Using covariates to sharpen bounds

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Early College High Schools

Autonomous small schools
managed by local school districts
in partnership with two- or four-year colleges

Students...

‣ begin taking college-level courses in ninth grade
‣ are expected to graduate high school with an associates’ degree and/or two years of transferable college credit

Goal: improve high school graduation and college going.

Hoped-for Mechanisms: environment promotes college going; powerful teaching and learning; small school size (400 students max); positive, supportive relationships between students & staff; effective leadership
And they were studied!

The **ECHS Study** is a randomized control study.

**44 lotteries** conducted at **19 schools** in **6 years**

First cohort entered 9\(^{\text{th}}\) grade in 2005

Data sources:
- North Carolina Dept. of Public Instruction (NCDPI)
- NC Community College System
- National Student Clearinghouse
Schools, scattered across the state

Each site x year combo is its own study.
We might naturally expect substantial variation in those stories.
Existing research: Positive ECHS treatment effects

<table>
<thead>
<tr>
<th>Outcome</th>
<th>N</th>
<th>Unadjusted Means</th>
<th>Adjusted ITT Estimate</th>
</tr>
</thead>
<tbody>
<tr>
<td>Grade 9 On Track</td>
<td>3855</td>
<td>97.1%</td>
<td>90.4%</td>
</tr>
<tr>
<td>College Credits Earned in HS</td>
<td>3402</td>
<td>21.7</td>
<td>2.5</td>
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<tr>
<td>Graduation</td>
<td>2941</td>
<td>87.3%</td>
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*p < 0.05

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Our principal strata: quality of counterfactual school

There are three types of school:

- **ECHS (E)** - the “treatment” schools
- **High-quality (HQ)** - high schools that are rated as being of good quality.
- **Low-quality (LQ)** - high schools that are not so rated.

**Our Research Question**

To what extent do the effects of Early College High Schools (ECHSs) vary according to the quality of the school students would have otherwise attended?
Students attend different types of high schools

- **Treatment**
  - ECHS: 86%
  - Low: 11%
  - High: 3%

- **Control**
  - ECHS: 83%
  - Low: 3%
  - High: 14%
Most students are on track in 9th grade

<table>
<thead>
<tr>
<th></th>
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<th>Not On Track</th>
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<tbody>
<tr>
<td>Treatment</td>
<td>94%</td>
<td>6%</td>
</tr>
<tr>
<td>Control</td>
<td>88%</td>
<td>12%</td>
</tr>
<tr>
<td>Total</td>
<td>91%</td>
<td>9%</td>
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Remember this tidbit:
Being near boundaries increases the chance of bounds being viable.
The school-attendance behavior gives our boxes

If we assign to **Treatment**, they go to...

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<td></td>
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<td>LQAT</td>
<td></td>
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*Defiers*
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If we assign to **Control**, they go to...

If we assign to **Treatment**, they go to...

- Defiers
- Irrelevant Alternatives
What we estimate, and what is mixed up

To get our **complier average treatment effects** we need:

- The mean outcomes of our LQC and HQC groups under control
- The mean outcomes of our LQC and HQC groups under treatment

For **control side:**

- Our LQCs are mixed with our LQATs, but we see our LQATs on the treatment side
- Same for our HQCs

For **treatment side:**

- Our LQCs and HQCs and EATs are all mixed up!
- We can see EATs on the control side, but we have no way to separate treated LQCs and HQCs
- We turn to bounds to assess what possible mean outcomes there could be for these two groups.
Bounds and always-takers

Under the **exclusion restriction** we can observe Always-Takers and know their mean outcome under either treatment arm.

This means we can “subtract them out.”

The presence of always-takers does not impact the bounds, only the estimation of the bounds.
Example of “Subtracting Out” to get the Control LQC mean

\[ \bar{Y}_{1b} \leftarrow \text{we see this directly in our data} \]

\[ \bar{Y}_{0b} = \frac{\pi_{Lg} \mu_{0Lc}}{\pi_{Lc} + \pi_{Lgat}} + \frac{\pi_{Lgat} \mu_{0Lgat}}{\pi_{Lc} + \pi_{Lgat}} \]

\[ \Delta \mu_{0Lc} = \frac{\pi_{Lg} + \pi_{Lgat}}{\pi_{Lc}} \bar{Y}_{0b} - \frac{\pi_{Lgat}}{\pi_{Lc}} \mu_{0Lgat} \]

\[ = \frac{\pi_{Lg}}{\pi_{Lg} - \pi_{Lc}} \bar{Y}_{0b} - \frac{\pi_{Lg}}{\pi_{Lg} - \pi_{Lc}} \bar{Y}_{1b} \]
Our problem (without always takers)

We have
- Low-Quality Compliers (LQC) and
- High-Quality Compliers (HQC)

We see
- Mean outcomes and proportions under the control condition
- One amorphous blob under the treatment condition

We want
- The mean outcome under the treatment condition
Calculating the bound (geometrically)

All our HQC + LQC

Area = 1

these are LQC kids

Area = \( \frac{\text{TP}_{\text{LQC}}}{\text{TP}_{\text{HQC}} + \text{TP}_{\text{LQC}}} \) = \( f_{\text{LQC}} \)

Let \( g \) = fraction of kids who are LQC and \( y = 1 \)

\[ M_{\text{LQC}} = \frac{g}{f_{\text{LQC}}} \]

How big? \( \min(f_{\text{LQC}}, \bar{Y}_{1e}) \)

How small? \( \max(0, (1 - \bar{Y}_{1e}) - (1 - f_{\text{LQC}})) = \max(0, f_{\text{LQC}} - \bar{Y}_{1e}) \)

Finally \( M_{\text{LQC}} \leq \min(1, \frac{\bar{Y}_{1e}}{f_{\text{LQC}}}) \)

\( M_{\text{LQC}} \geq \max(0, 1 - \frac{\bar{Y}_{1e}}{f_{\text{LQC}}}) \)

these kids did well (\( y = 1 \))

Area = \( \bar{Y}_{1e} \)

(assuming no EATs)
Our final bounds

We have

$$\mu_{lqc,1}^{low} - \mu_{lqc,0} \leq ITT_{lq} \leq \mu_{lqc,1}^{high} - \mu_{lqc,0}$$

with

$$\mu_{lqc,1}^{low} = \max \left( 0, \frac{1}{\pi_{lqc}} \bar{Y}_{1e} - \frac{1 - \pi_{lqc}}{\pi_{lqc}} \right)$$

$$\mu_{lqc,1}^{high} = \min \left( \frac{1}{\pi_{lqc}} \bar{Y}_{1e}, 1 \right)$$

$$\mu_{lqc,0} = \bar{Y}_{0lq}$$
Our final bounds in terms of observed quantities

We have

\[ \mu_{lqc,1}^{low} - \mu_{lqc,0} \leq I T T_{lq} \leq \mu_{lqc,1}^{high} - \mu_{lqc,0} \]

with

\[ B_{lqc} = \frac{1}{p_{0lq}} \bar{Y}_{1e} \]

\[ \mu_{lqc,1}^{low} = \max \left( 0, B_{lqc} - \frac{1 - p_{0lq}}{p_{0lq}} \right) \]

\[ \mu_{lqc,1}^{high} = \min \left( B_{lq}, 1 \right) \]

\[ \mu_{lqc,0} = \bar{Y}_{0lq} \]
A hypothetical population

This is a “perfect information” picture. We don’t actually know which people are on which side.
We see a mixture

Since we do not know who is a LQC or HQC, we could imagine them all being an average outcome...

This would give an overall, pooled estimate.
We see a mixture

We see an average of $\frac{35}{60} = 58\%$

Proportion of LQC is $\frac{36}{60} = 60\%$

Since we do not know who is a LQC or HQC, we could imagine them all being an average outcome...

This would give an overall, pooled estimate.
Or we can think “worst case”

Here we try and see how well of the LQC could be, best case scenario.

This is also how badly off the HQC could be. Their worst case scenario.
Our actual results

<table>
<thead>
<tr>
<th>ECHS Always Takers EATs</th>
<th>Low-quality compliers LQCcs</th>
<th>Low-quality compliers LQCcs</th>
<th>Low-quality Always Takers LQATs</th>
<th>High-quality Always Takers HQATs</th>
</tr>
</thead>
<tbody>
<tr>
<td>p = 3%</td>
<td>p = 72%</td>
<td>p = 72%</td>
<td>p = 12%</td>
<td>p = 3%</td>
</tr>
<tr>
<td>Yt = Yc = 100%</td>
<td>Yt = (95% - 100%)</td>
<td>Yt = (95% - 100%)</td>
<td>Yt = Yc = 84%</td>
<td>Yt = Yc = 90%</td>
</tr>
</tbody>
</table>

To get our estimates, we subtract the two means:

\[
\text{ITT}_{lqc} = +8 - +13\% \text{ effect}
\]

\[
\text{ITT}_{hqc} = -33 - +1\% \text{ effect}
\]
The tradeoff of treatment effects
## Comparing different stratifications

<table>
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<tr>
<th></th>
<th>High Quality Compliers</th>
<th>Low Quality Compliers</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Control Mean</td>
<td>Bounds on Tx Mean</td>
</tr>
<tr>
<td>No Strat</td>
<td>98.8</td>
<td>65.5 - 100</td>
</tr>
<tr>
<td>by achievement</td>
<td>98.3</td>
<td>76.7 - 100</td>
</tr>
<tr>
<td>by middle school</td>
<td>98.4</td>
<td>86.8 - 100</td>
</tr>
<tr>
<td>by both</td>
<td>97.8</td>
<td>87.2 - 100</td>
</tr>
</tbody>
</table>
Now say we had a “prognostic covariate”

For ECHS we used grade 8 test scores.
We use it to create four equal sized slices of our population.
Inside each slice we can identify the proportion of LQC and HQC (in this example the proportions turn out to be the same)
We then do our same bounding exercise within each slice...
Create our worst case (for HQC) by calling as many successes as we can LQCs

Our first two slices are just as bad as our original for HQC (a 0 lower bound)

But our next two are better!

We then combine (average) to get an overall bound.
Some calculations

\[ HAC: \quad 0 + 0 + \frac{1}{6} \cdot \frac{1}{4} + \frac{4}{6} \cdot \frac{1}{4} = \frac{5}{24} \approx 0.21 \]

\[ LAC: \quad \frac{5}{9} \cdot \frac{1}{4} + \frac{7}{9} \cdot \frac{1}{4} + 1 \cdot \frac{1}{4} + 1 \cdot \frac{1}{4} = \frac{30}{36} \approx 0.83 \]

So \( HAC \) lower bound \( \approx 0.21 \) from 0

\( LAC \) upper bound \( \approx 0.83 \) from 0.97
Some calculations

\[ H_{AC} : 0 + 0 + \frac{1}{6} \cdot \frac{1}{4} + \frac{4}{6} \cdot \frac{1}{4} = \frac{5}{24} \approx 0.21 \]

\[ L_{AC} : \frac{5}{9} \cdot \frac{1}{4} + \frac{7}{9} \cdot \frac{1}{4} + 1 \cdot \frac{1}{4} + 1 \cdot \frac{1}{4} = \frac{30}{36} \approx 0.83 \]

So, \( H_{AC} \) lower bound \( \rightarrow 0.21 \) from 0

\( L_{AC} \) upper bound \( \rightarrow 0.83 \) from 0.97

This is some serious improvement!

(There is nothing like an artificial example.)
What other kinds of variables might we consider for slicing up our data?
Here we have a “propensity score” covariate that predicts complier type.

Slice based on a covariate that is predictive of compliance type.

High values of $X_1$ indicate a higher proportion of LQC students.

In the ECHS study, this was the proportion of a student’s middle school cohort who went to a LQ high school.
We again look at worst-case within each slice of our sample.

Again, our bounds are tightened in some of our slices and not others.

But notice how the relatively bigger slices are the ones with the tighter bounds!

Averaging here is a bit trickier... Let’s talk about it.
Some further calculations

For $\text{HQC}$ we weight by $\#$ folks: $N_{\text{HQC}} = 24$

$$
\mu \geq 0 \frac{1}{24} + 0 \frac{4}{24} + \frac{6}{8} \frac{5}{24} + \frac{11}{11} \frac{11}{24} = \frac{6}{24} \implies \frac{3}{12} = 0.25
$$

For $\text{LQC}$

$$
\mu \leq \frac{9}{14} \frac{11}{36} + \frac{9}{11} \frac{11}{36} + 1 \frac{2}{36} + 1 \frac{4}{36} = \frac{29}{36} \approx 0.81
$$

Comparisons
No bounds:
$\text{HQC} > 0 \quad \text{LQC} < 0.97$

Prognostic bounds:
$\text{HQC} > 0.21 \quad \text{LQC} < 0.83$
Some further calculations

Full disclosure:
I kinda cheated. I tweaked the variables so that this covariate was better by a little bit than last.

(More exciting.)

But this does raise a question...

HQC > 0.21    LQC < 0.83

Prognostic bounds:
HQC > 0.21    LQC < 0.83
Choices, choices...

How do we decide what to stratify on?

Two primary choices:

- “prognostic score” - covariate predictive of outcome.
- “propensity score” - covariate predictive of compliance status.
Stratifying by Propensity Score
Different strata are differently effective for LQC and HQC
Stratifying by Prognostic Score: Tighter bounds at the margins

[-4.95, -1.41]  (-1.41, -0.63]  (-0.63, -0.025]

(-0.025, 0.58]  (0.58, 1.31]  (1.31, 5.94]
A simulation study

We will generate fake data indexed by some latent traits that govern school quality, compliance, and outcome.

We will then stratify these covariates and look at the resulting bound widths.

This could shed light on when covariates are powerful, or not.
The data generation process

\[ U_{1i}, U_{2i}, U_{3i} \sim N(0, 1) \text{ (all independent)} \]

\[ H_i = \begin{cases} 
  hq & U_{1i} > \Phi^{-1}(\gamma_0) \\
  lq & \text{else}
\end{cases} \]

\[ S_i(0) = \begin{cases} 
  e & U_{2i} > \Phi^{-1}(1 - \pi_{eat}) \\
  H_i & \text{else}
\end{cases} \]

\[ S_i(1) = \begin{cases} 
  e & U_{2i} > \Phi^{-1}(\gamma_1(H_i)) \\
  H_i & \text{else}
\end{cases} \]

\[ Y_i(z) = \logit^{-1}(\beta U_3 + \omega (S_i(z))) \]
The data generation process

\[ U_{1i}, U_{2i}, U_{3i} \sim N(0, 1) \text{ (all independent)} \]

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  lp & \text{else} 
\end{cases} \]

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The gamma and omega parameters govern marginal means, etc. They are functions of desired strata proportions and effect sizes. Beta governs strength of \( U_3 \) and the outcomes.
What we observe

Generate a population as above

We then observe *noisy versions* of the latent variables:

\[ X_{it} = U_{it} + \varepsilon_{it} \text{ with } \varepsilon_{it} \sim N(0, \sigma_t^2) \]

We then randomize our units to treatment and control, and observe the relevant potential outcomes to get our data:

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<tr>
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<th>Z</th>
<th>S</th>
<th>Yobs</th>
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<tr>
<td>1</td>
<td>1</td>
<td>2.27</td>
<td>-1.68</td>
<td>-1.38</td>
<td>0</td>
<td>lq</td>
</tr>
<tr>
<td>2</td>
<td>2</td>
<td>0.56</td>
<td>-0.17</td>
<td>-0.55</td>
<td>1</td>
<td>e</td>
</tr>
<tr>
<td>3</td>
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We vary these to get more or less useful Xs.
Bound Width for LQC Group

The graph shows the relationship between the width of the interval and the amount of noise in covariates, across different methods.

- **No strat** (red line) shows a consistent trend with a width of around 0.95 across all noise levels.
- **X1: principal** (yellow line) and **X2: compliance** (green line) exhibit a more pronounced increase in width as the noise increases, with **X1: principal** starting higher and **X2: compliance** lower.
- **X3: prognostic** (blue line) starts at a lower width and increases more steadily than **X2: compliance**.
- **X1+X3: princ + prog** (purple line) has the highest width, starting at around 0.9 and rising to 1.0 as the noise increases.

The x-axis represents the amount of noise in covariates, ranging from 0 to 3, and the y-axis represents the width of the interval.
Bound Width for HQC Group
Conclusions

Covariates can substantially sharpen bounds

- Both simulation studies and empirical examples (not shown) show substantial impact of using covariates.

Propensity scores seem to be more effective than prognostic scores

- In simulation studies, propensity scores “won”
- A caveat: We do not have easy way of “comparing predictiveness” so comparisons might not be fair.
- Cross-cutting data and using both does seem to be even more effective, however.
Thank you!