustment | | | | | |
|---|---|---|---|---|---|---|
| | (1) | | | (2) | | |
| | coef | (se) | | coef | (se) | |
| Parent education gaps intercept | 0.739 | (0.022) | *** | 0.752 | (0.021) | *** |
| Parent occupation gaps intercept | 0.772 | (0.024) | *** | 0.782 | (0.024) | *** |
| Household books gaps intercept | 0.829 | (0.028) | *** | 0.838 | (0.027) | *** |
| Level 1 - Gaps | | | | | | |
| Subject (ref=Reading): | | | | | | |
| Math | 0.018 | (0.006) | *** | 0.019 | (0.006) | *** |
| Science | 0.026 | (0.004) | *** | 0.026 | (0.004) | *** |
| SES variable quality measures | | | | | | |
| Number of categories (centered at 7) | -0.002 | 0.002 | | -0.001 | 0.002 | |
| ≥ 20% in bottom category | -0.025 | 0.014 | + | -0.033 | 0.016 | * |
| ≥ 20% in top category | -0.105 | 0.01 | *** | -0.116 | 0.009 | *** |
| Level 2 - Study-years | | | | | | |
| Age at testing (ref=14) | | | | | | |
| Age 10 at testing | -0.186 | (0.017) | *** | -0.185 | (0.017) | *** |
| Age 15 at testing | 0.014 | (0.015) | | 0.007 | (0.014) | |
| Cohort birth year × Parent education | 0.004 | (0.001) | *** | | | |
| Cohort birth year × Parent occupation | 0.006 | (0.001) | *** | | | |
| Cohort birth year × Books | 0.005 | (0.001) | *** | | | |
| Cohort birth year | | | | 0.005 | (0.001) | *** |
| Random effects | | | | | | |
| *Level 2 - Residual variance between studies in…* | | | | | | |
| Parent education intercepts | 0.01962 | | | 0.01991 | | |
| Parent occupation intercepts | 0.01561 | | | 0.01553 | | |
| Books intercepts | 0.01461 | | | 0.01514 | | |
| *Level 3 - Residual variance between countries in…* | | | | | | |
| Parent education intercepts | 0.02999 | | | 0.03059 | | |
| Parent occupation intercepts | 0.03917 | | | 0.03960 | | |
| Books intercepts | 0.06079 | | | 0.06187 | | |
| Parent education cohort slopes | 0.00003 | | | | | |
| Parent occupation cohort slopes | 0.00002 | | | | | |
| Books cohort slopes | 0.00004 | | | | | |
| Cohort slopes | | | | 0.00002 | | |
| N (Level 1 - gaps) | 4980 | | | 4980 | | |
| N (Level 2 - studies) | 1023 | | | 1023 | | |
| N (Level 3 - countries) | 100 | | | 100 | | |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: All models in this table specify known level 1 error variances and covariances, estimated via bootstrapping, consistent with models in the main text of the paper.

The estimates of the reliability of students' SES reports from recent years (for countries that collected SES from both students and parents) do vary substantially. Reliability estimates of students' reports of parent education from PISA 2012 range from 0.43 in Croatia to 0.91 in Portugal; of parent occupation range from 0.66 in Mexico to 0.92 for Croatia; and of household books from PIRLS 2011 range from 0.03 in Kuwait to 0.67 in Bulgaria. Yet for parent education and occupation, even the countries with the lowest estimated reliabilities barely reach what would need to be the *average* level of reliability for early cohorts in order to fully explain the global increase in SES achievement gaps. For household books, in contrast, there is more variability in the estimated accuracy of students' reports, and it is conceivable that average reliability could have increased by 37% in 49 years.

Estimated SES reliability could potentially also be compared across multiple waves of each study that collects a parent questionnaire (PIRLS 2001-2011 and PISA 2006-2012). Computed SES reliabilities for students and parents do appear to vary somewhat across waves, although it is not clear that these differences represent meaningful trends, given the small number of years and countries represented. It is also possible to compare trends in SES achievement gaps estimated from students' versus parents' reports across waves of PIRLS and PISA. These estimates are reported in Table H2. The estimated trends do not appear to differ systematically depending on whether they are estimated from parent or student reports (this comparison was done before adjusting for computed SES reliability, as gaps based on parent- and student-reported data will be nearly identical after adjustment by construction). Wald tests for joint null hypotheses that trends are equal for gaps based on parent- and student-reported SES cannot be rejected in all but one model. Correlations between country-specific random cohort slopes for gaps based on parent- and student-reported SES are also strongly positive all but one model (this excludes models predicting parent occupation gaps, where random cohort slopes could not be computed due to insufficient sample size). The exception in both cases is the model predicting parent education gaps in PISA science scores, which also showed negative trends in Appendix B. The results in Table H2 suggest this finding for science gaps based student-reported parent education may be inaccurate, as it is inconsistent with the results for science gaps based on parent-reported education, which are likely to be more reliably estimated.

Table H2. Trends in Gaps by Student-Reported and Parent-Reported SES, by Test

| Test instrument | PISA Math | PISA Reading | PISA Science | PISA Math | PISA Reading | PISA Science | PIRLS Reading |
|---|---|---|---|---|---|---|---|
| Test years | 2006-2012 | 2006-2012 | 2006-2012 | 2006-2012 | 2006-2012 | 2006-2012 | 2001-2011 |
| Cohort birth years | 1991-1997 | 1991-1997 | 1991-1997 | 1991-1997 | 1991-1997 | 1991-1997 | 1991-2001 |
| SES variable | Education | Education | Education | Occupation | Occupation | Occupation | Books |
| Parent education (parent report) intercept | 0.828 *** | 0.747 *** | 0.810 *** | | | | |
| | (0.066) | (0.069) | (0.070) | | | | |
| Parent education (student report) intercept | 0.747 *** | 0.681 *** | 0.744 *** | | | | |
| | (0.064) | (0.064) | (0.069) | | | | |
| Parent occupation (parent report) intercept | | | | 0.690 *** | 0.649 *** | 0.672 *** | |
| | | | | (0.102) | (0.102) | (0.107) | |
| Parent occupation (student report) intercept | | | | 0.673 *** | 0.629 *** | 0.661 *** | |
| | | | | (0.115) | (0.116) | (0.120) | |
| Household books (parent report) intercept | | | | | | | 0.807 *** |
| | | | | | | | (0.029) |
| Household books (student report) intercept | | | | | | | 0.756 *** |
| | | | | | | | (0.037) |
| Level 2 - Study-years | | | | | | | |
| Cohort birth year × Parent ed. (parent) | 0.019 ** | 0.029 *** | 0.014 * | | | | |
| | (0.006) | (0.009) | (0.006) | | | | |
| Cohort birth year × Parent ed. (student) | 0.012 *** | 0.015 * | -0.001 | | | | |
| | (0.003) | (0.006) | (0.007) | | | | |
| Cohort birth year × Parent occ. (parent) | | | | 0.013 * | 0.013 * | 0.008 | |
| | | | | (0.005) | (0.005) | (0.005) | |
| Cohort birth year × Parent occ. (student) | | | | 0.017 ** | 0.014 * | 0.010 + | |
| | | | | (0.006) | (0.007) | (0.005) | |
| Cohort birth year × Books (parent) | | | | | | | 0.005 * |
| | | | | | | | (0.003) |
| Cohort birth year × Books (student) | | | | | | | 0.009 ** |
| | | | | | | | (0.003) |
| N (Level 1 - gaps) | 66 | 66 | 66 | 28 | 28 | 28 | 198 |
| N (Level 2 - study-years) | 33 | 33 | 33 | 14 | 14 | 14 | 99 |
| N (Level 3 - countries) | 13 | 13 | 13 | 7 | 7 | 7 | 40 |
| p-value for H0: Cohort × SES (parent) = Cohort × SES (student) | 0.233 | 0.025 | 0.071 | 0.272 | >.500 | >.500 | 0.118 |
| Correlation between country random cohort slopes for parent- and student-reported SES | 0.761 | 0.744 | -0.066 | a | a | a | 0.951 |

+ p<.1, * p<.05, ** p<.01, *** p < .001 Note: To avoid very long computation times, all models in this table specify known level 1 error variances estimated using conventional non-bootstrap formulas and omit error covariances. This simplified specification appears to produce very similar results to models using bootstrapped error variances and covariances (see Appendix L).

[a] Country-level random cohort slopes could not be included in the models predicting parent occupation gaps due to an insufficient sample size.

Increased accuracy of the SES variables could result not only from improvements in students' knowledge of their family SES characteristics but also from improved questionnaire wording over time. The wording of the background questionnaires differs across the older IEA studies, the new IEA studies (TIMSS and PIRLS), and PISA. However, the wording of background questionnaires has changed very little across multiple waves of the trend studies (TIMSS, PIRLS, and PISA). Yet, as demonstrated in Appendix B, trends in SES achievement gaps estimated from each of these studies individually are still nearly always positive (though not always statistically significant, given the smaller sample sizes). Between 1995 and 2001, the IEA began including drawings of bookshelves in its TIMSS and PIRLS questionnaires for 4th grade students in order to assist them in estimating the number of books they have at home, which likely decreased measurement error and could bias estimates of books achievement gap trends upward. However, as we have seen, books achievement gaps also increased substantially for 8th grade and 15-year-old students, with no added drawings.

The one consistent exception to the increasing SES achievement gaps across all test instruments in Appendix B is trends estimated from PISA parent education, which are usually negative. This is especially surprising since 8th grade TIMSS tests a similar population in similar subjects and shows large increases in parent education achievement gaps. In the Appendix B results, it is difficult to discern whether the discrepancy between PISA and TIMSS 8th grade is the result of differences in samples, in test instruments, or in the measurement of SES. Table H3 estimates trends in parent education and household books achievement gaps for math and science in a constant sample of 42 countries that participated in at least two cycles each of PISA and TIMSS 8th grade. Trends in parent education gaps are close to 0 for PISA but are large and positive for TIMSS. Although trends in household books gaps are also larger for TIMSS than for PISA, they are nevertheless still large, positive, and significant for PISA. Wald tests for joint null hypotheses that trends are equal for PISA and TIMSS gaps can be rejected for both parent education and books gaps. Correlations between country-specific random cohort slopes for PISA and TIMSS gaps are weak for parent education but very strongly positive (over 0.90) for books gaps. These comparisons suggest that the difference in trends for PISA and TIMSS may be attributable not only to differences in test instruments and target populations but also likely due to differences in the measurement of SES, particularly parent education.

Table H3. Comparison of Trends in PISA and TIMSS 8th Grade 90/10 Parent Education and Household Books Gaps

|  | Math | Science |
|---|---|---|
| Parent education intercept × PISA | 0.720 *** | 0.713 *** |
|  | (0.033) | (0.031) |
| Parent education intercept × TIMSS | 0.765 *** | 0.738 *** |
|  | (0.032) | (0.034) |
| Household books intercept × PISA | 0.917 *** | 0.898 *** |
|  | (0.038) | (0.039) |
| Household books intercept × TIMSS | 0.799 *** | 0.800 *** |
|  | (0.037) | (0.044) |
| Level 2 - Study-years |  |  |
| Cohort birth year × Parent education × PISA | 0.001 | 0.000 |
|  | (0.002) | (0.002) |
| Cohort birth year × Parent education × TIMSS | 0.010 *** | 0.011 *** |
|  | (0.002) | (0.002) |
| Cohort birth year × Books × PISA | 0.005 ** | 0.007 *** |
|  | (0.002) | (0.002) |
| Cohort birth year × Books × TIMSS | 0.014 *** | 0.018 *** |
|  | (0.003) | (0.003) |
| N (Level 1 - gaps) | 711 | 711 |
| N (Level 2 - study-years) | 358 | 358 |
| N (Level 3 - countries) | 42 | 42 |
| p-value for H0: Cohort × Education × PISA = Cohort × Education × TIMSS | 0.003 | <0.001 |
| p-value for H0: Cohort × Books × PISA = Cohort × Books × TIMSS | 0.001 | <0.001 |
| Correlation between country random cohort slopes for PISA and TIMSS education gaps | -0.077 | 0.267 |
| Correlation between country random cohort slopes for PISA and TIMSS books gaps | 0.941 | 0.924 |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: To avoid very long computation times, all models in this table specify known level 1 error variances estimated using conventional non-bootstrap formulas and omit error covariances. This simplified specification appears to produce very similar results to models using bootstrapped error variances and covariances (see Appendix L).

Table H4 compares the parent education item wording from the student questionnaire for PISA and the student/parent questionnaires for the IEA studies (TIMSS and PIRLS).[7] TIMSS has a single education item for each parent; PISA has two items (schooling and higher education). TIMSS lists education levels in ascending order; PISA lists them in descending order. TIMSS includes an option for "I don't know"; PISA does not.[8] The highest two educational categories in TIMSS are "<ISCED Level 5A, first degree>" (i.e., BA) and "Beyond <ISCED Level 5A, first degree>" (i.e., MA, PhD, and professional degrees); the highest two categories in PISA are "<ISCED level 5A>" (i.e., BA, MA, and professional degrees) and

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[7] Beginning in 2015, TIMSS updated its parent education item wordings to reflect the new ISCED 2011 scheme. PISA had not yet made any update in its 2015 cycle.

[8] "I don't know" responses in TIMSS were treated as missing data and imputed as part of the multiple imputation procedure. TIMSS 8th grade still shows large and significant increases in parent education achievement gaps using unimputed data with listwise deletion of "I don't know" responses and other missing parent education data.

"<ISCED level 6>" (i.e., PhD).[9] Perhaps as a result of some or all of these differences, a substantially larger share of students select highest two categories (BA or above) in PISA than in TIMSS. Of the 29 countries participating in both PISA 2012 and TIMSS 2011 8th grade, 20 countries had a higher share of students reporting BA or above in PISA than TIMSS, by an average of 6 percentage points. For example, in Australia, 31% of TIMSS 8th grade 2011 students report BA or more, while 43% of PISA 2012 students report BA or more; in Finland, 42% of TIMSS 8th grade 2011 students report BA or more, while 55% of PISA 2012 students report BA or more. This pattern is in addition to the general increase in the share of students in the highest education categories seen across all datasets due to educational upgrading in the parents' generations. The larger share of students in the highest category in PISA means that the achievement at the 90th percentile of parent education is estimated with more error. This means that parent education achievement gaps may be underestimated in more recent years, and consequently that gap trends may be underestimated.

Table H4. Comparison of PISA and TIMSS parent education questionnaire wording

| PISA 2009-2015 student questionnaire | TIMSS 2003-2011 8th grade student questionnaire, TIMSS 2011 4th grade parent questionnaire, PIRLS 2006-2011 parent questionnaire |
|---|---|
| Q14. What is the <highest level of schooling> completed by your mother? <br> • <ISCED level 3A> <br> • <ISCED level 3B, 3C> <br> • <ISCED level 2> <br> • <ISCED level 1> <br> • She did not complete <ISCED level 1> <br> Q15. Does your mother have any of the following qualifications? <br> • <ISCED level 6> <br> • <ISCED level 5A> <br> • <ISCED level 5B> <br> • <ISCED level 4> | 6A. What is the highest level of education completed by your mother <or stepmother or female guardian>? <br> • Some <ISCED Level 1 or 2> or did not go to school <br> • <ISCED Level 2> <br> • <ISCED Level 3> <br> • <ISCED Level 4> <br> • <ISCED Level 5B> <br> • <ISCED Level 5A, first degree> <br> • Beyond <ISCED Level 5A, first degree> <br> • I don't know |

Models in the main text attempt to account for differences in variable quality across different studies or variables that may confound trend estimates. This is done by controlling for four SES variable quality measures (parent vs. student reporting, number of categories, 20% of more students in the bottom category, and 20% or more students in the top category). Model 1B in Table H5 omits these controls (Model 1 is reproduced from the main text for comparison). Results for Model 1B are nearly identical to those for Model 1. The trend in parent education

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[9] Due to likely inaccuracies in students' reports of very high levels of parental education in PISA, in all analyses, I combine the highest two educational categories in PISA: ISCED 5A (BA, MA, and professional degrees) and ISCED 6 (PhD). Very large shares of students in PISA report at least one parent holding a PhD (over 10% in some countries), far outnumbering national statistics on PhD attainment. When the ISCED 5A and 6 categories are kept separate, results are broadly similar, but PISA parent education gaps are even smaller than when the categories are combined, due to the large number of low-achieving students reporting parents holding PhDs.

achievement gaps increases very slightly, from 0.007 SD of achievement per year to 0.008 SD of achievement per year. Thus, the trend reported in the main text controlling for SES variable quality measures is a slightly more conservative estimate of the trend. As in the main text, a Wald joint test of the hypothesis that the trends are equal across all three SES variables cannot be rejected (p = 0.11).

Table H5. Comparison of Trends in Gaps, Omitting SES Variable Quality Measures

|  | (1) with SES quality measures | | | (1B) no SES quality measures | | |
|---|---|---|---|---|---|---|
|  | coef | (se) |  | coef | (se) |  |
| Parent education gaps intercept | 1.032 | (0.030) | *** | 0.989 | (0.028) | *** |
| Parent occupation gaps intercept | 0.958 | (0.030) | *** | 0.927 | (0.028) | *** |
| Household books gaps intercept | 1.299 | (0.041) | *** | 1.234 | (0.039) | *** |
| Level 1 - Gaps |  |  |  |  |  |  |
| Subject (ref=Reading) |  |  |  |  |  |  |
| Math | 0.020 | (0.007) | ** | 0.020 | (0.007) | ** |
| Science | 0.034 | (0.005) | *** | 0.034 | (0.005) | *** |
| SES variable quality measures |  |  |  |  |  |  |
| Parent-reported × Parent education | 0.132 | (0.030) | *** |  |  |  |
| Parent-reported × Parent occupation | 0.075 | (0.025) | ** |  |  |  |
| Parent-reported × Books | -0.039 | (0.029) |  |  |  |  |
| Number of categories (centered at 7) | 0.003 | (0.003) |  |  |  |  |
| ≥ 20% in bottom category | -0.065 | (0.021) | ** |  |  |  |
| ≥ 20% in top category | -0.135 | (0.013) | *** |  |  |  |
| Level 2 - Study-years |  |  |  |  |  |  |
| Age at testing (ref=14) |  |  |  |  |  |  |
| Age 10 at testing | -0.170 | (0.024) | *** | -0.136 | (0.022) | *** |
| Age 15 at testing | -0.024 | (0.020) |  | -0.034 | (0.020) | + |
| Cohort birth year × Parent education | 0.007 | (0.001) | *** | 0.008 | (0.001) | *** |
| Cohort birth year × Parent occupation | 0.007 | (0.001) | *** | 0.007 | (0.001) | *** |
| Cohort birth year × Books | 0.008 | (0.001) | *** | 0.008 | (0.001) | *** |
| Random effects |  |  |  |  |  |  |
| *Level 2 - Residual variance between studies in…* |  |  |  |  |  |  |
| Parent education intercepts | 0.03736 |  |  | 0.04505 |  |  |
| Parent occupation intercepts | 0.02322 |  |  | 0.02435 |  |  |
| Books intercepts | 0.03698 |  |  | 0.04404 |  |  |
| *Level 3 - Residual variance between countries in…* |  |  |  |  |  |  |
| Parent education intercepts | 0.05426 |  |  | 0.05021 |  |  |
| Parent occupation intercepts | 0.05227 |  |  | 0.04877 |  |  |
| Books intercepts | 0.11590 |  |  | 0.11679 |  |  |
| Parent education cohort slopes | 0.00004 |  |  | 0.00004 |  |  |
| Parent occupation cohort slopes | 0.00003 |  |  | 0.00003 |  |  |
| Books cohort slopes | 0.00007 |  |  | 0.00007 |  |  |
| N (Level 1 - gaps) | 5541 |  |  | 5541 |  |  |
| N (Level 2 - study-years) | 1026 |  |  | 1026 |  |  |
| N (Level 3 - countries) | 100 |  |  | 100 |  |  |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: All models in this table specify known level 1 error variances and covariances, estimated via bootstrapping, consistent with models in the main text of the paper.

Since the 90/10 percentile method may not perform as well when more than 20% of observations are in the bottom or top SES category (Reardon 2011a), the models in the main paper include two dummy variables indicating whether 20% or more of students fall into the bottom category (14.8% of country-study gaps) or the top category (38.7% of country-study gaps). Model 1C in Table H6 omits these gaps altogether (Model 1 is reproduced from the main text for comparison). Compared to Model 1, the sample size of Model 1C is substantially reduced—by more than half at level 1 due to the large shares of gaps with many students in the bottom or especially top categories, as reported above. While the estimates of the parent education and books gap trends remain relatively similar, the parent education gap trend is substantially reduced (from 0.007 SD of achievement per year to 0.004 SD per year). This is because parent education gaps are dropped for nearly all recent studies in high-income countries. I believe the estimate of the parent education gap trend in the main paper, while imprecise, is more accurate than the one in Model 1C omitting over half of the available data. Nevertheless, the estimate of the parent education gap trend is still positive and highly significant in Model 1C.

Table H6. Comparison of Trends in Gaps, Dropping Observations with 20% or More of Students in the Top or Bottom SES Variable Category

| | including all gaps (1) | | | dropping if ≥ 20% in top or bottom category (1C) | | |
|---|---|---|---|---|---|---|
| | coef | (se) | | coef | (se) | |
| Parent education gaps intercept | 1.032 | (0.030) | *** | 1.071 | (0.036) | *** |
| Parent occupation gaps intercept | 0.958 | (0.030) | *** | 0.998 | (0.037) | *** |
| Household books gaps intercept | 1.299 | (0.041) | *** | 1.366 | (0.041) | *** |
| Level 1 - Gaps | | | | | | |
|   Subject (ref=Reading) | | | | | | |
|     Math | 0.020 | (0.007) | ** | 0.004 | (0.010) | |
|     Science | 0.034 | (0.005) | *** | 0.034 | (0.006) | *** |
|   SES variable quality measures | | | | | | |
|     Parent-reported × Parent education | 0.132 | (0.030) | *** | 0.129 | (0.045) | ** |
|     Parent-reported × Parent occupation | 0.075 | (0.025) | ** | 0.049 | (0.035) | |
|     Parent-reported × Books | -0.039 | (0.029) | | -0.137 | (0.037) | *** |
|     Number of categories (centered at 7) | 0.003 | (0.003) | | 0.000 | (0.003) | |
|     ≥ 20% in bottom category | -0.065 | (0.021) | ** | | | |
|     ≥ 20% in top category | -0.135 | (0.013) | *** | | | |
| Level 2 - Study-years | | | | | | |
|   Age at testing (ref=14) | | | | | | |
|     Age 10 at testing | -0.170 | (0.024) | *** | -0.181 | (0.035) | *** |
|     Age 15 at testing | -0.024 | (0.020) | | 0.013 | (0.029) | |
|   Cohort birth year × Parent education | 0.007 | (0.001) | *** | 0.004 | (0.002) | ** |
|   Cohort birth year × Parent occupation | 0.007 | (0.001) | *** | 0.006 | (0.001) | *** |
|   Cohort birth year × Books | 0.008 | (0.001) | *** | 0.009 | (0.002) | *** |
| Random effects | | | | | | |
| *Level 2 - Residual variance between studies in…* | | | | | | |
|   Parent education intercepts | 0.03736 | | | 0.07398 | | |
|   Parent occupation intercepts | 0.02322 | | | 0.03327 | | |
|   Books intercepts | 0.03698 | | | 0.02828 | | |
| *Level 3 - Residual variance between countries in…* | | | | | | |
|   Parent education intercepts | 0.05426 | | | 0.04230 | | |
|   Parent occupation intercepts | 0.05227 | | | 0.05367 | | |
|   Books intercepts | 0.11590 | | | 0.09283 | | |
|   Parent education cohort slopes | 0.00004 | | | 0.00002 | | |
|   Parent occupation cohort slopes | 0.00003 | | | 0.00003 | | |
|   Books cohort slopes | 0.00007 | | | 0.00005 | | |
| N (Level 1 - gaps) | 5541 | | | 2577 | | |
| N (Level 2 - study-years) | 1026 | | | 820 | | |
| N (Level 3 - countries) | 100 | | | 91 | | |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: All models in this table specify known level 1 error variances and covariances, estimated via bootstrapping, consistent with models in the main text of the paper.

In summary, it is likely that changing measurement error in SES confounds estimates of true global changes in SES achievement gaps. But it is not clear that measurement error in SES is uniformly declining over time, as error may be increasing for some variables in some recent tests (i.e., PISA parent education). This would lead gap increases to be *under-* rather than overestimated, making the main trend estimates more conservative. To the extent that the reliability of SES has improved, this increase would have to be very large to fully account for the increase in SES achievement gaps. There is no direct evidence for changes in measurement error of SES, either in the data used in this study or in published research. Future research should identify older datasets containing both children's and parents' reports of family SES characteristics in order to examine whether reliability in student reporting may have increased over time and the size of this possible increase.

## J. Trend Models Run Separately by SES Variable

The models in the main text of the paper pool all gap types, regardless of which of the three SES variables was used to calculate gaps (parent education, parent occupation, or household books). Table J1 reports gap trends models run separately by the SES variable from which gaps were calculated. Results are similar to those for pooled models reported in the main text of the paper, though both the intercepts and the cohort birth year slopes are slightly smaller for all three SES variables in these models than in the main text. The discrepancy is mainly due to differences in coefficients for control variables (e.g., subject, SES quality measures, age), which are allowed to vary by SES variable in these models rather than constrained to take the same value for all gap types in the pooled models. However, it is not clear that the different results for control variables for different SES variables are meaningful, as each model in Table J1 includes a somewhat different set of study-years and countries. (Running the models only on country-study-years where all three SES variables are available would reduce the sample by 75%, eliminating a large amount of valuable information.)

In some cases, the differences in coefficients for control variables across models using different SES variables do appear meaningful. For example, the math coefficient in Table 1 in the main text of the paper is 0.02, meaning that, on average when pooling all gap types, math achievement gaps are larger than reading achievement gaps (the reference category) by about 0.02 SDs. In Table J1, the math coefficient when predicting parent education gaps is about 0.06, when predicting parent occupation gaps is about 0.03, and when predicting books gaps is about 0. It is easy to see why household books might truly have a stronger relationship with reading achievement compared to other SES variable-subject pairs. However, these coefficients may still overstate (or understate) the discrepancies across SES variables, as they are based on different samples. More importantly, though, the goal of the present study is not to describe the specific processes by which each type of achievement gap is generated, but instead to examine whether *trends* in gaps across cohorts are similar enough that they could plausibly be driven by the same underlying process of growing SES achievement gaps, broadly defined, and if so, to pool all gap types to obtain the most accurate estimate of that trend. Thus, the pooled models reported in the main text are preferred over those in Table J1, as they give the best summary of trends in SES achievement gaps, using the most complete data available.

The discrepancy between the results in Table J1 and those in the main text is also in small part due to the inability of the separate models to account for error covariances among different gap types, as the multivariate variance-known models do. However, this issue does not affect

results substantially, as the results in Appendix L show that pooled models omitting error covariances produce very similar results to models incorporating error covariances.

Table J1. Coefficients from Hierarchical Growth Models Predicting Achievement Gaps between 90th and 10th Percentiles of SES, Run Separately by SES Variable

| | Parent education | | Parent occupation | | Books | |
|---|---|---|---|---|---|---|
| Intercept | 1.016 *** | 1.054 *** | 0.921 *** | 0.903 *** | 1.211 *** | 1.200 *** |
| | (0.031) | (0.033) | (0.023) | (0.037) | (0.040) | (0.061) |
| **Level 1 - Gaps** | | | | | | |
| Subject (ref=Reading): | | | | | | |
| Math | 0.054 *** | 0.055 *** | 0.025 *** | 0.025 *** | -0.004 | -0.004 |
| | (0.010) | (0.010) | (0.007) | (0.007) | (0.009) | (0.009) |
| Science | 0.054 *** | 0.054 *** | 0.017 *** | 0.017 *** | 0.041 *** | 0.041 *** |
| | (0.008) | (0.008) | (0.005) | (0.005) | (0.007) | (0.007) |
| SES variable quality measures | | | | | | |
| Parent-reported | | 0.051 | | -0.051 + | | 0.096 *** |
| | | (0.045) | | (0.028) | | (0.024) |
| Number of categories (centered at 7) | | -0.002 | | 0.040 ** | | -0.045 |
| | | (0.003) | | (0.014) | | (0.028) |
| ≥ 20% in bottom category | | -0.117 ** | | 0.066 | | -0.120 ** |
| | | (0.044) | | (0.087) | | (0.037) |
| ≥ 20% in top category | | -0.101 *** | | -0.096 ** | | -0.107 *** |
| | | (0.020) | | (0.030) | | (0.023) |
| **Level 2 - Study-years** | | | | | | |
| Age at testing (ref=14) | | | | | | |
| Age 10 at testing | -0.003 | -0.060 | 0.017 | -0.003 | -0.215 *** | -0.247 *** |
| | (0.030) | (0.054) | (0.019) | (0.027) | (0.028) | (0.028) |
| Age 15 at testing | -0.154 *** | -0.112 *** | | | 0.105 *** | 0.127 ** |
| | (0.025) | (0.024) | | | (0.025) | (0.039) |
| Cohort birth year | 0.006 *** | 0.006 *** | 0.004 *** | 0.005 *** | 0.008 *** | 0.006 *** |
| | (0.001) | (0.001) | (0.001) | (0.001) | (0.001) | (0.002) |
| Level 2 residual variance in intercepts | 0.03390 | 0.03194 | 0.01852 | 0.01687 | 0.03395 | 0.03243 |
| Level 3 residual variance in intercepts | 0.04977 | 0.05164 | 0.03793 | 0.03616 | 0.11276 | 0.10128 |
| Level 3 residual variance in cohort slopes | 0.00005 | 0.00005 | 0.00004 | 0.00004 | 0.00008 | 0.00007 |
| N (Level 1 - gaps) | 1916 | 1916 | 1405 | 1405 | 2189 | 2189 |
| N (Level 2 - study-years) | 852 | 852 | 590 | 590 | 993 | 993 |
| N (Level 3 - countries) | 94 | 94 | 82 | 82 | 100 | 100 |

+ p<.1, * p<.05, ** p<.01, *** p < .001.
Note: In models predicting achievement gaps based on parent occupation, the dummy variable for age 15 is omitted, as there are very few age 14 assessments that collected parent occupation. Therefore, in the parent occupation models, the reference category is assessments of students who are age 14 *or* 15. All models in this table specify known level 1 error variances and covariances among gaps from different test subjects (but not different SES variables, as they are in separate models), estimated via bootstrapping.

## K. SES achievement gap trends by age and subject

Table K1 reports gap trends models estimating different cohort birth year trends by level of schooling (primary versus secondary), rather than by SES variable as in the main text models. SES achievement gaps have increased substantially for both primary school students (0.007 SD per year) and secondary school students (0.009 SD per year). Both estimates are statistically significant (p<0.001). Gaps for secondary school students tend to be slightly larger than those for primary school students, and gaps for secondary students have also increased somewhat more than those for primary students. However, the primary and secondary school trend estimates are similar in size; a joint test of the null hypothesis that the two cohort slopes are equal cannot be

rejected (p=0.148). Within countries, trends in gaps for primary and secondary school students tend to be similar: the correlation between country random cohort slopes for primary and secondary school gaps is positive and moderately strong (r=0.472). As differences in primary and secondary trends are small and comparing trends across age groups is not of substantive interest in the current project, I consider it justifiable to pool gaps from primary and secondary students in the main text models.

Table K2 reports gap trends models estimating different cohort birth year trends by test subject (math, science, or reading), rather than by SES variable as in the main text models. SES achievement gaps have increased substantially in reading tests (0.012 SD per year), math tests (0.010 SD per year), and science tests (0.012 SD per year). All three estimates are statistically significant (p<0.001). In the earliest birth cohorts, SES gaps in math achievement tended to be somewhat larger than those in reading or science achievement. However, SES gaps in reading and science achievement have increased at a slightly faster pace than those in math achievement, resulting in SES gaps that are very similar in size across the three subjects in the most recent cohorts. Yet the trend estimates for all three test subjects are similar in size; a joint test of the null hypothesis that the three cohort slopes are equal is only marginally significant (p=0.074). Within countries, trends in gaps based on different test subjects are very similar: pairwise correlations between country random slopes for different test subjects range from 0.785 to 0.854. As differences in trends by test subject are small and comparing trends across subjects is not of substantive interest in the current project, I consider it justifiable to pool gaps from all three test subjects in the main text models.

Table K1. Coefficients from Hierarchical Growth Models Predicting Achievement Gaps between 90th and 10th Percentiles of SES, Estimating Different Trends by Age (Primary or Secondary School)

| | coef | (se) | |
|---|---|---|---|
| Primary school gaps intercept | 0.988 | (0.068) | *** |
| Secondary school gaps intercept | 1.079 | (0.032) | *** |
| Level 1 - Gaps | | | |
| Subject (ref=Reading): | | | |
| Math | 0.021 | (0.007) | ** |
| Science | 0.035 | (0.005) | *** |
| SES variable (ref=Parent education) | | | |
| Parent occupation | -0.005 | (0.021) | |
| Household books | 0.214 | (0.035) | *** |
| SES variable quality measures | | | |
| Parent-reported | 0.021 | (0.063) | |
| Parent-reported × Parent occupation | -0.072 | (0.052) | |
| Parent-reported × Books | 0.002 | (0.071) | |
| Number of categories (centered at 7) | -0.013 | (0.005) | ** |
| ≥ 20% in bottom category | -0.212 | (0.038) | *** |
| ≥ 20% in top category | -0.035 | (0.022) | |
| Level 2 - Study-years | | | |
| Age 15 at testing | -0.185 | (0.029) | *** |
| Primary × Parent occupation | -0.040 | (0.046) | |
| Primary × Books | -0.112 | (0.065) | + |
| Cohort birth year × Primary | 0.007 | (0.002) | *** |
| Cohort birth year × Secondary | 0.009 | (0.001) | *** |
| Random effects | | | |
| *Residual variance between studies* | 0.02562 | | |
| *Residual variance between countries in…* | | | |
| Primary school intercepts | 0.05685 | | |
| Secondary school intercepts | 0.06327 | | |
| Primary school cohort slopes | 0.00003 | | |
| Secondary school cohort slopes | 0.00005 | | |
| N (Level 1 - gaps) | 5541 | | |
| N (Level 2 - study-years) | 1026 | | |
| N (Level 3 - countries) | 100 | | |
| p-value for H0: Cohort × Primary = Cohort × Secondary | 0.148 | | |
| Correlation between country random cohort slopes for primary and secondary gaps | 0.472 | | |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: All models in this table specify known level 1 error variances and covariances, estimated via bootstrapping, consistent with models in the main text of the paper.

Table K2. Coefficients from Hierarchical Growth Models Predicting Achievement Gaps between 90th and 10th Percentiles of SES, Estimating Different Trends by Test Subject (Math, Science, or Reading)

| | coef | (se) | |
|---|---|---|---|
| Reading gaps intercept | 1.025 | (0.032) | *** |
| Math gaps intercept | 1.037 | (0.031) | *** |
| Science gaps intercept | 1.047 | (0.032) | *** |
| Level 1 - Gaps | | | |
| SES variable (ref=Parent education) | | | |
| Parent occupation | -0.073 | (0.021) | *** |
| Household books | 0.362 | (0.033) | *** |
| SES variable quality measures | | | |
| Parent-reported | -0.034 | (0.052) | |
| Parent-reported × Parent occupation | -0.053 | (0.027) | + |
| Parent-reported × Books | -0.227 | (0.036) | *** |
| Number of categories (centered at 7) | -0.005 | (0.005) | |
| ≥ 20% in bottom category | -0.202 | (0.037) | *** |
| ≥ 20% in top category | -0.102 | (0.018) | *** |
| Level 2 - Study-years | | | |
| Age at testing (ref=14) | | | |
| Age 10 at testing | -0.135 | (0.031) | *** |
| Age 15 at testing | -0.018 | (0.019) | |
| Cohort birth year × Reading | 0.012 | (0.001) | *** |
| Cohort birth year × Math | 0.010 | (0.001) | *** |
| Cohort birth year × Science | 0.012 | (0.001) | *** |
| Random effects | | | |
| *Level 2 - Residual variance between studies in…* | | | |
| Reading intecepts | 0.03293 | | |
| Math intercepts | 0.03216 | | |
| Science intercepts | 0.03707 | | |
| *Level 3 - Residual variance between countries in…* | | | |
| Reading intercepts | 0.06441 | | |
| Math intercepts | 0.06027 | | |
| Science intercepts | 0.07121 | | |
| Reading cohort slopes | 0.00006 | | |
| Math cohort slopes | 0.00003 | | |
| Science cohort slopes | 0.00004 | | |
| N (Level 1 - gaps) | 5541 | | |
| N (Level 2 - study-years) | 1026 | | |
| N (Level 3 - countries) | 100 | | |
| p-value for H0: Cohort × Reading = Cohort × Math = Cohort × Science | 0.074 | | |
| Correlation between country random cohort slopes for reading and math gaps | 0.808 | | |
| Correlation between country random cohort slopes for reading and science gaps | 0.854 | | |
| Correlation between country random cohort slopes for math and science gaps | 0.785 | | |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: All models in this table specify known level 1 error variances and covariances, estimated via bootstrapping, consistent with models in the main text of the paper.

**L. Specification of trend model**

The coefficients for some control variables were omitted from the tables in the main text due to space constraints. Tables L1 and L2 report all coefficients from Table 2. Table L1 reports those from Model 3 (with country region interactions), and Table L2 reports all coefficients from Models 4-7.

Figure 3 in the main text showed estimated trends in SES achievement gaps for only 24 selected countries due to space constraints. Figures L1-L3 show estimated gap trends for a larger set of countries. Trend lines are derived from shrunken empirical Bayes estimates from Model 5 (Table 2 and Table L2). Countries with fewer than 7 study-years or whose data span fewer than 15 cohort birth years are excluded from the figures, as their trends cannot be precisely estimated. This leaves 55 displayed in the figures (of 100 countries in the full sample). For the remaining 45 countries with less available data, true country-specific gap trends are very uncertain. However, these countries still contribute to the estimation of the average global gap trend in the hierarchical growth curve models.

Table L1. All Coefficients from Model 3 Predicting Achievement Gaps between 90th and 10th Percentiles of SES with Interactions by Country Region

| | (3) Region Interactions | | |
|---|---|---|---|
| | coef | (se) | |
| Parent education gaps intercept | 1.126 | (0.027) | *** |
| Parent occupation gaps intercept | 1.111 | (0.029) | *** |
| Household books gaps intercept | 1.563 | (0.038) | *** |
| Level 1 - Gaps | | | |
|   Subject (ref=Reading) | | | |
|     Math | 0.020 | (0.007) | ** |
|     Science | 0.034 | (0.005) | *** |
|   SES variable quality measures | | | |
|     Parent-reported × Parent education | 0.111 | (0.031) | *** |
|     Parent-reported × Parent occupation | 0.069 | (0.024) | ** |
|     Parent-reported × Books | -0.018 | (0.026) | |
|     Number of categories (centered at 7) | 0.002 | (0.003) | |
|     ≥ 20% in bottom category | -0.045 | (0.020) | * |
|     ≥ 20% in top category | -0.147 | (0.013) | *** |
| Level 2 - Study-years | | | |
|   Age at testing (ref=14) | | | |
|     Age 10 at testing | -0.171 | (0.024) | *** |
|     Age 15 at testing | -0.028 | (0.020) | |
|   Cohort birth year | 0.008 | (0.001) | *** |
| Level 3 - Countries | | | |
|   Region (ref=Western) × Intercept interactions | | | |
|     Sub-Saharan Africa × Parent education | -0.307 | (0.089) | *** |
|     Sub-Saharan Africa × Parent occupation | -0.373 | (0.113) | ** |
|     Sub-Saharan Africa × Books | -0.707 | (0.143) | *** |
|     East Asia & Pacific × Parent education | -0.160 | (0.082) | + |
|     East Asia & Pacific × Parent occupation | -0.316 | (0.074) | *** |
|     East Asia & Pacific × Books | -0.402 | (0.097) | *** |
|     Middle East & N. Africa × Parent education | -0.248 | (0.078) | ** |
|     Middle East & N. Africa × Parent occupation | -0.317 | (0.069) | *** |
|     Middle East & N. Africa × Books | -0.731 | (0.082) | *** |
|     E. Europe & CIS × Parent education | -0.057 | (0.062) | |
|     E. Europe & CIS × Parent occupation | -0.091 | (0.051) | + |
|     E. Europe & CIS × Books | -0.162 | (0.064) | * |
|     Latin America & Caribbean × Parent education | 0.077 | (0.051) | |
|     Latin America & Caribbean × Parent occupation | -0.022 | (0.050) | |
|     Latin America & Caribbean × Books | -0.220 | (0.070) | ** |
|   Region (ref=Western) × Cohort interactions | | | |
|     Sub-Saharan Africa × Cohort | 0.004 | (0.004) | |
|     East Asia & Pacific × Cohort | -0.001 | (0.002) | |
|     Middle East & N. Africa × Cohort | -0.001 | (0.004) | |
|     E. Europe & CIS × Cohort | -0.001 | (0.002) | |
|     Latin America & Caribbean × Cohort | -0.009 | (0.004) | * |
| Random Effects | | | |
| *Level 2 - Residual variance between studies in…* | | | |
|   Parent education intercepts | 0.03785 | | |
|   Parent occupation intercepts | 0.02247 | | |
|   Books intercepts | 0.03825 | | |
| *Level 3 - Residual variance between countries in…* | | | |
|   Parent education intercepts | 0.04199 | | |
|   Parent occupation intercepts | 0.03086 | | |
|   Books intercepts | 0.05514 | | |
|   Cohort slopes | 0.00003 | | |
| N (Level 1 - gaps) | 5541 | | |
| N (Level 2 - study-years) | 1026 | | |
| N (Level 3 - countries) | 100 | | |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: See Table A1 for country regions.

Table L2. All Coefficients from Models 4-7 Predicting Achievement Gaps between 90th, 50th, and 10th Percentiles of SES with Interactions by Country Income Level

| | (4) Income Interaction | | (5) Quadratic | | (6) 90/50 Gap | | (7) 50/10 Gap | |
|---|---|---|---|---|---|---|---|---|
| | coef (se) | | coef (se) | | coef (se) | | coef (se) | |
| Parent education gaps intercept | 1.074 (0.035) *** | | 1.066 (0.037) *** | | 0.548 (0.019) *** | | 0.514 (0.022) *** | |
| Parent occupation gaps intercept | 1.054 (0.038) *** | | 1.046 (0.042) *** | | 0.564 (0.020) *** | | 0.481 (0.022) *** | |
| Household books gaps intercept | 1.434 (0.047) *** | | 1.426 (0.049) *** | | 0.622 (0.022) *** | | 0.806 (0.029) *** | |
| Level 1 - Gaps | | | | | | | | |
| Subject (ref=Reading) | | | | | | | | |
|   Math | 0.020 (0.007) ** | | 0.020 (0.007) ** | | 0.021 (0.004) *** | | 0.009 (0.004) * | |
|   Science | 0.034 (0.005) *** | | 0.034 (0.005) *** | | 0.022 (0.003) *** | | 0.013 (0.003) *** | |
| SES variable quality measures | | | | | | | | |
|   Parent-reported × Parent education | 0.113 (0.031) *** | | 0.112 (0.031) *** | | 0.105 (0.018) *** | | 0.016 (0.022) | |
|   Parent-reported × Parent occupation | 0.072 (0.024) ** | | 0.071 (0.024) ** | | 0.056 (0.017) ** | | 0.019 (0.017) | |
|   Parent-reported × Books | -0.018 (0.027) | | -0.018 (0.027) | | 0.140 (0.022) *** | | -0.143 (0.022) *** | |
|   Number of categories (centered at 7) | 0.002 (0.003) | | 0.002 (0.003) | | 0.002 (0.002) | | 0.001 (0.002) | |
|   ≥ 20% in bottom category | -0.061 (0.021) ** | | -0.061 (0.021) ** | | -0.055 (0.017) ** | | -0.009 (0.020) | |
|   ≥ 20% in top category | -0.148 (0.013) *** | | -0.147 (0.013) *** | | -0.053 (0.011) *** | | -0.086 (0.010) *** | |
| Level 2 - Study-years | | | | | | | | |
| Age at testing (ref=14) | | | | | | | | |
|   Age 10 at testing | -0.169 (0.024) *** | | -0.163 (0.024) *** | | -0.120 (0.016) *** | | -0.041 (0.014) ** | |
|   Age 15 at testing | -0.023 (0.020) | | -0.021 (0.020) | | -0.044 (0.013) ** | | 0.020 (0.012) + | |
| Cohort birth year | 0.007 (0.001) *** | | 0.007 (0.001) *** | | 0.001 (0.001) | | 0.006 (0.001) *** | |
| Cohort birth year$^2$ | | | 0.00005 (0.00005) | | | | | |
| Level 3 - Countries | | | | | | | | |
| Mid/low-income country × Intercept interactions | | | | | | | | |
|   Mid/low-income country × Parent education | -0.060 (0.048) | | -0.040 (0.048) | | 0.007 (0.025) | | -0.062 (0.029) * | |
|   Mid/low-income country × Parent occupation | -0.166 (0.047) *** | | -0.143 (0.050) ** | | -0.007 (0.027) | | -0.162 (0.028) *** | |
|   Mid/low-income country × Books | -0.251 (0.066) *** | | -0.231 (0.066) *** | | -0.054 (0.031) + | | -0.195 (0.042) *** | |
| Mid/low-income country × Cohort | 0.002 (0.002) | | 0.001 (0.002) | | 0.003 (0.001) * | | -0.001 (0.001) | |
| Mid/low-income country × Cohort$^2$ | | | -0.00022 (0.00010) * | | | | | |
| Random Effects | | | | | | | | |
| *Level 2 - Residual variance between studies in…* | | | | | | | | |
|   Parent education intercepts | 0.03822 | | 0.03677 | | 0.01671 | | 0.01362 | |
|   Parent occupation intercepts | 0.02302 | | 0.02154 | | 0.00877 | | 0.00634 | |
|   Books intercepts | 0.03839 | | 0.03734 | | 0.01715 | | 0.01594 | |
| *Level 3 - Residual variance between countries in…* | | | | | | | | |
|   Parent education intercepts | 0.05256 | | 0.05227 | | 0.01294 | | 0.01832 | |
|   Parent occupation intercepts | 0.04810 | | 0.05095 | | 0.01448 | | 0.01655 | |
|   Books intercepts | 0.10677 | | 0.10841 | | 0.02128 | | 0.03906 | |
|   Cohort slopes | 0.00003 | | 0.00003 | | 0.00001 | | 0.00001 | |
| N (Level 1 - gaps) | 5541 | | 5541 | | 5541 | | 5541 | |
| N (Level 2 - study-years) | 1026 | | 1026 | | 1026 | | 1026 | |
| N (Level 3 - countries) | 100 | | 100 | | 100 | | 100 | |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: "Middle/low income" countries had GDPs per capita of less than $6000 in 1980 (the reference category is high-income countries; see Table A1 for coding).

Figure L1. Estimated Quadratic Trends in 90/10 SES Achievement Gaps, Selected Countries

(Number of Study-Years in Parentheses)
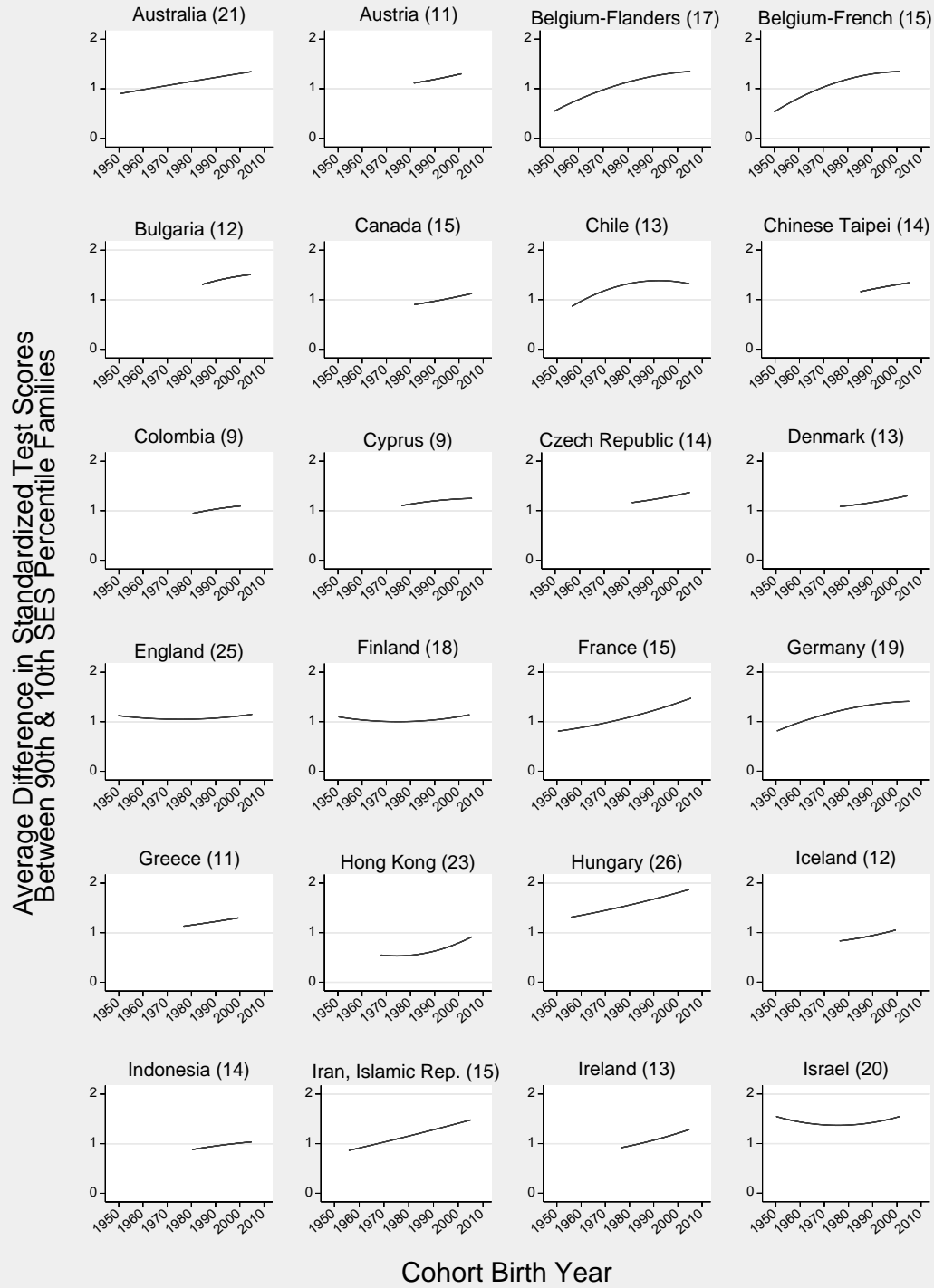
Notes: Countries sorted alphabetically. Trend lines are derived from shrunken empirical Bayes estimates from Model 5 (Table 2 and Table L2). Fixed values for control variables: SES=parent education, subject=math, all others=0 or reference category.

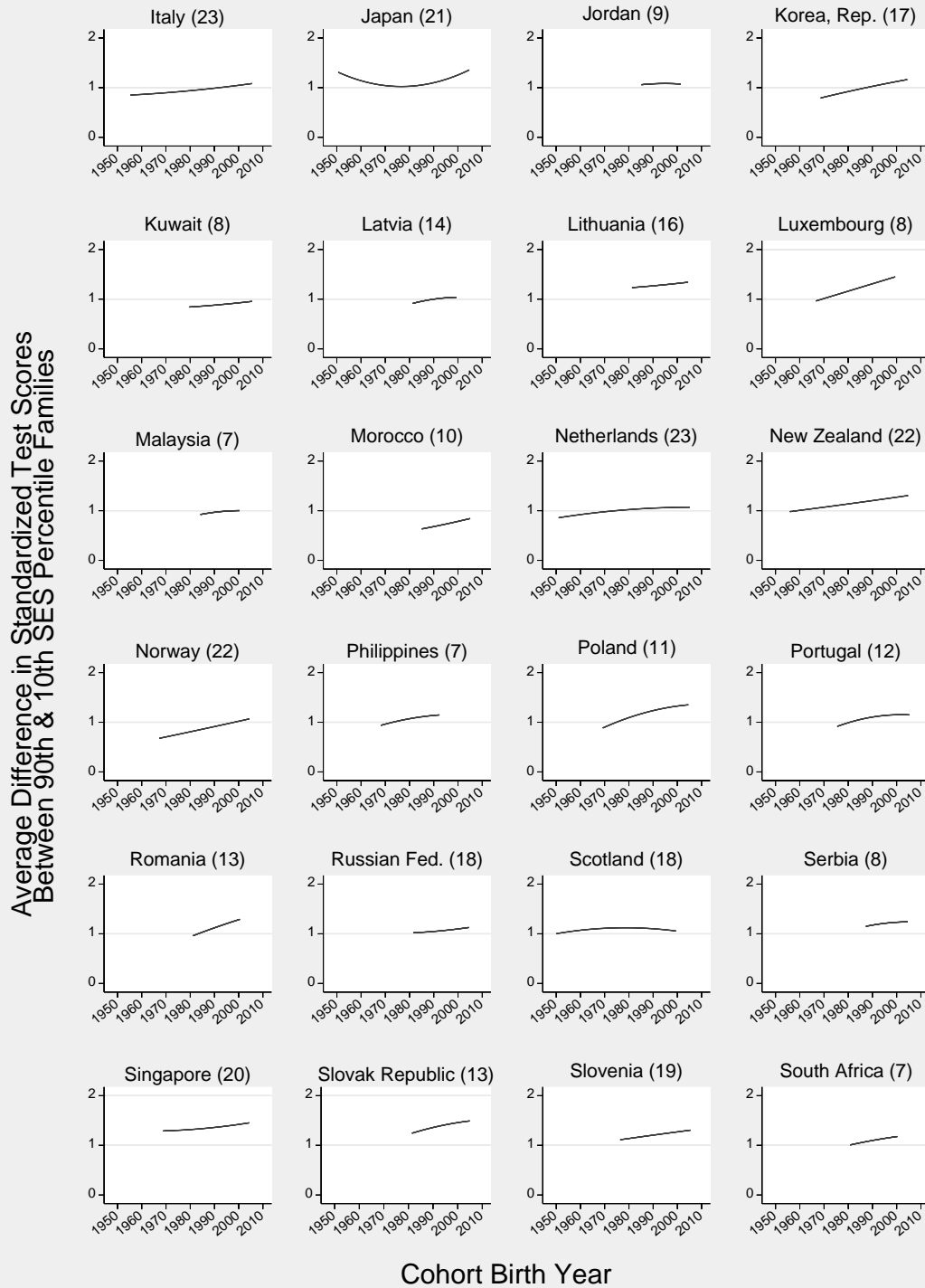Figure L2. Estimated Quadratic Trends in 90/10 SES Achievement Gaps, Selected Countries

(Number of Study-Years in Parentheses)

Notes: Countries sorted alphabetically. Trend lines are derived from shrunken empirical Bayes estimates from Model 5 (Table 2 and Table L2). Fixed values for control variables: SES=parent education, subject=math, all others=0 or reference category.

Figure L3. Estimated Quadratic Trends in 90/10 SES Achievement Gaps, Selected Countries
(Number of Study-Years in Parentheses)

Notes: Countries sorted alphabetically. Trend lines are derived from shrunken empirical Bayes estimates from Model 5 (Table 2 and Table L2). Fixed values for control variables: SES=parent education, subject=math, all others=0 or reference category.
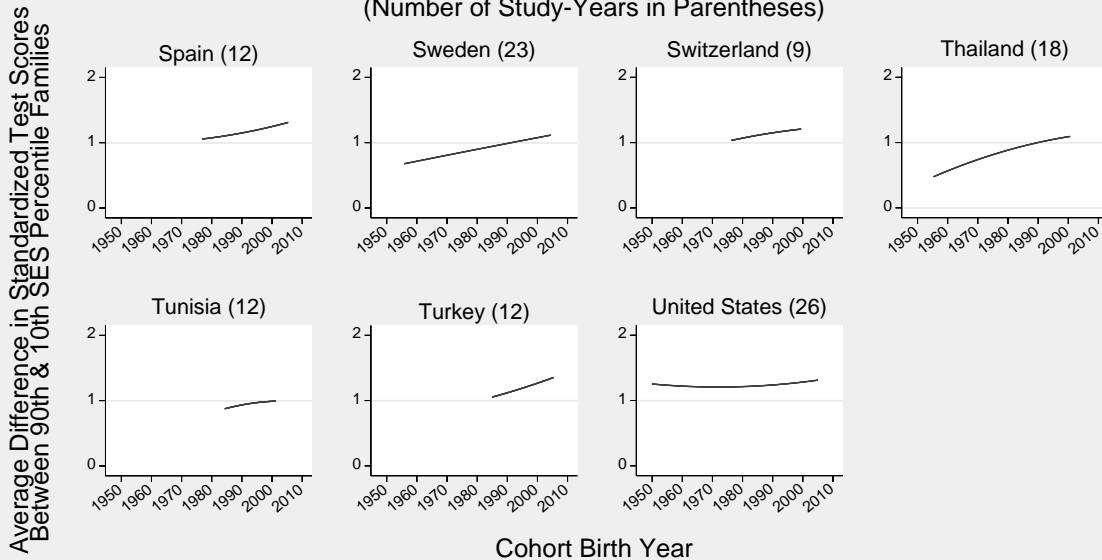
Tables L3 and L4 report coefficients from six alternate specifications of the main gap trend model from the main text of the paper (Model 2 in Table 1). The first column of table L3 replicates Model 2 for comparison. All models in the main paper are multivariate variance-known models that specify the known the Level 1 (within-study) variance structure as equal to the estimated error variance-covariance matrix among different gap types in each country-study. Thus, the models take into account that the errors of different gap types within the same country-study are not independent. Error variances and covariances are estimated via bootstrapping. Due to the very long computation time of 1000 bootstrap estimates of 5541 SES achievement gaps, Models 2B and 2C examine the robustness of the results to excluding error covariances from level 1 while still specifying known error variances estimated via bootstrapping (Model 2B) and to specifying known error variances estimated using the conventional normality-assumption-based formula (following Reardon (2011b)) rather than bootstrapping (Model 2C). Results for both models are nearly identical to Model 2. Model 2D omits any known error variance from level 1 and freely estimates level 1 variance within the hierarchical model, meaning that the model no longer gives greater weight to more precisely-estimated gaps. Results again are nearly identical. As omitting error covariances and using conventional non-bootstrap error variances appears to produce very similar results, some subsequent models (2E and 2F) and several others throughout the appendices use this alternate level 1 variance specification to avoid excessive computation time.

Table L3. Coefficients from Hierarchical Growth Models Predicting Achievement Gaps between 90th and 10th Percentiles of Parent Education (Variety of Trend Model Specifications)

| | (2) | (2B) | (2C) | (2D) |
|---|---|---|---|---|
| | Main text (with BS error covariance) | With BS, no error covariance | Non-BS, no error covariance | No weights |
| Parent education intercept | 1.039 *** | 1.034 *** | 1.044 *** | 1.035 *** |
| | (0.030) | (0.030) | (0.031) | (0.030) |
| Parent occupation intercept | 0.964 *** | 0.959 *** | 0.962 *** | 0.955 *** |
| | (0.030) | (0.030) | (0.031) | (0.029) |
| Household books intercept | 1.294 *** | 1.291 *** | 1.301 *** | 1.292 *** |
| | (0.041) | (0.041) | (0.040) | (0.040) |
| **Level 1 - Gaps** | | | | |
| Subject (ref=Reading): | | | | |
| Math | 0.020 ** | 0.018 * | 0.006 | 0.014 + |
| | (0.007) | (0.007) | (0.008) | (0.007) |
| Science | 0.034 *** | 0.031 *** | 0.026 *** | 0.033 *** |
| | (0.005) | (0.006) | (0.007) | (0.006) |
| SES variable quality measures | | | | |
| Parent-reported × Parent education | 0.112 *** | 0.111 *** | 0.080 ** | 0.117 *** |
| | (0.031) | (0.030) | (0.027) | (0.031) |
| Parent-reported × Parent occupation | 0.073 ** | 0.069 ** | 0.051 * | 0.074 ** |
| | (0.024) | (0.023) | (0.021) | (0.025) |
| Parent-reported × Books | -0.017 | -0.017 | -0.067 ** | -0.001 |
| | (0.026) | (0.026) | (0.023) | (0.029) |
| Number of categories (centered at 7) | 0.002 | 0.003 | 0.003 | 0.003 |
| | (0.003) | (0.003) | (0.003) | (0.003) |
| ≥ 20% in bottom category | -0.063 ** | -0.062 ** | -0.066 ** | -0.072 *** |
| | (0.021) | (0.021) | (0.021) | (0.020) |
| ≥ 20% in top category | -0.146 *** | -0.143 *** | -0.136 *** | -0.144 *** |
| | (0.013) | (0.013) | (0.012) | (0.013) |
| **Level 2 - Study-years** | | | | |
| Age at testing (ref=14) | | | | |
| Age 10 at testing | -0.168 *** | -0.163 *** | -0.147 *** | -0.166 *** |
| | (0.024) | (0.024) | (0.023) | (0.024) |
| Age 15 at testing | -0.023 | -0.021 | -0.027 | -0.022 |
| | (0.020) | (0.020) | (0.019) | (0.020) |
| Cohort birth year | 0.007 *** | 0.007 *** | 0.008 *** | 0.007 *** |
| | (0.001) | (0.001) | (0.001) | (0.001) |
| *Level 1 - Residual variance between gaps* | | | | 0.00774 |
| *Level 2 - Residual variance between studies in…* | | | | |
| Parent education gaps | 0.03831 | 0.03972 | 0.02719 | 0.04135 |
| Parent occupation gaps | 0.02284 | 0.02309 | 0.01877 | 0.02297 |
| Household books gaps | 0.03823 | 0.03881 | 0.03187 | 0.04111 |
| *Level 3 - Residual variance between countries in…* | | | | |
| Parent education gaps | 0.05362 | 0.05231 | 0.05517 | 0.05121 |
| Parent occupation gaps | 0.05330 | 0.05305 | 0.05490 | 0.05326 |
| Household books gaps | 0.12149 | 0.12167 | 0.12270 | 0.11918 |
| Cohort slopes | 0.00003 | 0.00003 | 0.00003 | 0.00003 |
| N (Level 1 - gaps) | 5541 | 5541 | 5541 | 5541 |
| N (Level 2 - study-years) | 1026 | 1026 | 1027 | 1027 |
| N (Level 3 - countries) | 100 | 100 | 100 | 100 |

+ p<.1, * p<.05, ** p<.01, *** p < .001

Model 2E uses gaps computed without multiple imputation of missing student-level data; listwise deletion of student observations with missing data on any variables is used instead. Level 1 (gaps) sample sizes are slightly smaller. Point estimates for coefficients are similar, though the standard errors on the intercepts and cohort trend increase substantially, and the cohort trend loses significance. These changes are due to increased heterogeneity across countries in gap trends estimated when dropping missing data (though note that imputation models were estimated separately for each country-study).

Model 2F uses gaps for which the 90th and 10th percentiles of each SES categorical variable were interpolated from linear rather than cubic weighted least squares models. Level 1 (gaps) and Level 2 (study-years) sample sizes are larger because a slightly larger number of gaps can be estimated reliably from these simpler models. The cohort birth year coefficient is slightly larger (i.e., the estimated increase in SES achievement gaps is larger) because the additional observations are very small gaps in early years. Cubic models are retained as the preferred models, as the trend estimates are more conservative, cubic gap functions allow more flexibility in the shape of the relationship between SES and achievement, and for comparability with Reardon (2011b).

Model 2G uses gaps between the 75th and 25th percentiles rather than the 90th and 10th percentiles of each SES variable. As expected, the intercepts, cohort birth year coefficient, and coefficients for most control variables are smaller than in models predicting 90/10 gaps, as 75/10 gaps are smaller in magnitude. However, the size of the cohort birth year coefficient is still substantial (0.005 SD per year) and highly significant. These results demonstrate that the finding of increasing SES achievement gaps is robust to using an alternate gap measure that may be more precisely estimated, as the 75th and 25th percentiles are less likely to be extrapolated outside the available SES data in high-income countries with large numbers of students in the top parent education and occupation categories. Additional models predicting 75/50 and 50/25 gaps (not shown) reveal that the estimated cohort coefficient for 75/50 gaps (0.002) is slightly smaller than the estimated cohort coefficient for 50/25 gaps (0.003). Both coefficients are highly significant ($p < 0.001$). Thus, the finding that gaps have increased more between the middle and bottom of the SES distribution than between the middle and top of the SES distribution—modeled in the main paper text using 50/10 and 90/50 gaps—is also robust to using alternative percentiles of the SES distribution. The 90/10, 90/50, and 50/10 gaps are still retained as the preferred results in the main paper text for comparability with Reardon (2011b).

Model 2H omits the random effect for the cohort birth year slope and instead estimates a fixed cohort slope for all countries. Results are very similar. The random cohort slope is retained as the preferred model, as a chi-square test shows the variance in the random cohort slopes to be significantly different from 0 in the main models and all other trend model specifications.

Table L4. Coefficients from Hierarchical Growth Models Predicting Achievement Gaps between 90th and 10th (or 75th and 25th) Percentiles of Parent Occupation (Variety of Trend Model Specifications)

| | (2E) No imputation | (2F) Linear gaps | (2G) 75/25 | (2H) Fixed slope |
|---|---|---|---|---|
| Parent education intercept | 1.063 *** | 0.992 *** | 0.561 *** | 1.036 *** |
| | (0.225) | (0.029) | (0.019) | (0.031) |
| Parent occupation intercept | 1.000 *** | 0.923 *** | 0.555 *** | 0.961 *** |
| | (0.178) | (0.029) | (0.021) | (0.030) |
| Household books intercept | 1.311 *** | 1.276 *** | 0.794 *** | 1.290 *** |
| | (0.174) | (0.040) | (0.022) | (0.042) |
| **Level 1 - Gaps** | | | | |
| Subject (ref=Reading): | | | | |
| Math | -0.007 | 0.013 + | 0.002 | 0.020 ** |
| | (0.010) | (0.007) | (0.006) | (0.007) |
| Science | 0.023 ** | 0.020 *** | 0.005 | 0.034 *** |
| | (0.009) | (0.005) | (0.006) | (0.005) |
| SES variable quality measures | | | | |
| Parent-reported × Parent education | 0.084 ** | 0.067 * | 0.001 | 0.118 *** |
| | (0.031) | (0.026) | (0.002) | (0.031) |
| Parent-reported × Parent occupation | 0.013 | 0.046 * | 0.078 *** | 0.080 ** |
| | (0.025) | (0.021) | (0.023) | (0.025) |
| Parent-reported × Books | -0.075 ** | -0.087 *** | -0.005 | -0.012 |
| | (0.027) | (0.022) | (0.016) | (0.027) |
| Number of categories (centered at 7) | 0.002 | 0.004 | -0.076 *** | 0.002 |
| | (0.006) | (0.002) | (0.017) | (0.003) |
| ≥ 20% in bottom category | -0.049 * | -0.037 + | 0.004 | -0.065 ** |
| | (0.023) | (0.019) | (0.015) | (0.021) |
| ≥ 20% in top category | -0.129 *** | -0.104 *** | -0.069 *** | -0.148 *** |
| | (0.016) | (0.011) | (0.009) | (0.013) |
| **Level 2 - Study-years** | | | | |
| Age at testing (ref=14) | | | | |
| Age 10 at testing | -0.142 *** | -0.136 *** | -0.126 *** | -0.165 *** |
| | (0.024) | (0.022) | (0.015) | (0.024) |
| Age 15 at testing | -0.031 | -0.020 | -0.010 | -0.021 |
| | (0.022) | (0.019) | (0.012) | (0.020) |
| Cohort birth year | 0.007 | 0.008 *** | 0.005 *** | 0.007 *** |
| | (0.006) | (0.001) | (0.001) | (0.001) |
| *Level 1 - Residual variance between gaps* | | | | |
| *Level 2 - Residual variance between studies in…* | | | | |
| Parent education gaps | 0.02784 | 0.02385 | 0.01715 | 0.04126 |
| Parent occupation gaps | 0.01356 | 0.01764 | 0.00615 | 0.02701 |
| Household books gaps | 0.03188 | 0.02749 | 0.01261 | 0.04209 |
| *Level 3 - Residual variance between countries in…* | | | | |
| Parent education gaps | 0.05741 | 0.05079 | 0.01342 | 0.05533 |
| Parent occupation gaps | 0.04812 | 0.05529 | 0.01952 | 0.05526 |
| Household books gaps | 0.12037 | 0.12524 | 0.03298 | 0.12563 |
| Cohort slopes | 0.00003 | 0.00003 | 0.00002 | |
| N (Level 1 - gaps) | 5521 | 5584 | 5199 | 5541 |
| N (Level 2 - study-years) | 1026 | 1033 | 1015 | 1026 |
| N (Level 3 - countries) | 100 | 100 | 99 | 100 |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: To avoid very long computation times, Models 2E, 2F, and 2G specify known level 1 error variances estimated using conventional non-bootstrap formulas and omit error covariances. This simplified specification appears to produce very similar results to models using bootstrapped error variances and covariances (see Table L3). Model 2H uses bootstrapped error variances and covariances.

# M. Specification of multivariate model

The coefficients for some control variables were omitted from the tables in the main text due to space constraints. Table M1 reports all coefficients from Table 3 (with time-varying country covariates). Additionally, Table M1 reports two further models predicting 90/50 and 50/10 SES achievement gaps (Models 9 and 10, respectively). The coefficients for the time-varying country covariates are similar in direction whether predicting 90/50 or 50/10 gaps, but the magnitude and significance of coefficients tends to be larger when predicting 90/50 gaps than 50/10 gaps. Additionally, by comparing the level 2 residual variances for these full models to reduced models that include study fixed effects and controls but no country covariates (not shown), we can compare how well the covariates account for changes in gaps at the top relative to the bottom of the SES distribution. Compared to a reduced model predicting 90/50 gaps, the country covariates in Model 9 explain an additional 5%, 10%, and 6% of the within-country variance in 90/50 gaps based on parent education, occupation and books. In contrast, the covariates in Model 10 explain only an additional 0.2%, 0%, and 5% of the within-country variance in 50/10 gaps based on parent education, occupation, and books. This is consistent with the fact that most of the covariate coefficients are larger and more significant when predicting 90/50 than 50/10 gaps, and indicates that the country characteristics examined in this study do a better job of explaining changes in gaps at the top of the SES distribution than at the bottom. As 50/10 gaps have increased greatly in most countries (more than 90/50 gaps have increased), a large share of the within-country variance in 50/10 gaps is explained by the study-year fixed effects. Net of this large global secular increase, however, the country covariates do not tell us much about which countries experience larger or smaller increases in 50/10 gaps.

Table M1. Unstandardized Coefficients from Hierarchical Growth Models Predicting Achievement Gaps between $90^{th}$, $50^{th}$, and $10^{th}$ Percentiles of SES

| | (8) 90/10 Gap | | (9) 90/50 Gap | | (10) 50/10 Gap | |
|---|---|---|---|---|---|---|
| | coef | (se) | coef | (se) | coef | (se) |
| Parent education gaps intercept | 0.970 | (0.046) *** | 0.552 | (0.029) *** | 0.401 | (0.032) *** |
| Parent occupation gaps intercept | 0.969 | (0.049) *** | 0.584 | (0.030) *** | 0.375 | (0.032) *** |
| Household books gaps intercept | 1.405 | (0.057) *** | 0.645 | (0.033) *** | 0.753 | (0.034) *** |
| **Level 1 - Gaps** | | | | | | |
| Subject controls (ref=Reading) | | | | | | |
| Math | 0.020 | (0.007) ** | 0.021 | (0.004) *** | 0.007 | (0.004) * |
| Science | 0.033 | (0.006) *** | 0.022 | (0.003) *** | 0.012 | (0.003) *** |
| SES variable quality measures | | | | | | |
| Parent-reported × Parent education | 0.023 | (0.039) | 0.071 | (0.022) ** | -0.049 | (0.028) + |
| Parent-reported × Parent occupation | -0.026 | (0.032) | 0.002 | (0.021) | -0.038 | (0.022) + |
| Parent-reported × Books | -0.128 | (0.039) ** | 0.084 | (0.025) *** | -0.212 | (0.026) *** |
| Number of categories (centered at 7) | 0.000 | (0.003) | 0.001 | (0.002) | 0.001 | (0.002) |
| ≥ 20% in bottom category | -0.055 | (0.023) * | -0.047 | (0.019) * | 0.001 | (0.021) |
| ≥ 20% in top category | -0.149 | (0.015) *** | -0.061 | (0.012) *** | -0.080 | (0.011) *** |
| **Level 2 - Study-years** | | | | | | |
| Study fixed effects (ref=TIMSS 2003 Grade 8) | yes | | yes | | yes | |
| School enrollment (proportion) | 0.486 | (0.107) *** | 0.276 | (0.067) *** | 0.203 | (0.068) ** |
| Immigrant background (proportion) | 0.226 | (0.250) | 0.130 | (0.145) | 0.090 | (0.134) |
| GDP per capita (logged) | 0.055 | (0.059) | 0.024 | (0.049) | 0.039 | (0.039) |
| Income inequality (Gini) | -1.913 | (0.887) * | -1.380 | (0.476) ** | -0.554 | (0.489) |
| Mid/low-income country × Income inequality | 2.539 | (1.129) * | 1.687 | (0.628) ** | 0.753 | (0.650) |
| Age when tracking begins | -0.037 | (0.016) * | -0.023 | (0.011) * | -0.015 | (0.008) + |
| Private school enrollment (proportion) | 0.240 | (0.249) | 0.147 | (0.142) | 0.070 | (0.159) |
| Expecting higher education (proportion) | -0.029 | (0.094) | -0.043 | (0.062) | 0.017 | (0.046) |
| **Level 3 - Countries** | | | | | | |
| Mid/low-income country × Intercept interactions | | | | | | |
| Mid/low-income country × Parent education | 0.192 | (0.063) ** | 0.071 | (0.033) * | 0.138 | (0.037) *** |
| Mid/low-income country × Parent occupation | 0.060 | (0.058) | 0.056 | (0.035) | 0.017 | (0.033) |
| Mid/low-income country × Books | -0.082 | (0.068) | -0.016 | (0.040) | -0.059 | (0.043) |
| Mean school enrollment | 0.640 | (0.317) * | 0.139 | (0.179) | 0.531 | (0.162) ** |
| Mean proportion immigrant background | 0.134 | (0.261) | -0.057 | (0.146) | 0.172 | (0.140) |
| Mean GDP per capita (logged) | 0.142 | (0.041) *** | 0.078 | (0.024) ** | 0.070 | (0.022) ** |
| Mean income inequality | 0.605 | (0.317) + | 0.594 | (0.169) *** | -0.008 | (0.174) |
| Mean age when tracking begins | -0.035 | (0.011) ** | -0.020 | (0.005) *** | -0.018 | (0.007) * |
| Mean private school enrollment | 0.043 | (0.099) | 0.006 | (0.058) | 0.029 | (0.057) |
| Mean proportion expecting higher education | -0.255 | (0.169) | -0.113 | (0.087) | -0.154 | (0.101) |
| **Random effects** | | | | | | |
| *Level 2 - Residual variance between studies in…* | | | | | | |
| Parent education gaps | 0.038 | | 0.017 | | 0.015 | |
| Parent occupation gaps | 0.021 | | 0.008 | | 0.006 | |
| Household books gaps | 0.037 | | 0.015 | | 0.014 | |
| *Level 3 - Residual variance between countries in…* | | | | | | |
| Parent education gaps | 0.028 | | 0.007 | | 0.010 | |
| Parent occupation gaps | 0.027 | | 0.009 | | 0.010 | |
| Household books gaps | 0.049 | | 0.014 | | 0.016 | |
| N (Level 1 - gaps) | 4604 | | 4604 | | 4604 | |
| N (Level 2 - study-years) | 855 | | 855 | | 855 | |
| N (Level 3 - countries) | 78 | | 78 | | 78 | |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Notes: "Middle/low income" countries had GDPs per capita of less than $6000 in 1980 (the reference category is high-income countries; see Appendix Table A1 for coding). All level 2 time-varying country covariates are mean-centered within countries, meaning results can be interpreted very similarly to a model with country fixed effects (as well as study-year fixed effects, included at level 2).

Tables M2 and M3 report intercorrelations among all country covariates, at the country-study-year level (n=855; Table M2) and at the country level (n=78; Table M3). None of the correlations is strong enough to cause concern about collinearity in the multivariate models.

Table M2. Correlations Among Time-Varying Country Covariates (n=855)

|  | School enrollment (proportion) | Immigrant background (proportion) | GDP per capita (logged) | Income inequality (Gini) | Age when tracking begins | Private school enrollment (proportion) | Expecting higher education (proportion) |
|---|---|---|---|---|---|---|---|
| School enrollment (proportion) | 1 | | | | | | |
| Immigrant background (proportion) | 0.258 | 1 | | | | | |
| GDP per capita (logged) | 0.587 | 0.441 | 1 | | | | |
| Income inequality (Gini) | -0.443 | -0.281 | -0.466 | 1 | | | |
| Age when tracking begins | 0.077 | 0.067 | 0.028 | 0.169 | 1 | | |
| Private school enrollment (proportion) | 0.022 | 0.165 | 0.135 | -0.052 | -0.214 | 1 | |
| Expecting higher education (proportion) | -0.063 | -0.060 | -0.086 | 0.432 | 0.307 | 0.017 | 1 |

Table M3. Correlations Among Country-Level Mean Variables (N=78)

|  | Mean school enrollment | Mean proportion immigrant background | Mean GDP per capita (logged) | Mean income inequality | Mean age when tracking begins | Mean private school enrollment | Mean proportion expecting higher education |
|---|---|---|---|---|---|---|---|
| Mean school enrollment | 1.000 | | | | | | |
| Mean proportion immigrant background | 0.320 | 1 | | | | | |
| Mean GDP per capita (logged) | 0.647 | 0.475 | 1 | | | | |
| Mean income inequality | -0.520 | -0.388 | -0.457 | 1 | | | |
| Mean age when tracking begins | 0.089 | 0.045 | -0.015 | 0.082 | 1 | | |
| Mean private school enrollment | -0.009 | 0.089 | 0.151 | -0.025 | -0.262 | 1 | |
| Mean proportion expecting higher education | -0.288 | -0.251 | -0.398 | 0.438 | 0.265 | 0.074 | 1 |

Table M4 further examines possible collinearity among time-varying country covariates by reporting seven models, each predicting 90/10 SES achievement gaps from only one time-varying country covariate at a time. The coefficients for all time-varying country covariates are very similar when each is entered individually to in models where all seven covariates are entered together. Thus, collinearity among country covariates does not appear to influence the results.

In the main text of the paper, I examined whether certain types of countries tend to have greater increases in SES achievement gaps by interacting two country-level variables (region and income level) with the cohort birth year slope. Table M5 further explores whether cohort birth year slopes systematically vary by other available country-level variables. As the country-level sample size with complete data is only 78, I run each interaction model separately. The coefficients of interest are the 'Country variable × Cohort interactions' near the bottom of the table. All coefficients are close to 0 and non-significant.

Table M4. Unstandardized Coefficients from Hierarchical Growth Models Predicting Achievement Gaps between 90th and 10th Percentiles of SES, Entering Covariates One at a Time

| | Enrollment | | Immigration | | GDP per capita | | Income inequality | |
|---|---|---|---|---|---|---|---|---|
| | coef | (se) | coef | (se) | coef | (se) | coef | (se) |
| Parent education gaps intercept | 1.065 | (0.043) *** | 1.082 | (0.040) *** | 0.969 | (0.045) *** | 1.016 | (0.044) *** |
| Parent occupation gaps intercept | 1.066 | (0.047) *** | 1.083 | (0.046) *** | 0.970 | (0.049) *** | 1.015 | (0.046) *** |
| Household books gaps intercept | 1.498 | (0.057) *** | 1.514 | (0.055) *** | 1.401 | (0.055) *** | 1.449 | (0.056) *** |
| **Level 1 - Gaps** | | | | | | | | |
| Subject controls (ref=Reading) | | | | | | | | |
| Math | 0.020 | (0.007) ** | 0.020 | (0.007) ** | 0.020 | (0.007) ** | 0.020 | (0.007) ** |
| Science | 0.033 | (0.006) *** | 0.033 | (0.006) *** | 0.033 | (0.006) *** | 0.033 | (0.006) *** |
| SES variable quality measures | | | | | | | | |
| Parent-reported × Parent education | 0.027 | (0.039) | 0.027 | (0.039) | 0.025 | (0.039) | 0.020 | (0.039) |
| Parent-reported × Parent occupation | -0.022 | (0.033) | -0.024 | (0.032) | -0.025 | (0.032) | -0.024 | (0.032) |
| Parent-reported × Books | -0.123 | (0.039) ** | -0.124 | (0.040) ** | -0.127 | (0.039) ** | -0.136 | (0.040) *** |
| Number of categories (centered at 7) | 0.000 | (0.003) | 0.000 | (0.003) | 0.000 | (0.003) | -0.002 | (0.003) |
| ≥ 20% in bottom category | -0.059 | (0.022) ** | -0.055 | (0.022) * | -0.056 | (0.022) * | -0.057 | (0.022) ** |
| ≥ 20% in top category | -0.154 | (0.015) *** | -0.153 | (0.015) *** | -0.152 | (0.015) *** | -0.146 | (0.015) *** |
| **Level 2 - Study-years** | | | | | | | | |
| Study fixed effects (ref=TIMSS 2003 Grade 8) | yes | | yes | | yes | | yes | |
| School enrollment (proportion) | 0.505 | (0.110) *** | | | | | | |
| Immigrant background (proportion) | | | 0.313 | (0.301) | | | | |
| GDP per capita (logged) | | | | | 0.076 | (0.068) | | |
| Income inequality (Gini) | | | | | | | -1.623 | (0.838) + |
| Mid/low-income country × Income inequality | | | | | | | 2.072 | (1.078) + |
| **Level 3 - Countries** | | | | | | | | |
| Mid/low-income country × Intercept interactions | | | | | | | | |
| Mid/low-income country × Parent education | 0.021 | (0.061) | -0.038 | (0.050) | 0.168 | (0.062) ** | -0.070 | (0.063) |
| Mid/low-income country × Parent occupation | -0.110 | (0.056) * | -0.161 | (0.051) ** | 0.039 | (0.057) | -0.195 | (0.059) *** |
| Mid/low-income country × Books | -0.254 | (0.072) *** | -0.314 | (0.069) *** | -0.102 | (0.066) | -0.347 | (0.078) *** |
| Mean school enrollment | 0.583 | (0.291) * | | | | | | |
| Mean proportion immigrant background | | | 0.358 | (0.279) | | | | |
| Mean GDP per capita (logged) | | | | | 0.178 | (0.036) *** | | |
| Mean income inequality | | | | | | | 0.140 | (0.305) |
| **Random effects** | | | | | | | | |
| *Level 2 - Residual variance between studies in…* | | | | | | | | |
| Parent education gaps | 0.03892 | | 0.03980 | | 0.03957 | | 0.04075 | |
| Parent occupation gaps | 0.02241 | | 0.02219 | | 0.02206 | | 0.02300 | |
| Household books gaps | 0.03788 | | 0.03949 | | 0.04030 | | 0.04011 | |
| *Level 3 - Residual variance between countries in…* | | | | | | | | |
| Parent education gaps | 0.03905 | | 0.04544 | | 0.03646 | | 0.04400 | |
| Parent occupation gaps | 0.03498 | | 0.03694 | | 0.02962 | | 0.03736 | |
| Household books gaps | 0.07527 | | 0.08838 | | 0.06152 | | 0.08712 | |
| N (Level 1 - gaps) | 4604 | | 4604 | | 4604 | | 4604 | |
| N (Level 2 - study-years) | 855 | | 855 | | 855 | | 855 | |
| N (Level 3 - countries) | 78 | | 78 | | 78 | | 78 | |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Notes: "Middle/low income" countries had GDPs per capita of less than $6000 in 1980 (the reference category is high-income countries; see Appendix Table A1 for coding). All level 2 time-varying country covariates are mean-centered within countries, meaning results can be interpreted very similarly to a model with country fixed effects (as well as study-year fixed effects, included at level 2). All models in this table specify known level 1 error variances and covariances, estimated via bootstrapping, consistent with models in the main text of the paper.

Model M4 (cont.) Unstandardized Coefficients from Hierarchical Growth Models Predicting Achievement Gaps between 90th and 10th Percentiles of SES, Entering Covariates One at a Time

| | Tracking | | | Private | | | Higher ed. Expect. | | |
|---|---|---|---|---|---|---|---|---|---|
| | coef | (se) | | coef | (se) | | coef | (se) | |
| Parent education gaps intercept | 1.098 | (0.039) | *** | 1.094 | (0.040) | *** | 1.083 | (0.043) | *** |
| Parent occupation gaps intercept | 1.098 | (0.042) | *** | 1.095 | (0.043) | *** | 1.084 | (0.044) | *** |
| Household books gaps intercept | 1.531 | (0.052) | *** | 1.527 | (0.054) | *** | 1.515 | (0.052) | *** |
| Level 1 - Gaps | | | | | | | | | |
| Subject controls (ref=Reading) | | | | | | | | | |
| Math | 0.020 | (0.007) | ** | 0.020 | (0.007) | ** | 0.020 | (0.007) | ** |
| Science | 0.033 | (0.006) | *** | 0.033 | (0.006) | *** | 0.033 | (0.006) | *** |
| SES variable quality measures | | | | | | | | | |
| Parent-reported × Parent education | 0.026 | (0.039) | | 0.025 | (0.038) | | 0.027 | (0.038) | |
| Parent-reported × Parent occupation | -0.025 | (0.032) | | -0.026 | (0.032) | | -0.025 | (0.032) | |
| Parent-reported × Books | -0.126 | (0.040) | ** | -0.126 | (0.039) | ** | -0.125 | (0.039) | ** |
| Number of categories (centered at 7) | 0.000 | (0.003) | | 0.000 | (0.003) | | 0.000 | (0.003) | |
| ≥ 20% in bottom category | -0.052 | (0.022) | * | -0.054 | (0.022) | * | -0.056 | (0.022) | * |
| ≥ 20% in top category | -0.150 | (0.015) | *** | -0.152 | (0.015) | *** | -0.151 | (0.015) | *** |
| Level 2 - Study-years | | | | | | | | | |
| Study fixed effects (ref=TIMSS 2003 Grade 8) | yes | | | yes | | | yes | | |
| Age when tracking begins | -0.032 | (0.015) | * | | | | | | |
| Private school enrollment (proportion) | | | | 0.206 | (0.295) | | | | |
| Expecting higher education (proportion) | | | | | | | -0.054 | (0.101) | |
| Level 3 - Countries | | | | | | | | | |
| Mid/low-income country × Intercept interactions | | | | | | | | | |
| Mid/low-income country × Parent education | -0.068 | (0.045) | | -0.057 | (0.046) | | -0.037 | (0.054) | |
| Mid/low-income country × Parent occupation | -0.194 | (0.045) | *** | -0.181 | (0.045) | *** | -0.160 | (0.049) | *** |
| Mid/low-income country × Books | -0.346 | (0.063) | *** | -0.332 | (0.066) | *** | -0.311 | (0.067) | *** |
| Mean age when tracking begins | -0.028 | (0.013) | * | | | | | | |
| Mean private school enrollment | | | | 0.213 | (0.108) | * | | | |
| Mean proportion expecting higher education | | | | | | | -0.311 | (0.176) | + |
| Random effects | | | | | | | | | |
| *Level 2 - Residual variance between studies in…* | | | | | | | | | |
| Parent education gaps | 0.03950 | | | 0.03947 | | | 0.03939 | | |
| Parent occupation gaps | 0.02165 | | | 0.02223 | | | 0.02209 | | |
| Household books gaps | 0.03962 | | | 0.04009 | | | 0.04010 | | |
| *Level 3 - Residual variance between countries in…* | | | | | | | | | |
| Parent education gaps | 0.04093 | | | 0.04483 | | | 0.04541 | | |
| Parent occupation gaps | 0.03644 | | | 0.03729 | | | 0.03597 | | |
| Household books gaps | 0.08150 | | | 0.08966 | | | 0.08300 | | |
| N (Level 1 - gaps) | 4604 | | | 4604 | | | 4604 | | |
| N (Level 2 - study-years) | 855 | | | 855 | | | 855 | | |
| N (Level 3 - countries) | 78 | | | 78 | | | 78 | | |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Notes: "Middle/low income" countries had GDPs per capita of less than $6000 in 1980 (the reference category is high-income countries; see Appendix Table A1 for coding). All level 2 time-varying country covariates are mean-centered within countries, meaning results can be interpreted very similarly to a model with country fixed effects (as well as study-year fixed effects, included at level 2). All models in this table specify known level 1 error variances and covariances, estimated via bootstrapping, consistent with models in the main text of the paper.

Table M5. Unstandardized Coefficients from Hierarchical Growth Curve Models Predicting Achievement Gaps between 90th and 10th Percentiles, Adding Interactions by Country Variables

| | Enrollment | | Immigration | | GDP per capita | | Income inequality | |
|---|---|---|---|---|---|---|---|---|
| | coef | (se) | coef | (se) | coef | (se) | coef | (se) |
| Parent education gaps intercept | 1.064 | (0.031) *** | 1.067 | (0.033) *** | 1.063 | (0.031) *** | 1.066 | (0.033) *** |
| Parent occupation gaps intercept | 0.993 | (0.029) *** | 0.995 | (0.030) *** | 0.988 | (0.028) *** | 0.994 | (0.031) *** |
| Household books gaps intercept | 1.331 | (0.039) *** | 1.338 | (0.044) *** | 1.322 | (0.036) *** | 1.335 | (0.043) *** |
| Level 1 - Gaps | | | | | | | | |
|   Subject controls (ref=Reading) | | | | | | | | |
|     Math | 0.020 | (0.007) ** | 0.020 | (0.007) ** | 0.020 | (0.007) ** | 0.020 | (0.007) ** |
|     Science | 0.033 | (0.006) *** | 0.033 | (0.006) *** | 0.033 | (0.006) *** | 0.033 | (0.006) *** |
|   SES variable quality measures | | | | | | | | |
|     Parent-reported × Parent education | 0.109 | (0.038) ** | 0.108 | (0.037) ** | 0.107 | (0.037) ** | 0.109 | (0.037) ** |
|     Parent-reported × Parent occupation | 0.055 | (0.027) * | 0.055 | (0.028) * | 0.053 | (0.027) + | 0.055 | (0.028) * |
|     Parent-reported × Books | -0.024 | (0.030) | -0.020 | (0.030) | -0.024 | (0.030) | -0.021 | (0.030) |
|     Number of categories (centered at 7) | 0.003 | (0.003) | 0.002 | (0.003) | 0.002 | (0.003) | 0.002 | (0.003) |
|     ≥ 20% in bottom category | -0.049 | (0.024) * | -0.061 | (0.025) * | -0.047 | (0.024) + | -0.055 | (0.024) * |
|     ≥ 20% in top category | -0.155 | (0.014) *** | -0.157 | (0.013) *** | -0.156 | (0.014) *** | -0.156 | (0.014) *** |
| Level 2 - Study-years | | | | | | | | |
|   Age at testing (ref=14) | | | | | | | | |
|     Age 10 at testing | -0.153 | (0.026) *** | -0.152 | (0.026) *** | -0.151 | (0.026) *** | -0.152 | (0.026) *** |
|     Age 15 at testing | -0.014 | (0.023) | -0.013 | (0.023) | -0.015 | (0.023) | -0.014 | (0.023) |
|   Cohort birth year | 0.007 | (0.001) *** | 0.007 | (0.001) *** | 0.007 | (0.001) *** | 0.007 | (0.001) *** |
| Level 3 - Countries | | | | | | | | |
|   Country variable × Intercept interactions | | | | | | | | |
|     Mean school enrollment × Parent education | 0.597 | (0.194) ** | | | | | | |
|     Mean school enrollment × Parent occupation | 0.810 | (0.252) ** | | | | | | |
|     Mean school enrollment × Books | 1.635 | (0.287) *** | | | | | | |
|     Mean proportion immigrant background × Parent education | | | 0.358 | (0.292) | | | | |
|     Mean proportion immigrant background × Parent occupation | | | 0.934 | (0.248) *** | | | | |
|     Mean proportion immigrant background × Books | | | 1.288 | (0.395) ** | | | | |
|     Mean GDP per capita (logged) × Parent education | | | | | 0.091 | (0.026) *** | | |
|     Mean GDP per capita (logged) × Parent occupation | | | | | 0.152 | (0.027) *** | | |
|     Mean GDP per capita (logged) × Books | | | | | 0.291 | (0.034) *** | | |
|     Mean income inequality × Parent education | | | | | | | -0.020 | (0.241) |
|     Mean income inequality × Parent occupation | | | | | | | -0.510 | (0.221) * |
|     Mean income inequality × Books | | | | | | | -1.326 | (0.318) *** |
|   Country variable × Cohort interactions | | | | | | | | |
|     Mean school enrollment × Cohort | 0.003 | (0.009) | | | | | | |
|     Mean proportion immigrant background × Cohort | | | 0.013 | (0.015) | | | | |
|     Mean GDP per capita (logged) × Cohort | | | | | 0.001 | (0.001) | | |
|     Mean income inequality × Cohort | | | | | | | -0.010 | (0.012) |
| Random Effects | | | | | | | | |
| *Level 2 - Residual variance between studies in...* | | | | | | | | |
|   Parent education intercepts | 0.04087 | | 0.04074 | | 0.04074 | | 0.04071 | |
|   Parent occupation intercepts | 0.02232 | | 0.02233 | | 0.02223 | | 0.02237 | |
|   Books intercepts | 0.04054 | | 0.04042 | | 0.04049 | | 0.04051 | |
| *Level 3 - Residual variance between countries in...* | | | | | | | | |
|   Parent education intercepts | 0.03693 | | 0.04246 | | 0.0372 | | 0.04387 | |
|   Parent occupation intercepts | 0.03238 | | 0.0371 | | 0.02761 | | 0.04022 | |
|   Books intercepts | 0.06803 | | 0.09794 | | 0.05272 | | 0.09185 | |
|   Cohort slopes | 0.00004 | | 0.00004 | | 0.00003 | | 0.00004 | |
| N (Level 1 - gaps) | 4604 | | 4604 | | 4604 | | 4604 | |
| N (Level 2 - study-years) | 855 | | 855 | | 855 | | 855 | |
| N (Level 3 - countries) | 78 | | 78 | | 78 | | 78 | |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: All models in this table specify known level 1 error variances and covariances, estimated via bootstrapping, consistent with models in the main text of the paper.

Table M5 (cont). Unstandardized Coefficients from Hierarchical Growth Curve Models Predicting Achievement Gaps between 90th and 10th Percentiles, Adding Interactions by Country Variables

| | Tracking | | | Private | | | Higher ed. Expect. | | |
|---|---|---|---|---|---|---|---|---|---|
| | coef | (se) | | coef | (se) | | coef | (se) | |
| Parent education gaps intercept | 1.068 | (0.032) | *** | 1.065 | (0.033) | *** | 1.065 | (0.033) | *** |
| Parent occupation gaps intercept | 0.995 | (0.031) | *** | 0.993 | (0.031) | *** | 0.993 | (0.031) | *** |
| Household books gaps intercept | 1.341 | (0.044) | *** | 1.338 | (0.045) | *** | 1.334 | (0.044) | *** |
| Level 1 - Gaps | | | | | | | | | |
|   Subject controls (ref=Reading) | | | | | | | | | |
|     Math | 0.020 | (0.007) | ** | 0.020 | (0.007) | ** | 0.020 | (0.007) | ** |
|     Science | 0.033 | (0.006) | *** | 0.033 | (0.006) | *** | 0.033 | (0.006) | *** |
|   SES variable quality measures | | | | | | | | | |
|     Parent-reported × Parent education | 0.108 | (0.037) | ** | 0.109 | (0.037) | ** | 0.108 | (0.037) | ** |
|     Parent-reported × Parent occupation | 0.055 | (0.028) | * | 0.055 | (0.027) | * | 0.054 | (0.028) | * |
|     Parent-reported × Books | -0.021 | (0.030) | | -0.021 | (0.030) | | -0.022 | (0.030) | |
|     Number of categories (centered at 7) | 0.002 | (0.003) | | 0.002 | (0.003) | | 0.002 | (0.003) | |
|     ≥ 20% in bottom category | -0.064 | (0.025) | ** | -0.063 | (0.024) | * | -0.058 | (0.025) | * |
|     ≥ 20% in top category | -0.156 | (0.014) | *** | -0.155 | (0.014) | *** | -0.155 | (0.014) | *** |
| Level 2 - Study-years | | | | | | | | | |
|   Age at testing (ref=14) | | | | | | | | | |
|     Age 10 at testing | -0.152 | (0.026) | *** | -0.153 | (0.026) | *** | -0.152 | (0.026) | *** |
|     Age 15 at testing | -0.014 | (0.023) | | -0.014 | (0.023) | | -0.014 | (0.023) | |
|   Cohort birth year | 0.007 | (0.001) | *** | 0.007 | (0.001) | *** | 0.007 | (0.001) | *** |
| Level 3 - Countries | | | | | | | | | |
|   Country variable × Intercept interactions | | | | | | | | | |
|     Mean age when tracking begins × Parent education | -0.044 | (0.013) | *** | | | | | | |
|     Mean age when tracking begins × Parent occupation | -0.027 | (0.014) | * | | | | | | |
|     Mean age when tracking begins × Books | -0.058 | (0.020) | ** | | | | | | |
|     Mean private school enrollment × Parent education | | | | 0.232 | (0.121) | + | | | |
|     Mean private school enrollment × Parent occupation | | | | 0.278 | (0.122) | * | | | |
|     Mean private school enrollment × Books | | | | 0.165 | (0.157) | | | | |
|     Mean proportion expecting higher education × Parent education | | | | | | | -0.269 | (0.192) | |
|     Mean proportion expecting higher education × Parent occupation | | | | | | | -0.615 | (0.152) | *** |
|     Mean proportion expecting higher education × Books | | | | | | | -1.167 | (0.279) | *** |
|     Mean age when tracking begins × Cohort | -0.001 | (0.001) | | | | | | | |
|     Mean private school enrollment × Cohort | | | | -0.003 | (0.007) | | | | |
|     Mean proportion expecting higher education × Cohort | | | | | | | -0.011 | (0.007) | + |
| Random Effects | | | | | | | | | |
| *Level 2 - Residual variance between studies in…* | | | | | | | | | |
|   Parent education intercepts | 0.04075 | | | 0.0407 | | | 0.04072 | | |
|   Parent occupation intercepts | 0.0224 | | | 0.02249 | | | 0.02227 | | |
|   Books intercepts | 0.04058 | | | 0.04035 | | | 0.04044 | | |
| *Level 3 - Residual variance between countries in…* | | | | | | | | | |
|   Parent education intercepts | 0.03695 | | | 0.04189 | | | 0.04211 | | |
|   Parent occupation intercepts | 0.04095 | | | 0.0403 | | | 0.0366 | | |
|   Books intercepts | 0.0959 | | | 0.10571 | | | 0.08689 | | |
|   Cohort slopes | 0.00003 | | | 0.00004 | | | 0.00003 | | |
| N (Level 1 - gaps) | 4604 | | | 4604 | | | 4604 | | |
| N (Level 2 - study-years) | 855 | | | 855 | | | 855 | | |
| N (Level 3 - countries) | 78 | | | 78 | | | 78 | | |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: All models in this table specify known level 1 error variances and covariances, estimated via bootstrapping, consistent with models in the main text of the paper.

The country covariate models in the main text of the paper enter a series of study-year dummies in order to account for secular trends across study-years. An alternative way to account for secular trends in SES achievement gaps would be to include the cohort birth year variable in the model rather than study-year fixed effects. Table M6 reports such a model (Model 8B), along with a model omitting country covariates but using the same analytic sample of 78 countries with complete covariate data (Model 2H), and the original version of Model 8 with study-year fixed effects from the main paper text for comparison. Model 2H shows that the cohort birth year trend is identical in the reduced analytic sample of countries with full covariate data to in the full sample of 100 countries. Comparing across Models 8B and 8, coefficents for time-varying country covariates are broadly similar. Some coefficients change direction, but only those that are not significantly different from 0 in either model. The strongest predictor of increasing SES achievement gaps remains increasing school enrollment. Model 8B including both country covariates and cohort birth year also allows us to examine a different question: to what extent do the time-varying country covariates explain the average global increase in SES achievement gaps? Comparing across Model 2H and Model 8B, the cohort birth year coefficient decreases from 0.007 to 0.001, an 86% reduction. Thus, the included country covariates do explain a great deal of the average global trend in SES achievement gaps. Model 8 is preferred over Model 8B because in addition to controlling for secular time trends in SES achievement gaps, it also accounts for possible differences across studies in the measurement error of SES achievement gaps (either due to unreliability or low quality of achievement or SES measures, as discussed in Appendices D and H). If we assume that all countries are likely to suffer from similar data reliability or quality issues in the same studies—a reasonable assumption because SES survey item wording is very similar across countries in a given year of a study—then these dummies control for study-specific biases in the estimation of SES achievement gaps. Thus, the coefficient estimates for covariates in Model 8 are less likely to be confounded by uneven quality of the outcome variablew than those in Model 8B.

Table M1. Comparison between Multivariate Models Predicting 90/10 Achievement Gaps from Country Covariates and Cohort Birth Year or Study-Year Dummies

| | (2H) Analytic sample | | (8B) Cohort year | | (8) Study-year FE | |
|---|---|---|---|---|---|---|
| | coef | (se) | coef | (se) | coef | (se) |
| Parent education gaps intercept | 1.099 | (0.030) *** | 0.978 | (0.039) *** | 0.970 | (0.046) *** |
| Parent occupation gaps intercept | 1.097 | (0.031) *** | 0.975 | (0.042) *** | 0.969 | (0.049) *** |
| Household books gaps intercept | 1.522 | (0.040) *** | 1.391 | (0.047) *** | 1.405 | (0.057) *** |
| **Level 1 - Gaps** | | | | | | |
| Subject controls (ref=Reading) | | | | | | |
| Math | 0.020 | (0.007) ** | 0.020 | (0.007) ** | 0.020 | (0.007) ** |
| Science | 0.033 | (0.006) *** | 0.033 | (0.006) *** | 0.033 | (0.006) *** |
| SES variable quality measures | | | | | | |
| Parent-reported × Parent education | 0.108 | (0.037) ** | 0.112 | (0.037) ** | 0.023 | (0.039) |
| Parent-reported × Parent occupation | 0.054 | (0.028) * | 0.061 | (0.027) * | -0.026 | (0.032) |
| Parent-reported × Books | -0.022 | (0.030) | -0.017 | (0.029) | -0.128 | (0.039) ** |
| Number of categories (centered at 7) | 0.002 | (0.003) | 0.002 | (0.003) | 0.000 | (0.003) |
| ≥ 20% in bottom category | -0.055 | (0.024) * | -0.057 | (0.025) * | -0.055 | (0.023) * |
| ≥ 20% in top category | -0.157 | (0.014) *** | -0.156 | (0.014) *** | -0.149 | (0.015) *** |
| **Level 2 - Study-years** | | | | | | |
| Age at testing (ref=14) | | | | | | |
| Age 10 at testing | -0.153 | (0.026) *** | -0.203 | (0.031) *** | | |
| Age 15 at testing | -0.014 | (0.023) | -0.029 | (0.022) | | |
| Cohort birth year | 0.007 | (0.001) *** | 0.001 | (0.003) | | |
| Study fixed effects (ref=TIMSS 2003 Grade 8) | | | | | yes | |
| School enrollment (proportion) | | | 0.539 | (0.145) *** | 0.486 | (0.107) *** |
| Immigrant background (proportion) | | | -0.215 | (0.283) | 0.226 | (0.250) |
| GDP per capita (logged) | | | 0.127 | (0.061) * | 0.055 | (0.059) |
| Income inequality (Gini) | | | -0.723 | (0.997) | -1.913 | (0.887) * |
| Mid/low-income country × Income inequality | | | 1.908 | (1.138) + | 2.539 | (1.129) * |
| Age when tracking begins | | | -0.027 | (0.016) + | -0.037 | (0.016) * |
| Private school enrollment (proportion) | | | -0.042 | (0.255) | 0.240 | (0.249) |
| Expecting higher education (proportion) | | | 0.050 | (0.093) | -0.029 | (0.094) |
| **Level 3 - Countries** | | | | | | |
| Mid/low-income country × Intercept interactions | | | | | | |
| Mid/low-income country × Parent education | -0.059 | (0.048) | 0.206 | (0.059) *** | 0.192 | (0.063) ** |
| Mid/low-income country × Parent occupation | -0.185 | (0.045) *** | 0.071 | (0.057) | 0.060 | (0.058) |
| Mid/low-income country × Books | -0.331 | (0.063) *** | -0.044 | (0.066) | -0.082 | (0.068) |
| Mean school enrollment | | | 0.652 | (0.282) * | 0.640 | (0.317) * |
| Mean proportion immigrant background | | | 0.068 | (0.239) | 0.134 | (0.261) |
| Mean GDP per capita (logged) | | | 0.143 | (0.039) *** | 0.142 | (0.041) *** |
| Mean income inequality | | | 0.653 | (0.303) * | 0.605 | (0.317) + |
| Mean age when tracking begins | | | -0.040 | (0.010) *** | -0.035 | (0.011) ** |
| Mean private school enrollment | | | 0.004 | (0.089) | 0.043 | (0.099) |
| Mean proportion expecting higher education | | | -0.155 | (0.165) | -0.255 | (0.169) |
| **Random effects** | | | | | | |
| *Level 2 - Residual variance between studies in…* | | | | | | |
| Parent education gaps | 0.04067 | | 0.03993 | | 0.03824 | |
| Parent occupation gaps | 0.02244 | | 0.02222 | | 0.02065 | |
| Household books gaps | 0.04051 | | 0.03738 | | 0.03683 | |
| *Level 3 - Residual variance between countries in…* | | | | | | |
| Parent education gaps | 0.04228 | | 0.02359 | | 0.02822 | |
| Parent occupation gaps | 0.03519 | | 0.02440 | | 0.02732 | |
| Household books gaps | 0.08273 | | 0.04353 | | 0.04944 | |
| Cohort slopes | 0.00003 | | 0.00004 | | | |
| N (Level 1 - gaps) | 4604 | | 4604 | | 4604 | |
| N (Level 2 - study-years) | 855 | | 855 | | 855 | |
| N (Level 3 - countries) | 78 | | 78 | | 78 | |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Notes: "Middle/low income" countries had GDPs per capita of less than $6000 in 1980 (the reference category is high-income countries; see Appendix Table A1 for coding). All level 2 time-varying country covariates are mean-centered within countries, meaning results can be interpreted very similarly to a model with country fixed effects (as well as study-year fixed effects, included at level 2). All models in this table specify known level 1 error variances and covariances, estimated via bootstrapping, consistent with models in the main text of the paper.

Finally, Table M7 reports multivariate results estimated using a fixed effects rather than a mixed effects (hierarchical growth curve) model. The model includes country fixed effects, meaning that country-level covariates are omitted due to collinearity. The model is estimated as follows:

$$\hat{G}_{ij} = \gamma_{00} + \gamma_{10}Y_{ij} + (\boldsymbol{X}_{ij})\mathbf{B} + \boldsymbol{\Gamma}_j + \boldsymbol{\Lambda}_{ij},$$

[M1]

where $\hat{G}_{ij}$ is the estimated gap in country $j$ in country-study-year $i$, $\gamma_{10}$ is the coefficient for cohort birth year $Y_{ij}$, $\boldsymbol{X}_{ij}$ is a vector of time-varying country covariates in country-year $i$, $\boldsymbol{\Gamma}$ is a vector of country dummy variables, $\boldsymbol{\Lambda}_{ij}$ is a vector of dummy variables indicating age at testing and test subject, $\boldsymbol{B}$ is a vector of coefficients for the time-varying country covariates. The country fixed effects are estimated using weighted least squares (weighted by the inverse squared standard error associated with each gap estimate, estimated via bootstrapping) and are reported with robust Huber-White standard errors.

Results for the country fixed effects models are very similar to those in the main results for Models 2 and 8 of the main text, both in terms of the cohort trend estimates in Model 2J (and for the reduced analytic sample in Model 2K) and the covariate results in Model 8C. Recall that time-varying country covariates in the hierarchical growth curve models are mean-centered within countries, meaning that they have an interpretation very similar to a fixed effects model.

Table M7. Country Fixed Effects Models Predicting 90/10 Achievement Gaps from Country Covariates and Cohort Birth Year

| | (2J) Country FE - full sample | | | (2K) Country FE - analytic sample | | | (8C) Country & study FE, covariates | | |
|---|---|---|---|---|---|---|---|---|---|
| | coef | (se) | | coef | (se) | | coef | (se) | |
| SES variable (ref=Parent education) | | | | | | | | | |
|   Parent occupation | -0.083 | (0.011) | *** | -0.063 | (0.012) | *** | -0.009 | (0.012) | |
|   Household books | 0.391 | (0.012) | *** | 0.399 | (0.013) | *** | 0.492 | (0.014) | *** |
| Subject (ref=Reading) | | | | | | | | | |
|   Math | 0.012 | (0.008) | | 0.011 | (0.008) | | 0.019 | (0.008) | * |
|   Science | 0.025 | (0.008) | *** | 0.023 | (0.008) | ** | 0.031 | (0.008) | *** |
| SES variable quality measures | | | | | | | | | |
|   Parent-reported × Parent education | 0.083 | (0.017) | *** | 0.071 | (0.018) | *** | -0.018 | (0.019) | |
|   Parent-reported × Parent occupation | -0.080 | (0.018) | *** | -0.080 | (0.019) | *** | -0.055 | (0.018) | ** |
|   Parent-reported × Books | -0.206 | (0.022) | *** | -0.156 | (0.022) | *** | -0.187 | (0.022) | *** |
|   Number of categories (centered at 7) | 0.006 | (0.004) | | 0.004 | (0.004) | | 0.002 | (0.004) | |
|   ≥ 20% in bottom category | -0.220 | (0.014) | *** | -0.213 | (0.015) | *** | -0.152 | (0.015) | *** |
|   ≥ 20% in top category | -0.128 | (0.008) | *** | -0.144 | (0.009) | *** | -0.124 | (0.009) | *** |
| Age at testing (ref=14) | | | | | | | | | |
|   Age 10 at testing | -0.136 | (0.017) | *** | -0.145 | (0.018) | *** | | | |
|   Age 15 at testing | -0.058 | (0.012) | *** | -0.058 | (0.013) | *** | | | |
| Cohort birth year | 0.007 | (0.001) | *** | 0.007 | (0.001) | *** | | | |
| Study fixed effects (ref=TIMSS 2003 Grade 8) | | | | | | | yes | | |
| School enrollment (proportion) | | | | | | | 0.397 | (0.077) | *** |
| Immigrant background (proportion) | | | | | | | 0.362 | (0.143) | * |
| GDP per capita (logged) | | | | | | | -0.051 | (0.049) | |
| Income inequality (Gini) | | | | | | | -2.176 | (0.449) | *** |
| Mid/low-income country × Income inequality | | | | | | | 3.429 | (0.629) | *** |
| Age when tracking begins | | | | | | | -0.035 | (0.009) | *** |
| Private school enrollment (proportion) | | | | | | | 0.025 | (0.144) | |
| Expecting higher education (proportion) | | | | | | | -0.095 | (0.045) | * |
| Mid/low-income country × SES interactions | | | | | | | | | |
|   Mid/low-income country × Parent occupation | | | | | | | -0.113 | (0.015) | *** |
|   Mid/low-income country × Books | | | | | | | -0.212 | (0.017) | *** |
| Country fixed effects | yes | | | yes | | | yes | | |
| Intercept | 1.152 | (0.013) | *** | 1.178 | (0.014) | *** | 1.979 | (0.451) | *** |
| Adjusted $R^2$ | 0.677 | | | 0.675 | | | 0.711 | | |
| N (gaps) | 5541 | | | 4604 | | | 4604 | | |
| N (study-years) | 1026 | | | 855 | | | 855 | | |
| N (countries) | 100 | | | 78 | | | 78 | | |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: All models in this table are estimated via weighted least squares (weighted by the inverse squared standard error associated with each gap estimate, estimated via bootstrapping).

**N. Trends in SES Achievement Gaps Using a Rank-Based Measure of Achievement**

The analyses in the main paper text standardize all achievement scores to a mean of 0 and standard deviation of 1 within each country-study-year-subject before calculating SES achievement gaps. Thus, trends in SES achievement gaps cannot be influenced by changes in the dispersion of achievement, as the dispersion is constrained to be constant over time. However, trends could still be influenced by changes in the *shape* of the achievement distribution. An examination of achievement distributions shows that, with very few exceptions, the distribution of achievement scores is close to normal in every country-study-year-subject. As an additional robustness check that trends in SES achievement gaps are not influenced by changes in the shape of the achievement distribution, I also estimate trends in gaps after converting achievement to a rank-based measure, which has a constant uniform distribution in every country-study-year-subject. Table N1 reports results of models estimating trends in SES achievement gaps computed from achievement rank. Within each country-study-year-subject, rather than standardizing achievement to a mean of 0 and standard deviation of 1, I instead convert achievement into percentiles scaled from 0 to 1. Using this transformed achievement measure, I then compute 90/10 gaps using the same method described in the main text of the paper.

The results for trends in SES gaps in achievement rank are very similar to those for trends in SES gaps in standardized achievement. The magnitude of all coefficients is smaller because gaps based on achievement rank are smaller than gaps based on standardized achievement. (A gap of 1.0 would indicate a gap of 100 percentile points of achievement between the 90th and 10th percentiles of SES, compared with a gap of 1.0 standard deviations of achievement between the 90th and 10th percentiles of SES.) Yet the overall story is very similar. 90/10 gaps based on all three SES measures have increased significantly ($p < .001$).

The trends in gaps in achievement rank also partially relax the assumption that achievement must be interval scaled. However, ranks are interval scaled only when interpreted *as ranks*; they are not necessarily interval scaled with respect to students' true achievement—the assumption that must hold in order for standardized achievement to be interval scaled in z-scores. Although it is not possible to test the interval scaling assumption directly, the consistency of results across the gaps using standardized achievement (in the main text), rank-based achievement in Table N1, and the original scales of PISA, TIMSS, and PIRLS (Appendix B) provides some indication that results are robust to the achievement scale used.

Table N1. Hierarchical Growth Models Predicting 90/10 SES Achievement Gaps Using a Rank-Based Measure of Achievement

| | (1) | | | (2) | | |
| | 3 Cohort Slopes | | | 1 Cohort Slope | | |
| | coef | (se) | | coef | (se) | |
|---|---|---|---|---|---|---|
| Parent education gaps intercept | 0.300 | (0.009) | *** | 0.303 | (0.009) | *** |
| Parent occupation gaps intercept | 0.272 | (0.009) | *** | 0.275 | (0.009) | *** |
| Household books gaps intercept | 0.373 | (0.012) | *** | 0.373 | (0.012) | *** |
| Level 1 - Gaps | | | | | | |
| Subject (ref=Reading): | | | | | | |
| Math | -0.001 | (0.002) | | -0.001 | (0.002) | |
| Science | 0.006 | (0.002) | *** | 0.006 | (0.002) | *** |
| SES variable quality measures | | | | | | |
| Parent-reported × Parent education | 0.025 | (0.007) | *** | 0.019 | (0.008) | * |
| Parent-reported × Parent occupation | 0.016 | (0.006) | ** | 0.015 | (0.006) | * |
| Parent-reported × Books | -0.029 | (0.008) | *** | -0.020 | (0.007) | ** |
| Number of categories (centered at 7) | 0.001 | (0.001) | | 0.001 | (0.001) | |
| ≥ 20% in bottom category | -0.014 | (0.006) | * | -0.015 | (0.006) | * |
| ≥ 20% in top category | -0.029 | (0.003) | *** | -0.033 | (0.003) | *** |
| Level 2 - Study-years | | | | | | |
| Age at testing (ref=14) | | | | | | |
| Age 10 at testing | -0.042 | (0.007) | *** | -0.042 | (0.007) | *** |
| Age 15 at testing | -0.007 | (0.006) | | -0.007 | (0.006) | |
| Cohort birth year × Parent education | 0.0019 | (0.0003) | *** | | | |
| Cohort birth year × Parent occupation | 0.0021 | (0.0003) | *** | | | |
| Cohort birth year × Books | 0.0024 | (0.0004) | *** | | | |
| Cohort birth year | | | | 0.0022 | (0.0003) | *** |
| Random effects | | | | | | |
| *Level 2 - Residual variance between studies in…* | | | | | | |
| Parent education intercepts | 0.002230 | | | 0.002270 | | |
| Parent occupation intercepts | 0.001560 | | | 0.001580 | | |
| Books intercepts | 0.002650 | | | 0.002750 | | |
| *Level 3 - Residual variance between countries in…* | | | | | | |
| Parent education intercepts | 0.004780 | | | 0.004730 | | |
| Parent occupation intercepts | 0.004630 | | | 0.004750 | | |
| Books intercepts | 0.010300 | | | 0.010660 | | |
| Parent education cohort slopes | 0.000004 | | | | | |
| Parent occupation cohort slopes | 0.000003 | | | | | |
| Books cohort slopes | 0.000005 | | | | | |
| Cohort slopes | | | | 0.000003 | | |
| N (Level 1 - gaps) | 5541 | | | 5541 | | |
| N (Level 2 - study-years) | 1027 | | | 1027 | | |
| N (Level 3 - countries) | 100 | | | 100 | | |

+ p<.1, * p<.05, ** p<.01, *** p < .001. Note: To avoid very long computation times, all models in this table specify known level 1 error variances estimated using conventional non-bootstrap formulas and omit error covariances. This simplified specification appears to produce very similar results to models using bootstrapped error variances and covariances (see Appendix L).

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